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# Defining the climatic signal in stream salinity trends using the Interdecadal Pacific Oscillation and its rate of change

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## Abstract

The impact of landuse on stream salinity is difficult to separate from decadal climatic variability, as the decadal scale climatic cycles in ground water and stream hydrology have similar wavelengths to the landuse pattern. These hydrological cycles determine the stream salinity through accumulation or release of salt in the landscape. The Interdecadal Pacific Oscillation (IPO) has been investigated before as an indicator of hydrological and related time series in the southern hemisphere. This study presents a new approach, which uses the rate of change in the IPO, rather than just its absolute value, to define an indicator for the climate component of ambient shallow groundwater tables and corresponding stream salinity. Representative time series of water table and stream salinity indicators are compiled, using an extensive but irregular database covering a very wide geographical area. These are modelled with respect to the IPO and its rate of change to derive climatic indicators. The effect of removing the decadal climatic influence from stream salinity trends is demonstrated.

## 1 Introduction

Variations or trends in salinity represented by Electrical Conductivity (EC) are often measured in streams and have the potential to be used as an indicator of landuse sustainability. Such measurements are relatively low cost and are convenient to automate; they also give important insights into the incipient development of dryland salinity, impacts of land clearing, irrigation, drainage alterations, and urbanisation. However, such sources can only be relied on to define human impacts if they can be differentiated from variations that result from broad climate fluctuations.

The state of Queensland encompasses a considerable land area. It extends from 9° S to 29° S in latitude, and from 138° E to 154° E in longitude, and occupies most of the northeast quadrant of Australia. The influence of climatic fluctuations on stream salinity was demonstrated when conductivity trends were calculated for about 500 gauging sta-

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tions throughout the state over a twenty year period between 1970 and 1990 (QDPI, 1994). The results revealed a complex but consistent trend pattern evident in many streams, despite a very diverse range of topography, geology and landuse. Subsequently it was found that stream EC and shallow unconfined water tables are related to climatic indicators such as the Southern Oscillation Index (SOI) and Interdecadal Pacific Oscillation (IPO) (McNeil and Cox, 2002).

It was conceptually reasonable that these trends were at least partly related to hydrological regimes of ground and surface water interaction. Salt is released into catchments through weathering, precipitation and dry fallout. Losses occur through runoff, wind blown dust, or percolation through the soil to the water table. The balance between sources and sinks of salt may vary according to prevailing rainfall and temperature patterns which control the dynamic interactions between ground and surface water. Queensland rainfall is exceptionally variable both temporally and spatially (Nicholls, 1988; Loewe and Radok, 1948), with cycles covering decadal time scales. Major floods are often followed by a period of dry years, which may terminate in a prolonged drought.

Baseflow is affected by groundwater and is generally more saline than overland flow, due to prolonged contact with weathering minerals, so it raises the average salinity of streams. It is hypothesised that a flood flushes accumulated salts from the surface and upper soil horizons, resulting in large flows of dilute surface water. The wet conditions also recharge the groundwater, which can result in a greater contribution of baseflow over the next few drier years. If the dry period continues, baseflow is lost as water tables drop below the streambeds, a feature typical of alluvial aquifers. This causes the streams to become more ephemeral, surface water dominated, and consequently less saline. The declining salinity trend continues slowly, until the next flood event, when there is a sharp drop followed by the subsequent rise. This hypothesis would be verified if water tables, streamflow, and stream salinity could be related to climatic variables in a manner consistent with the observed patterns of occurrence.

A clear link between climate and streamflow has been established by a number of authors including Chiew et al. (1998) and Lough (1991), but apart from Vaccaro

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(1993) who reported groundwater levels to be sensitive to climatic variability, and Evans et al. (2001) who related salt concentration in lakes to variation in the North Atlantic Oscillation (NAO), there have been few studies to date on the numerical relationship between decadal climate indicators and water tables or surface water salinity.

5 The purpose of this study is to verify that climatic trends are present in the records of Queensland stream salinity and its associated hydrology, using a new approach which incorporates both the Interdecadal Pacific Oscillation (IPO) and its rate of change ( $\Delta$ IPO) to derive decadal scale climatic indicators.

## 2 Choice of climatic indicators

10 The influence of climate on hydrology and subsequently stream salinity is not solely related to precipitation, because temperature and evapotranspiration affect runoff, soil moisture storage and ultimately groundwater levels (Chiew et al., 1998). Therefore a more general and geographically broadscale climate indicator is required than could be provided by local observations such as running means of rainfall.

15 The most commonly used general climatic indicator is the El Nino/Southern Oscillation Index phenomenon (ENSO/SOI). It has been linked to precipitation and streamflow throughout the world by Philander (1990) and others. In addition there are studies on many specific regions, for instance: USA (Gray, 1990; Kahya and Dracup, 1993; Zorn and Waylen, 1997); Peru (Henderson et al., 1990); New Zealand (McKerchar et al., 1996; McKerchar et al., 1998; Gordon, 1986; Mullan, 1995); Belgium (Gellens and Roulin, 1998) and western Africa (Gray, 1990). A number of authors including McBride and Nicholls (1983), Ropelewski and Halpert (1987, 1989), Lough (1991), and Chiew et al. (1998) have indicated the effect of ENSO/SOI on Queensland's rainfall. Chiew et al. (1998) and Lough (1991) also established that as in the case of rainfall, there  
20 was a clear link between Australian streamflow and the ENSO/SOI, but Lough (1991), Cordery et al. (1993), Burt and Shahgedanova (1998), and Power et al. (1999; 2005) have all observed that the relationship of ENSO/SOI with hydrology over time is defi-

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nately non-stationary. This leads to problems in fitting long term hydrological time series based on the SOI alone, indicating that its interaction with other global climatic influences should also be considered.

Other modes of climatic variability have been detected on decadal time scales that are controlled by Sea Surface Temperature (SST) anomaly patterns. These may be more aligned than the rapidly fluctuating ENSO/SOI to the longer term trends required to extract the climate signal from landuse in natural systems. They have been identified through complex empirical orthogonal functions (CEOF) (Trenberth and Hurrell, 1994; Rowell and Zwiers, 1999; Zhang et al., 1997). CEOF is a statistical technique used to detect structure revealed by very slow decay of percent variance in noisy data. The most significant climatic signals that have been identified to date by CEOF analysis of global SST data are a secular trend representing global warming (IPCC, 1996) and the Inter-decadal Pacific Oscillation (IPO) described by Folland et al. (1999, 2002a, 2002b) and Power et al. (1999, 1998). Similarities have been noted between the ENSO/SOI and the IPO, although they were derived independently (Folland et al., 2002a; 2002b). The IPO has been proved to modulate ENSO/SOI climate variability in the Pacific region. (Salinger et al., 2001; Power et al., 2005) In particular, during positive phases of the IPO ( $IPO > 0.5$ ) El Nino events are noted to be more common, whereas La Niña events tend to occur more under the negative IPO phase ( $\leq 0.5$ ).

## 2.1 The rate of change of the IPO as a climatic indicator

Recent studies (Kiem et al., 2002; Kiem and Franks, 2004; Franks, 2002; Jones and Everingham, 2005) and (McKeon et al., 2004) have incorporated both the IPO and the ENSO/SOI in their analyses of climatic effects on hydrology by dividing the IPO into positive, negative and neutral phases as recommended by Power et al. (1999), and modelling the likelihood of events on ENSO/SOI separately within each phase. However, a more continuous indicator was required to analyse the effect of climate on a time series. It was also decided that a simple climatic indicator would be desirable to model the effect on stream salinity and groundwater levels. Whereas the effect

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of ENSO/IPO on stream flow has been well established as described above, there have been few climatic studies on groundwater levels and associated stream salinity. It is expected that reasonable correlations based on a simple model would be more convincing than a more complex model with greater skill.

5 The Burdekin River is one of the biggest rivers in Queensland, and has a flow history extending back to the 1920s. It is also reasonably indicative of general hydrological conditions over a wide area of Queensland (Allan, 1985). Figure 1 shows annual flows in the Burdekin compared with the record of the IPO. The IPO has been estimated since 1870, and appears to rise irregularly then fall rapidly in phases of approximately  
10 50 years. At the present time, the long term (50 year) trend is rising. It is apparent that the general conditions of flow, particularly as indicated by the 3 year geometric mean, can be related to the behaviour of the IPO. Flows tend to be above the median when the IPO is changing rapidly, particularly when it is rising. As the rate of change slows, flows diminish, and droughts are most common when the rate of change is close to  
15 static. On the other hand, very high flows can occur, sometimes anomalously, when a rapid slowing of the IPO leading to a change in direction takes place. Examples of such events are 1928, 1935, 1941, 1950–51, 1959, 1973–75, 1982, 1992, and 2001. Table 1 lists the expected stream flow according to the behaviours of the IPO. The similarity was considered sufficient to trial the IPO and its rate of change as the sole  
20 indicators to test for climatic variability in stream salinity and water levels.

### 3 Representative time series to model the climate signal

#### 3.1 Data used in the study

The data used in this study are from the hydrological and chemical databases held by the Queensland Department of Natural Resources, Mines and Water (NRMW). Although these data are variable in content, reliability and periodicity of sampling, it was  
25 expected that these would be evened out by the use of monthly medians which took

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account of the whole population. The surface water quality database contains nearly 50 000 EC measurements, collected from the 815 gauging stations since about 1962. The average number of samples per site is around 70, usually collected about four times a year. Some streamflow records date back to about 1920.

5 The groundwater data base contains about 500 000 individual water level (WL) measurements from over 30 000 bores, most of which were only sampled once or twice, usually when first drilled. Although shallow groundwater is used throughout the state, the great majority of bores are concentrated in irrigated alluvial plains which have mostly been cleared and developed. Bores are particularly concentrated within irriga-  
10 tion areas, and therefore at risk of interference from pumping or regional depression, or rises due to deforestation, watering, or disturbance to natural drainage patterns.

### 3.2 Producing annual time series for EC, water levels, and climate indicators

Time series were produced from monthly medians of EC and water tables shallower than about 30 m. All non-artesian WL measurements no deeper than 30 m, and all  
15 EC measurements taken at gauged sites within the particular month were included, regardless of location. Annual time series were based on the climatic years (October of the preceding year through to September) to avoid splitting the (Southern Hemisphere) summer wet season, as recommended by Loewe and Radok (1948). Annual values were calculated if all months were represented for the climatic year. The arithmetic  
20 mean was used in the case of WL, and the geometric means in the case of EC, in line with observed distributions. Finally, a smoothing technique (3 year moving geometric mean) was used to produce series in which seasonal variation was smoothed and the impact of irregularity and outliers minimised.

This exercise produced an almost continuous annual time series for EC since 1960,  
25 and WL since about 1950, but not all years were considered to be sufficiently representative for use in constructing the climatic model. The representative time series contained only years with a reasonably comprehensive and consistent coverage of the state, and at least 200 measurements in the case of EC, and 1000 in the case of ground-

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water. Figure 2 summarises the distribution of EC measurement sites in Queensland over the period of record. The representative period for modelling was taken as 1971–2002 for EC, and 1968–2001 for WL. It is accepted that the distribution of gauging stations shown on Fig. 2 indicates that the EC time series would be most representative of the coastal and central regions. However, the similarity of the pattern with that observed throughout the state in QDPI (1994) indicates that it is general enough for a preliminary study aimed at verifying the existence of climatic influence in the trends rather than closely defining them. The pattern also resembles plots prepared by Jolly and Chin (1991) for bores in undisturbed areas of the Northern Territory to the west of Queensland which have been monitored since the 1950s.

The January value of the slowly varying IPO was selected to represent the climatic year in the middle of the wet season. The  $\Delta$ IPO was taken as the difference between annual values. Figure 3 displays these climatic indicators with the EC and WL time series.

## 4 Modelling EC and water levels with the IPO and its rate of change

Modelling was carried out on the representative WL and EC time series through linear and second order polynomial regression with the IPO and its rate of change. Nonlinearity in the relationship between climatic indicators and hydrological time series has been noted by other authors such as Power et al. (2005). They pointed out that whereas the upper limits of hydrological events are not confined, lower limits are generally limited. The model was then applied to all years since 1950 in the case of WL and since 1960 in the case of EC.

### 4.1 Groundwater levels

Modelling the WL time series was relatively simple. The first stage was a second order polynomial regression using the  $\Delta$ IPO (Fig. 3), which accounted for 50% of the total

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variance. The equation is:

$$WL = 4.1633\Delta IPO^2 + 2.0116\Delta IPO - 8.1148 \quad (1)$$

The residuals from the first model were then plotted against the IPO to include the longer term trend. This produced the polynomial:

$$5 \text{ Residual}WL = 0.42IPO^2 - 0.47IPO - 0.3298. \quad (2)$$

The final model of 3 years moving average water tables for bores less than 30 m was

$$WL = 0.42IPO^2 - 0.47IPO + 4.1633\Delta IPO^2 + 2.0116\Delta IPO - 8.4446 \quad (3)$$

The correlation coefficient of the model with the time series 19, as shown on Fig. 4 was  $R^2=0.81$ . Figure 5 compares the model and the WL series over time

## 10 4.2 Stream EC levels

Modelling the EC time series was less straightforward than the WL. For a start, the  $\Delta IPO$  appears to vary stochastically about its median over the representative period, and the WL also displays little overall trend. However, as is clear from Fig. 2, the EC series, whilst similar in phase to the WL, shows a significant downward trend, in opposition to the upward trend shown by the IPO over the period. Walsh et al. (2002) have also noted a decline in long term Queensland stream flows. Because stream salinity reflects the interactions between ground and surface waters, it was decided to first examine any component of the EC series that might be related to the long term gradient of the IPO. The period selected to reflect long term change was 50 years, 15 1950 to 2000, based on the apparent phases in the IPO. This trend was removed from the EC data (available 1960 to 2000) by regression against the IPO trend, leading to the following equation, which accounted for 45% of the variance:

$$20 \text{ Trendin}EC = -65.998(1950 - \text{CurrentYear}) + 392.2. \quad (4)$$

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The detrended EC, or residual component of the EC series, was then compared with the  $\Delta$ IPO, as shown on Fig. 7. The series values for negative  $\Delta$ IPO are also plotted with the  $\Delta$ IPO sign reversed. This plot shows that the series is positively correlated with  $\Delta$ IPO when  $\Delta$ IPO is positive, but negatively correlated when  $\Delta$ IPO is negative. This bimodal behaviour is understandable, because high residuals corresponding to low EC occur during extreme rises and falls of the IPO, but low residuals are only conceptually likely when the IPO is approaching stability. This bimodality is related to the way that very wet conditions and very prolonged dry spells both reduce stream EC, as discussed in the Introduction. It is difficult to fit a simple algorithm to the untransformed series, but the reversed sign values fit sufficiently well with the positive values to indicate that the rate of change of the IPO is sufficient to base the model on, regardless of the direction of change. Therefore, the detrended EC series was modelled on the absolute value of  $\Delta$ IPO using second order polynomial regression, which accounted for 34 percent of the remaining variance. The equation is:

$$\text{DetrendedEC} = -322.3 \times |\Delta\text{IPO}|^2 + 321.82 \times |\Delta\text{IPO}| - 42.263 \quad (5)$$

The final model was thus determined as:

$$\text{EC} = -66 \times (\text{CurrentYear} - 1950) - 322 \times (|\Delta\text{IPO}|)^2 + 322 \times |\Delta\text{IPO}| + 350 \quad (6)$$

Figure 8 shows that this relationship has a correlation coefficient of over 50% ( $R^2=0.53$ ) with the 3-year moving average of EC measured at gauging stations. Figure 9 shows the model reasonably reflects EC trend direction in the years before 1970, where the data was not sufficiently representative to be included in the model calibration.

#### 4.2.1 Field evidence of the modelled climate signal

In the case of the WL representative series, it was found by McNeil and Cox (2002) that despite the effects of pumping, the average extensively measured bore shows a slight positive association with the climate indicator of just under 10%. In fact, 30% of bores show a correlation of at least 30%, and more than 10% are over 50% correlated.

Some examples of highly correlated water level records from various parts of the state are shown on Fig. 10 from McNeil and Cox (2002) with the representative WL series (which has since been refined).

Recent studies have suggested that the EC records of many Queensland streams show a 10% to 50% correlation with the representative EC series, with downstream catchments and shorter term periods (5 to 10 years) apparently being the most influenced by climatic cycles. Although the time series was derived predominantly from the southeast, the patterns in the series have also been found to be mirrored at various sites in most parts of the state, in continuance with those patterns observed in QDPI (1994). One such site was selected to trial the removal of climate trend from the EC record.

## 5 The climate signal in stream salinity trends

The main objective of the study is the separation of landuse and climate EC trends using a representative control series in which the climatic signal had been verified. Because field EC data is collected several times a year and affected by season, the monthly version of the representative EC series was preferred as the appropriate climate control after being smoothed by a three month moving average. For testing, a gauging station with a long EC history was selected in a catchment which was extensively developed several decades ago. This was GS422316 on the Condamine River at Chinchilla, which has 117 EC readings collected between 1963 and 2004. These readings were paired with the representative EC value for the closest date.

For this exercise, a test was required that would detect monotonic trend in stream EC records with non-normal distributions, seasonality and missing values. The rank based seasonal Kendall test defined by Hirsch et al. (1982) was selected, having been already trialled on a state-wide basis in QDPI (1994). Very large short term variability occurs in stream EC simply from changes in flow, so it is removed from the data before a test for long term trend is applied. This was carried out using an algorithm derived

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by Thorburn et al. (1992) which produces an S shaped curve based on the assumption that the salinity will be asymptotic to that of the groundwater during baseflow, and to that of superficial runoff during high flood.:

$$EC = \frac{EC_1 - EC_i}{1 + BQ^{1/m}} + EC_i \quad (7)$$

where EC is salinity measured in the stream in  $\mu\text{S}/\text{cm}$ ,  $EC_1$  is the median salinity under lowest flow conditions,  $EC_i$  is the salinity approached by surface runoff during highest flows, Q is streamflow in  $\text{m}^3 \text{s}^{-1}$ , and B and m are constants derived by optimisation.

A climate correction was then applied to the flow corrected data, based on linear regression with the representative ECs. The Kendall test was run on the flow corrected data, both with and without climate correction, with the results summarised on Fig. 11. Based on flow correction alone, there was a falling trend, with the null hypothesis of no trend being only 1% probable, but when the climate correction was added, no significant trend remained. A Spearman rank correlation between the site data and their paired monthly representative ECs gave a correlation coefficient of greater than 40%. This strongly suggests that the falling trend is climatic rather than related to the local landuse, which has been stable in recent times.

## 6 Conclusions

A new approach has been demonstrated to define the climate signal in ambient groundwater level and stream salinity records, based on the rate of change of the IPO. Although this research is in its early stages, it is evident that such broadscale representative hydrological and stream salinity time series are useful for defining the medium and long term responses of natural systems to both climate and landuse. In future the control series might be refined to be more representative in space and time. Further mathematical analysis of the control series is needed to relate them more effectively to

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climate processes. Also, back casting or extending the control time series, based on historical or simulated climatic variables, would help assess the full natural ranges of parameters, or the likely impacts of future climatic conditions.

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**Table 1.** Predictability of decadal flow trends in the Burdekin River, based on behaviour of the IPO as illustrated in Fig. 1.

Decade	Behaviour of the IPO	Expected Burdekin flow regime.
1920–1930	Rising, quite rapidly in mid years	At least median, above in mid years
1930–1940	Fall then rise, but fairly static	Usually low, particularly in early years
1940–1950	Falling rapidly	Above median
1950–1960	Rising, rapidly in later years	Above median, particularly in later years
1960–1970	Stationary	Very low
1970–1980	Slow fall early, then very rapid rise	Low initially, then well above median
1980–1990	Slowing rise then slow fall	Dropping and remaining below median
1990–2000	Slow rise followed by rapid fall	Around median, driest in mid years

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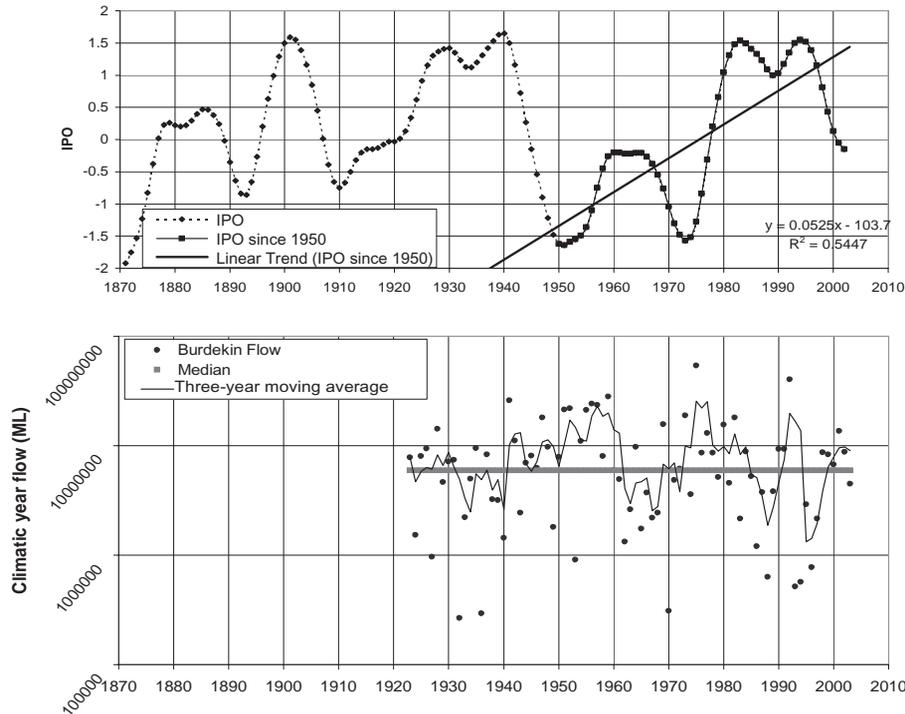
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**Fig. 1.** Comparison of the long term IPO and recorded annual flows in the Burdekin River. The tendency for the IPO to go through a generally rising and falling phase on about a 50 year cycle is observed, as well as the tendency for the river flow to be above the median when the IPO is changing rapidly, particularly when the change is positive.

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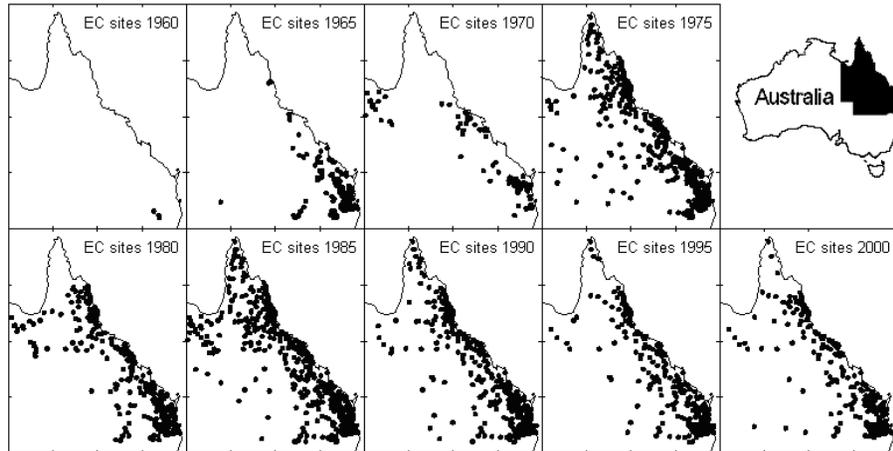
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**Fig. 2.** Distribution of EC samples stored in the NRME database, representing Queensland over time.

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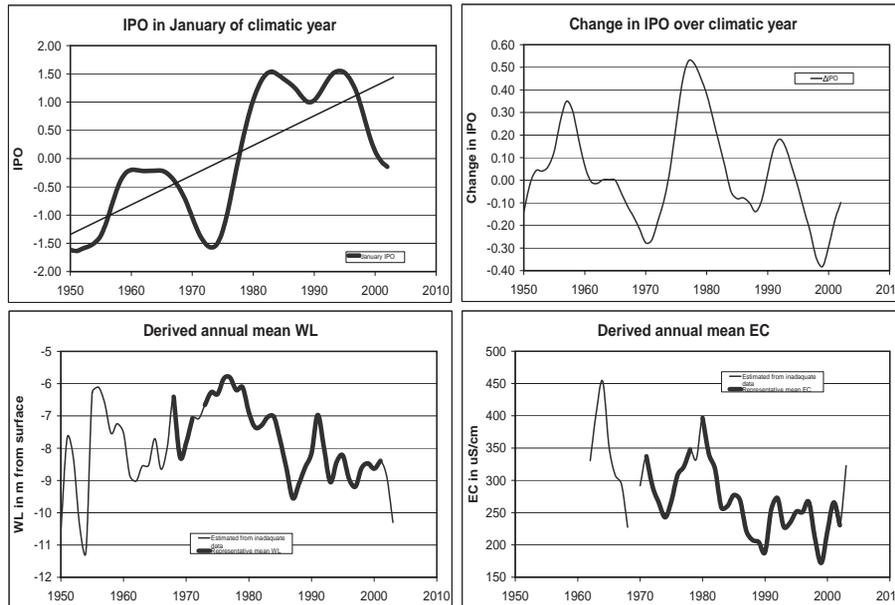
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**Fig. 3.** Climatic variables IPO and  $\Delta$ IPO, and State-wide representative water tables and stream salinity since 1950, as used for the modelling.

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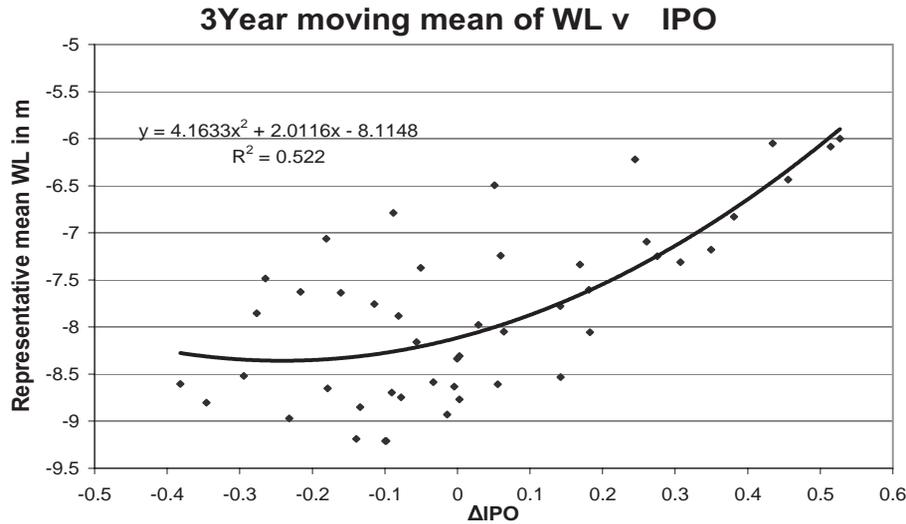


Fig. 4. Groundwater levels modelled against  $\Delta IPO$  using nonlinear (quadratic) regression.

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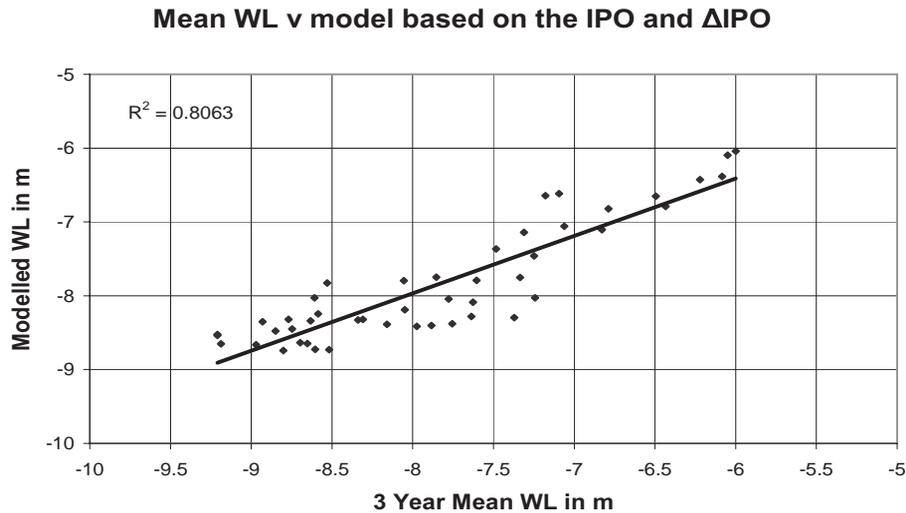
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**Fig. 5.** Three-year moving mean groundwater levels compared with the model based on the IPO and its rate of change,  $\Delta IPO$ .

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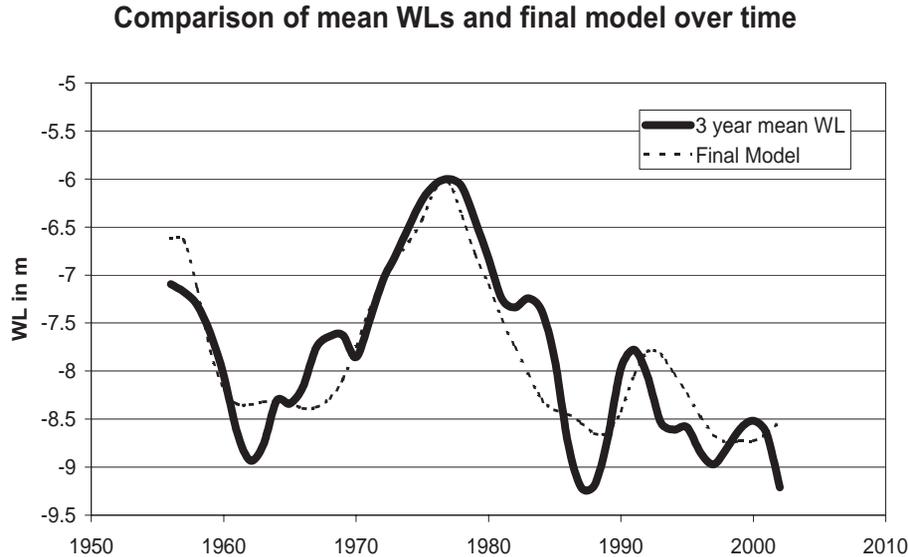
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**Fig. 6.** Three year moving mean groundwater levels compared with the model over time verifying a broad scale climatic component in the series.

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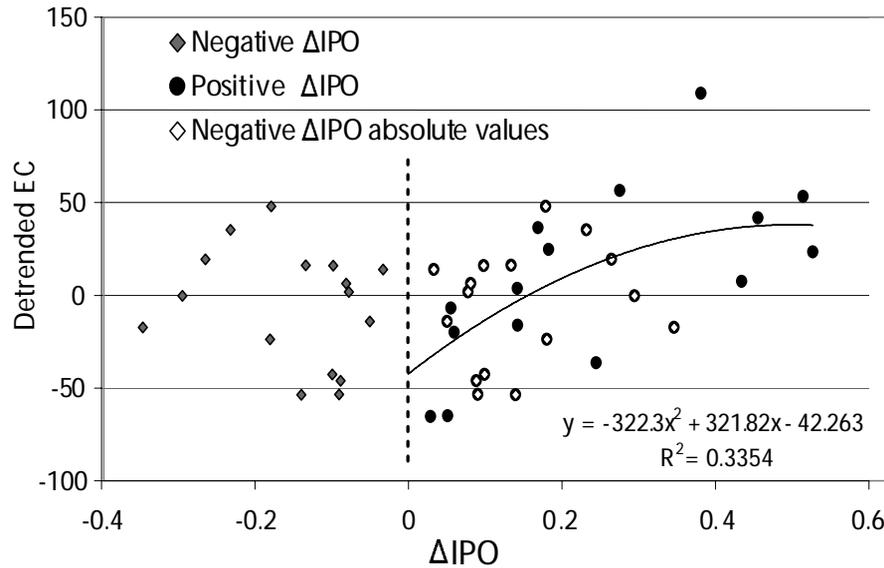
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**Fig. 7.** EC residuals modelled against the absolute value of  $\Delta\text{IPO}$  using nonlinear (quadratic) regression. The 3 year moving mean ECs were first detrended using the linear trend in the IPO since 1950.

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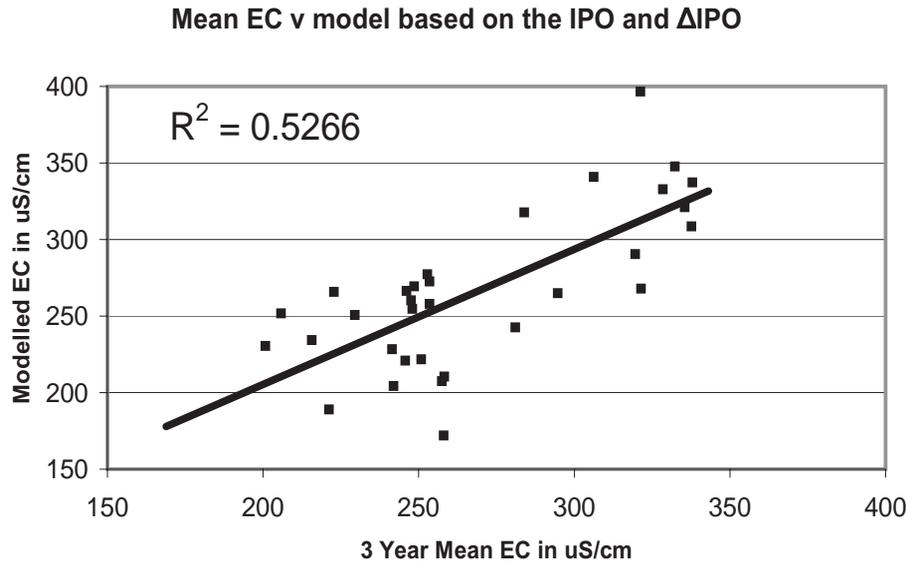
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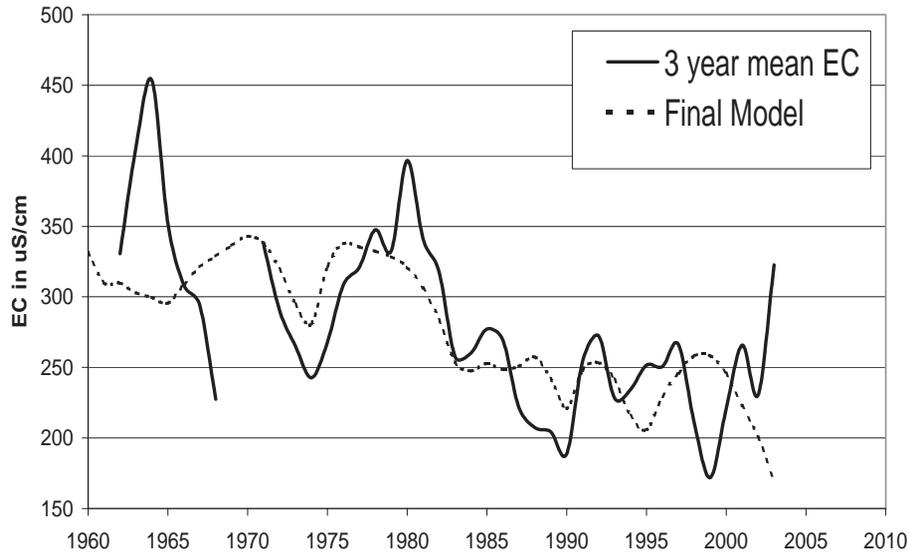


**Fig. 8.** Three year moving mean ECs compared with the model based on the IPO and its rate of change,  $\Delta IPO$ .

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**Fig. 9.** Three year moving mean EC levels compared with the model over time verifying a broad scale climatic component in the series.

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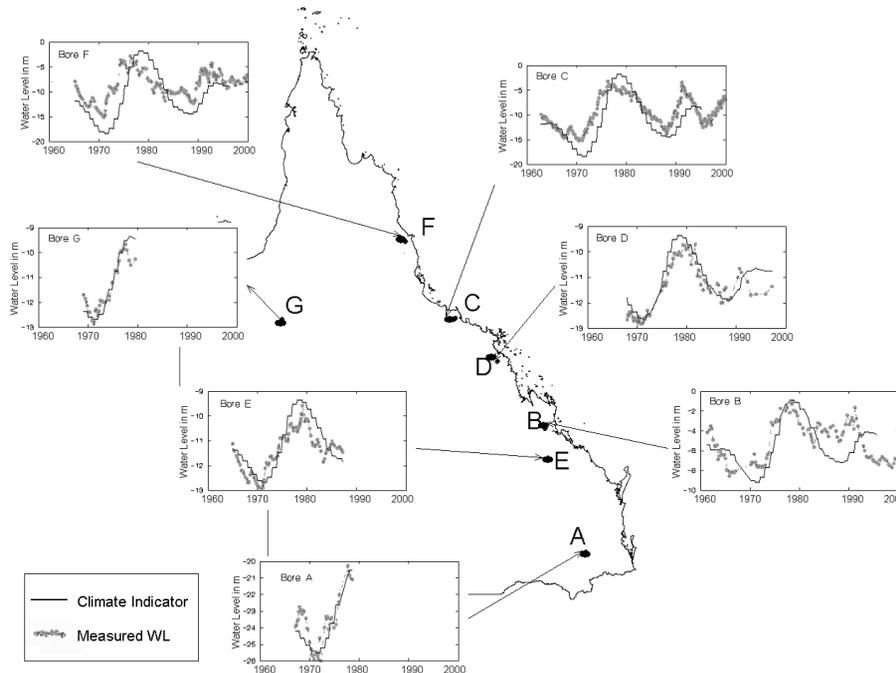
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**Fig. 10.** Selection of bore water level records showing influence of the climatic indicator over a wide range of Queensland.

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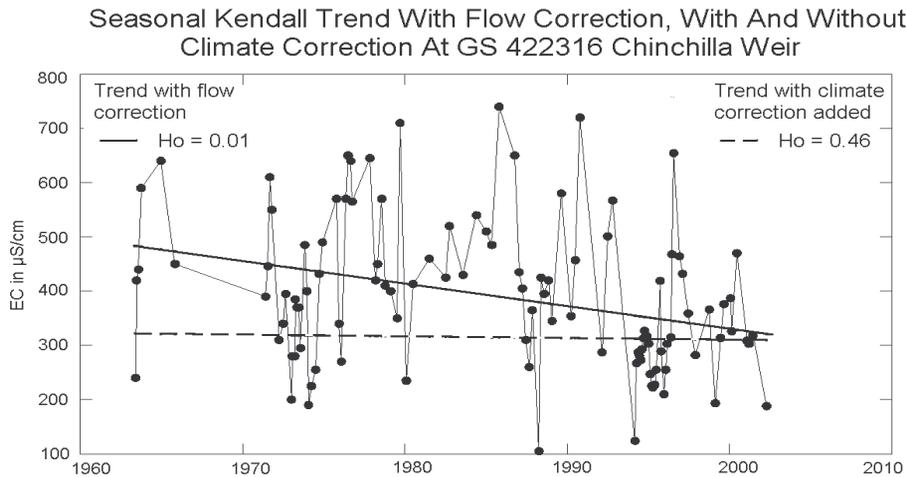
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**Fig. 11.** An example of a stream with an apparent falling EC trend, but with a general pattern that resembled the derived climate indicator time series. When the effect of climate was removed by linear regression with the climate indicator, no significant trend remained.

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