

# Numerical approximation of Backward Stochastic Differential Equations with Jumps

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**Abstract.** In this paper we propose a numerical method to approximate the solution of a Backward Stochastic Differential Equations with Jumps (BSDEJ). This method is based on the construction of a discrete BSDEJ driven by a complete system of three orthogonal discrete time-space martingales, the first a random walk converging to a Brownian motion; the second, another random walk, independent of the first one, converging to a Poisson process. The solution of this discrete BSDEJ is shown to weakly converge to the solution of the continuous time BSDEJ. An application to partial integro-differential equations is given.

**Keywords:** Backward SDEs with jumps, Skorokhod topology, Poisson Process, Monte Carlo method.

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# 1 Introduction

Consider a stochastic process  $\{Y_t, Z_t, U_t: 0 \leq t \leq T\}$  that is the solution of a Backward Stochastic Differential Equation with Jumps (in short BSDEJ) of the form

$$Y_t = \xi + \int_t^T f(s, Y_s, Z_s, U_s) ds - \int_t^T Z_s dB_s - \int_{(t,T] \times \mathbb{R}} U_s(x) \tilde{N}(ds, dx), \quad 0 \leq t \leq T. \quad (1.1)$$

Here  $\{B_t: 0 \leq t \leq T\}$  is a one dimensional standard Brownian motion;  $\tilde{N}(dt, dx) = N(dt, dx) - \nu(dt, dx)$  is a compensated Poisson random measure defined in  $[0, T] \times (\mathbb{R} \setminus \{0\})$ . Both processes are defined on a stochastic basis  $\mathcal{B} = (\Omega, \mathcal{F}_T, \mathbb{F} = \{\mathcal{F}_t\}_{0 \leq t \leq T}, \mathbb{P})$ , and, as usual in this framework, we assume that they are independent. The *terminal condition*  $\xi$  is a  $\mathcal{F}_T$ -measurable random variable in  $L^q(\mathbb{P})$ ,  $q > 2$  (see condition **(B)** in section 2); the *coefficient*  $f$  is a continuous, bounded and Lipschitz function  $f: [0, T] \times \mathbb{R}^3 \rightarrow \mathbb{R}$  (see condition **(A)** in Section 2). The problem of solving a BSDEJ given the terminal condition  $\xi$ , the coefficient  $f$ , and the driving processes  $B$  and  $\tilde{N}$ , defined on a stochastic basis  $\mathcal{B}$ , is to find three adapted processes  $\{Y_t, Z_t, U_t: 0 \leq t \leq T\}$  such that (1.1) holds.

In this paper — once the existence and uniqueness of the solution of a BSDEJ as described above is known — we propose a numerical method that approximates the solution of the equation (1.1) in the case when  $\tilde{N}$  is a compensated Poisson process.

Nonlinear backward stochastic differential equations of the form

$$Y_t = \xi + \int_t^T f(s, Y_s, Z_s) ds - \int_t^T Z_s dW_s, \quad 0 \leq t \leq T,$$

were first introduced by E. Pardoux and S. Peng [15] in 1990. Under some Lipschitz conditions on the generator  $f$ , the authors stated the first existence and uniqueness result. Later on, the same authors in 1992 developed the BSDE theory relaxing the hypotheses that ensure the existence and uniqueness on this type of equations in [16]. Many subsequent efforts have been made in order to relax further the assumptions on the coefficient  $f(s, y, z)$ , and many applications in mathematical finance have been proposed.

The backward stochastic differential equations with jumps theory begins with an existence result obtained by S. Tang and X. Li [12]. The authors

stated such a theorem when the generator satisfies some Lipschitz condition. Another relevant contribution on BSDEJ is the paper by R. Situ [20]. In [2], G. Barles, R. Buckdahn and E. Pardoux consider a BSDEJ when the driving noises are a Brownian motion and an independent Poisson random measure. They show the existence and uniqueness of the solution, and in addition, they establish a link with a partial integro-differential equation (in short PIDE).

A relevant problem in the theory of BSDEs is to propose implementable numerical methods to approximate the solution of such equations. Several efforts have been made in this direction as well. For example, in the Markovian case, J. Douglas, J. Ma and P. Protter [9] proposed a numerical method for a class of forward-backward SDEs, based on a four step scheme developed by J. Ma, P. Protter and J. Yong [14]. On the other hand, D. Chevance [6] proposed a numerical method for BSDEs. In [21], J. Zhang proposed a numerical scheme for a class of backward stochastic differential equations with possible path-dependent terminal values. See also [4, 11], among others, where numerical methods for decoupled forward-backward differential equations are proposed, and [1, 3, 13] for backward differential equations.

In the present work we propose to approximate the solution of a BSDEJ driven by a Brownian Motion and an independent compensated Poisson process, through the solution of a discrete backward equation, following the approach proposed for BSDE by Briand, Delyon and Mémin in [5]. The algorithm to compute this approximation is simple.

However, the rate of convergence is difficult to establish. The difficulty comes from the representation theorem. Not surprisingly, studying closely this term requires sophisticated tools such as Mallavian calculus, on which the work of B. Bouchard and R. Elie relies [3]. Here, we prefer to use relatively simple tools.

The rest of the paper is organized as follows. In section 2 we present the problem, propose an approximation process and a simple algorithm to compute it, and present the main convergence result of the paper. Section 3 is devoted to the proof of the previous main result, in the adequate topology. In section 4 we present an application of the main result to a decoupled system of a stochastic differential equation and a backward stochastic differential equation, resulting in a numerical approximation of the solution of an associated partial integro-differential equation.

## 2 Main result

In order to propose an implementable numerical scheme we consider that the Poisson random measure is simply generated by the jumps of a Poisson process. For the sake of simplicity of notation we work in the time interval  $[0, 1]$ . We then consider a Poisson process  $\{N_t : 0 \leq t \leq 1\}$  with intensity  $\lambda$  and jump epochs  $\{\tau_k : k = 0, 1, \dots\}$ . The random measure is then

$$\tilde{N}(dt, dx) = \sum_{k=1}^{N_1} \delta_{(\tau_k, 1)}(dt, dx) - \lambda dt \delta_1(dx),$$

where  $\delta_a$  denotes the Dirac delta at the point  $a$ . We also denote  $\tilde{N}_t = N_t - \lambda t$ . As a consequence, the unknown function  $U_t(x)$  that in principle depends on the jump magnitude  $x$  becomes  $U_t = U_t(1)$ , and our BSDEJ in (1.1) becomes

$$\begin{aligned} Y_t &= \xi + \int_t^1 f(s, Y_s, Z_s, U_s) ds - \int_t^1 Z_s dB_s - \sum_{i=N_t+1}^{N_1} U_{T_i} + \lambda \int_t^1 U_s ds, \\ &= \xi + \int_t^1 f(s, Y_s, Z_s, U_s) ds - \int_t^1 Z_s dB_s - \int_{(t,1]} U_s d\tilde{N}_s, \end{aligned} \quad (2.2)$$

for  $0 \leq t \leq 1$ . In order to ensure existence and uniqueness of the solution of this equation we consider the following conditions on the coefficient:

- (A) There exists  $K \geq 0$  and a bounded, non-decreasing continuous function  $\Lambda$  with  $\Lambda(0) = 0$  such that

$$\begin{aligned} &|f(s_1, y_1, z_1, u_1) - f(s_2, y_2, z_2, u_2)| \\ &\leq \Lambda(s_2 - s_1) + K(|y_1 - y_2| + |z_1 - z_2| + |u_1 - u_2|). \end{aligned}$$

In what respects the terminal condition, we assume:

- (B) The random variable  $\xi$  is  $\mathcal{F}_T$ -measurable and  $\mathbb{E}[|\xi|^q] < \infty$  for some  $q > 2$ .

### 2.1 Discrete time BSDE with jumps

We propose to approximate the solution of the BSDEJ in (2.2) by the solution of a *discrete* backward stochastic differential equation with jumps in

a discrete stochastic basis with a filtration generated by two independent, centered random walks. In order to obtain the representation property a third martingale is considered. The convergence of this approximation relies on the fact that the first random walk converges to the driving Brownian motion, and the second to the driving compensated Poisson process. Let us define these two random walks.

For  $n \in \mathbb{N}$  we introduce the first random walk  $\{W_k^n : k = 0, \dots, n\}$  by

$$W_0^n = 0, \quad W_k^n = \frac{1}{\sqrt{n}} \sum_{i=1}^k \epsilon_i^n \quad (k = 1, \dots, n), \quad (2.3)$$

where  $\epsilon_1^n, \dots, \epsilon_n^n$  are independent symmetric Bernoulli random variables:

$$\mathbb{P}(\epsilon_k^n = 1) = \mathbb{P}(\epsilon_k^n = -1) = 1/2, \quad (k = 1, \dots, n).$$

The second random walk  $\{\tilde{N}_k^n : k = 0, \dots, n\}$  is non symmetric, and given by

$$\tilde{N}_0^n = 0, \quad \tilde{N}_k^n = \sum_{i=1}^k \eta_i^n \quad (k = 1, \dots, n), \quad (2.4)$$

where  $\eta_1^n, \dots, \eta_n^n$  are independent and identically distributed random variables with probabilities, for each  $k$ , given by

$$\mathbb{P}(\eta_k^n = \kappa_n - 1) = 1 - \mathbb{P}(\eta_k^n = \kappa_n) = \kappa_n \quad (k = 1, \dots, n), \quad (2.5)$$

where  $\kappa_n = e^{-\lambda/n}$ . We assume that both sequences  $\epsilon_1^n, \dots, \epsilon_n^n$  and  $\eta_1^n, \dots, \eta_n^n$  are defined on the original probability space  $(\Omega, \mathcal{F}, \mathbb{P})$  (that can be enlarged if necessary), and that they are mutually independent. The (discrete) filtration in the probability space is  $\mathbb{F}^n = \{\mathcal{F}_k^n : k = 0, \dots, n\}$  with  $\mathcal{F}_0^n = \{\Omega, \emptyset\}$  and  $\mathcal{F}_k^n = \sigma\{\epsilon_1^n, \dots, \epsilon_k^n, \eta_1^n, \dots, \eta_k^n\}$  for  $k = 1, \dots, n$ .

In this discrete stochastic basis, given an  $\mathcal{F}_{k+1}^n$ -measurable random variable  $y_{k+1}$ , to represent the martingale difference  $m_{k+1} := y_{k+1} - \mathbb{E}(y_{k+1} | \mathcal{F}_k^n)$  we need a set of three strongly orthogonal martingales. This is a motivation to introduce a third martingale increments sequence  $\{\mu_k^n = \epsilon_k^n \eta_k^n : k = 0, \dots, n\}$ . In this context there exist unique  $\mathcal{F}_k^n$ -measurable random variables  $z_k, u_k, v_k$  such that

$$m_{k+1} = y_{k+1} - \mathbb{E}(y_{k+1} | \mathcal{F}_k^n) = \frac{1}{\sqrt{n}} z_k \epsilon_{k+1}^n + u_k \eta_{k+1}^n + v_k \mu_{k+1}^n, \quad (2.6)$$

that can be computed as

$$z_k = \mathbb{E}(\sqrt{n}y_{k+1}\epsilon_{k+1}^n \mid \mathcal{F}_k^n) = \sqrt{n} \mathbb{E}(y_{k+1}\epsilon_{k+1}^n \mid \mathcal{F}_k^n), \quad (2.7)$$

$$u_k = \frac{\mathbb{E}(y_{k+1}\eta_{k+1}^n \mid \mathcal{F}_k^n)}{\mathbb{E}((\eta_{k+1}^n)^2 \mid \mathcal{F}_k^n)} = \frac{1}{\kappa_n(1 - \kappa_n)} \mathbb{E}(y_{k+1}\eta_{k+1}^n \mid \mathcal{F}_k^n), \quad (2.8)$$

$$v_k = \frac{\mathbb{E}(y_{k+1}\mu_{k+1}^n \mid \mathcal{F}_k^n)}{\mathbb{E}((\mu_{k+1}^n)^2 \mid \mathcal{F}_k^n)} = \frac{1}{\kappa_n(1 - \kappa_n)} \mathbb{E}(y_{k+1}\mu_{k+1}^n \mid \mathcal{F}_k^n).$$

(Observe that the martingales are orthogonal but not orthonormal, hence the normalization.) With this information we proceed to formulate the discrete BSDEJ. Consider then a square integrable and  $\mathcal{F}_n^n$ -measurable terminal condition  $\xi^n$ , the same coefficient  $f$  as in Eq (2.2), and denote  $h = 1/n$ . Consider the discrete time backward stochastic differential equation with terminal condition  $y_n^n := y_{t_n}^n = \xi^n$ , and for  $k = n - 1, \dots, 0$  given by

$$y_k^n = \xi^n + \sum_{i=k}^{n-1} hf(t_i^n, y_i^n, z_i^n, u_i^n) - \sum_{i=k}^{n-1} \left( \sqrt{h}z_i^n \epsilon_{i+1}^n + u_i^n \eta_{i+1}^n + v_i^n \mu_{i+1}^n \right), \quad (2.9)$$

where  $y_i^n := y_{t_i}^n$ , and  $t_i^n = i/n$  ( $i = 0, \dots, n$ ). A solution to (2.9) is an  $\mathbb{F}^n$ -adapted sequence  $\{y_k^n, z_k^n, u_k^n, v_k^n : k = 0, \dots, n\}$  such that  $y_n^n = \xi^n$  and (2.9) holds.

Equation (2.9) is equivalent to a backward recursive system, beginning by  $y_n^n = \xi^n$ , followed by

$$\begin{aligned} y_k^n &= y_{k+1}^n + hf(t_k^n, y_k^n, z_k^n, u_k^n) - \sqrt{h}z_k^n \epsilon_{k+1}^n - u_k^n \eta_{k+1}^n - v_k^n \mu_{k+1}^n \\ &= y_{k+1}^n + hf(t_k^n, y_k^n, z_k^n, u_k^n) - m_{k+1}^n, \end{aligned} \quad (2.10)$$

for  $k = n - 1, \dots, 0$ . In view of the representation property (2.6), this last equation (2.10) is equivalent to

$$y_k^n = \mathbb{E}(y_{k+1}^n \mid \mathcal{F}_k^n) + hf(t_k^n, y_k^n, z_k^n, u_k^n). \quad (2.11)$$

The solution can be computed by the following simple algorithm:

- (I) Set  $(y_n^n, z_n^n, u_n^n, v_n^n) = (\xi^n, 0, 0, 0)$ .
- (II) For  $k$  from  $n - 1$  down to  $0$ , compute  $\mathbb{E}(y_{k+1}^n \mid \mathcal{F}_k^n)$ . Compute  $z_k^n$  and  $u_k^n$  as in (2.7) and (2.8) and solve (2.11) to find  $y_k^n$ , using a fixed point principle.

This algorithm shows the existence and uniqueness of the solution of the discrete BSDEJ defined in (2.9), as it is always possible to solve it using the fixed point principle for  $n$  large enough, such that  $Kh = K/n < 1$ . Bounds on the solution are given below in Section 3.1.

*Remark 1* (Computing conditional expectations). Our scheme is very simple to implement. However, as for any scheme that solves BSDEs numerically, it requires to compute conditional expectations. Various methods have then been proposed. In our case, for a function  $\Phi: \mathbb{R}^{2k+2} \rightarrow \mathbb{R}$  we use the formula

$$\begin{aligned} \mathbb{E}[\Phi(\epsilon_1^n, \dots, \epsilon_{k+1}^n, \eta_1^n, \dots, \eta_{k+1}^n) \mid \mathcal{F}_k^n] \\ &= \frac{\kappa_n}{2} \Phi(\epsilon_1^n, \dots, \epsilon_k^n, 1, \eta_1^n, \dots, \eta_k^n, \kappa_n - 1) \\ &+ \frac{\kappa_n}{2} \Phi(\epsilon_1^n, \dots, \epsilon_k^n, -1, \eta_1^n, \dots, \eta_k^n, \kappa_n - 1) \\ &+ \frac{1 - \kappa_n}{2} \Phi(\epsilon_1^n, \dots, \epsilon_k^n, 1, \eta_1^n, \dots, \eta_k^n, \kappa_n) \\ &+ \frac{1 - \kappa_n}{2} \Phi(\epsilon_1^n, \dots, \epsilon_k^n, -1, \eta_1^n, \dots, \eta_k^n, \kappa_n). \end{aligned} \quad (2.12)$$

*Remark 2.* In principle our proposal can be implemented for a  $d$ -dimensional Brownian motion and a Poisson process with a jump measure of the form  $\Pi(dz) = \sum_{i=1, \dots, \ell} \pi_i \mu_{x_i}(dz)$ . In this case, there exist a set of  $2^{d+\ell} - 1$  martingales with the representation property (2.6). It should be taken into account that the computation of conditional expectation in (II) involves the sum of  $2^{d+\ell}$  terms.

## 2.2 The main convergence result

Consider the continuous time version  $\{Y_t^n, Z_t^n, U_t^n, V_t^n: 0 \leq t \leq 1\}$  of the solution  $\{y_i^n, z_i^n, u_i^n, v_i^n: i = 0, \dots, n\}$  of the discrete equation (2.10) defined by

$$Y_t^n = y_{[tn]}^n, \quad Z_t^n = z_{[tn]}^n, \quad U_t^n = u_{[tn]}^n, \quad V_t^n = v_{[tn]}^n, \quad (2.13)$$

for  $t \in [0, 1]$ . Note that  $Y_t^n, Z_t^n, U_t^n$  and  $V_t^n$  are measurable w.r.t. the  $\sigma$ -algebra  $\mathcal{F}_{[nt]}^n$  when  $t \in [0, 1]$ . The discrete BSDEJ in (2.9), denoting  $c_n(t) = [nt]/n$  ( $0 \leq t \leq 1$ ), can be written as  $Y_1^n = \xi^n$  (the square integrable and

$\mathcal{F}_t^n$ -measurable terminal condition), and

$$Y_t^n = \xi^n + \int_{(t,1]} f(c_n(s), Y_{s-}^n, Z_{s-}^n, U_{s-}^n) dc_n(s) - \int_{(t,1]} Z_{s-}^n dW_s^n - \int_{(t,1]} U_{s-}^n d\tilde{N}_s^n - \int_{(t,1]} V_{s-}^n d\tilde{M}_s^n$$

for  $t \in [0, 1]$ , and  $\tilde{M}_s^n = W_s^n \times \tilde{N}_s^n$ . With our notations (2.10) becomes

$$Y_{t_i^n}^n = Y_{t_{i+1}^n}^n + \frac{1}{n} f(t_i^n, Y_{t_i^n}^n, Z_{t_i^n}^n, U_{t_i^n}^n) - \frac{1}{\sqrt{n}} Z_{t_i^n}^n \epsilon_{i+1}^n - U_{t_i^n}^n \eta_{i+1}^n - V_{t_i^n}^n \mu_{i+1}^n. \quad (2.14)$$

In Section 2.1 and Remark 1, we gave a numerical scheme to compute the solution  $\{Y_{t_i^n}^n, Z_{t_i^n}^n, U_{t_i^n}^n : 0 \leq t \leq 1\}$ .

We finally state the convergence assumption on the terminal conditions of the approximating equations:

**(H)** For some  $q > 2$ ,  $\sup_{n \in \mathbb{N}} \mathbb{E}[|\xi^n|^q] + \mathbb{E}[|\xi|^q]$  and  $\mathbb{E}[|\xi^n - \xi|^2] \xrightarrow{n \rightarrow \infty} 0$ .

**Theorem 1.** *Under the assumptions (A), (B) and (H), the set of processes  $(\xi^n, Y^n, \int_0^\cdot Z_s^n ds, \int_0^\cdot U_s^n ds)$  converges in the  $J_1$ -Skorokhod topology, and in probability, towards the solution  $(\xi, Y, \int_0^\cdot Z_s ds, \int_0^\cdot U_s ds)$  of the BSDEJ (1.1).*

### 3 Proofs

Our theorem is inspired in the main result of [5]. The proof follows, when possible, the main steps of the proof of this result. The main difference appears due to the fact that the underlying representation theorem for the simple symmetric Bernoulli random walk does not take place in our case, being necessary to consider a complete system of orthogonal martingales in order to have this representation property, as we have seen in equation (2.10). The idea is then to consider for both the discrete and the continuous time equations the approximations provided by the Picard's method.

In the continuous case denote  $Y^{\infty,0} = Z^{\infty,0} = U^{\infty,0} = 0$  and define  $\{Y^{\infty,p+1}, Z^{\infty,p+1}, U^{\infty,p+1} : 0 \leq t \leq 1\}$  inductively as the solution of the backward differential equation

$$Y_t^{\infty,p+1} = \xi + \int_t^1 f(s, Y_s^{\infty,p}, Z_s^{\infty,p}, U_s^{\infty,p}) ds - \int_t^1 Z_s^{\infty,p+1} dB_s - \int_{(t,1]} U_s^{\infty,p+1} d\tilde{N}_s.$$

In the discrete case, given  $n$  denote  $y_k^{n,0} = z_k^{n,0} = u_k^{n,0} = v_k^{n,0} = 0$  for  $k = 0, \dots, n$  and define  $\{y_k^{n,p+1}, z_k^{n,p+1}, u_k^{n,p+1}, v_k^{n,p+1} : k = 0, \dots, n\}$  inductively as the solution of the backward difference equation with terminal condition  $y_n^{n,p+1} = \xi^n$  and backwards recursion defined by

$$\begin{aligned} y_k^{n,p+1} &= y_{k+1}^{n,p+1} + hf(t_k^n, y_k^{n,p}, z_k^{n,p}, u_k^{n,p}) - \sqrt{h} z_k^{n,p+1} \epsilon_{k+1}^n - u_k^{n,p+1} \eta_{k+1}^n \\ &\quad - v_k^{n,p+1} \mu_{k+1}^n \\ &= y_{k+1}^{n,p+1} + hf(t_k^n, y_k^{n,p}, z_k^{n,p}, u_k^{n,p}) - m_{k+1}^{n,p+1}. \end{aligned} \quad (3.15)$$

If we consider the continuous time versions of the discrete Picard approximations as defined in (2.13), our method of proof relies on the decompositions

$$\begin{aligned} Y^n - Y &= (Y^n - Y^{n,p}) + (Y^{n,p} - Y^{\infty,p}) + (Y^{\infty,p} - Y), \\ Z^n - Z &= (Z^n - Z^{n,p}) + (Z^{n,p} - Z^{\infty,p}) + (Z^{\infty,p} - Z), \\ U^n - U &= (U^n - U^{n,p}) + (U^{n,p} - U^{\infty,p}) + (U^{\infty,p} - U), \end{aligned}$$

where  $(Y^n, Z^n, U^n)$  (resp.  $(Y^{n,p}, Z^{n,p}, U^{n,p})$ ) are the càdlàg processes on  $[0, 1]$  associated to  $(y^n, z^n, u^n)$  (resp.  $(y^{n,p}, z^{n,p}, u^{n,p})$ ) as in (2.13).

We prove the convergence of the discrete solution as  $n \rightarrow \infty$  to the solution of (1.1), by proving the uniform convergence in the the Picard iteration principle, as well as the convergence of the approximation of this solution given by this iteration principle at each step when the time step is refined.

### 3.1 Convergence of the Picard's iteration procedure in the discrete case

With standard computations, we have that

$$\sup_{p>0} \mathbb{E} \left[ \sup_{t \in [0,1]} |Y_t^{\infty,p}|^2 + \int_0^1 |Z_s^{\infty,p}|^2 ds + \lambda \int_0^1 |U_s^{\infty,p}|^2 ds \right] < +\infty \quad (3.16)$$

and

$$\mathbb{E} \left[ \sup_{t \in [0,1]} |Y_t^{\infty,p} - Y_t|^2 + \int_0^1 |Z_s^{\infty,p} - Z_s|^2 ds + \lambda \int_0^1 |U_s^{\infty,p} - U_s|^2 ds \right] \xrightarrow{p \rightarrow \infty} 0.$$

We now present similar results for the discrete approximations.

The next lemma evaluates the convergence rate of the Picard approximation sequence  $(y^{n,p}, z^{n,p}, u^{n,p})$  to  $(y^n, z^n, u^n)$  in the discrete scheme, and shows this rate is uniform in the time step  $1/n$ . It also provides the convergence of the discrete auxiliary sequence  $(v_k^{n,p})$  to zero.

We denote  $(y, z, u) = \{(y_k, z_k, u_k) : k = 0, \dots, n\}$  (with or without superscript  $n$ ) and, for  $\gamma > 1$  introduce the norms

$$\begin{aligned} \|(y^n, z^n, u^n)\|_\gamma^2 &:= \mathbb{E} \left( \sup_{0 \leq k \leq n} \gamma^{k/n} |y_k^n|^2 \right) + \frac{1}{n} \mathbb{E} \left( \sum_{k=0}^{n-1} \gamma^{k/n} |z_k^n|^2 \right) \\ &\quad + \mathbb{E} \left( \kappa_n (1 - \kappa_n) \sum_{k=0}^{n-1} \gamma^{k/n} |u_k^n|^2 \right), \\ \|v^n\|_\gamma^2 &:= \mathbb{E} \left( \kappa_n (1 - \kappa_n) \sum_{k=0}^{n-1} \gamma^{k/n} |v_k^n|^2 \right). \end{aligned}$$

**Lemma 1.** *There exists  $\gamma > 1$  and  $n_0 \in \mathbb{N}$  such that for all  $n \geq n_0$  and  $p \in \mathbb{N}^*$ ,*

$$\begin{aligned} &\| (y^{n,p+1} - y^{n,p}, z^{n,p+1} - z^{n,p}, u^{n,p+1} - u^{n,p}) \|_\gamma^2 + \|v^{n,p+1} - v_p^n\|_\gamma^2 \\ &\leq \frac{1}{4} \| (y^{n,p} - y^{n,p-1}, z^{n,p} - z^{n,p-1}, u^{n,p} - u^{n,p-1}) \|_\gamma^2. \end{aligned}$$

*Proof.* We introduce some notation. As  $n$  remains fixed during the main part of the proof, it will be omitted in the notations whenever possible. We denote

$$\delta x_k^{p+1} := x_k^{n,p+1} - x_k^{n,p}$$

for a quantity  $x = y, z, u, v, m$ . With this notation observe that

$$\delta y_k^{p+1} = \delta y_{k+1}^{p+1} + \frac{1}{n} \delta f_k^p - \delta m_{k+1}^{p+1}, \quad (3.17)$$

where  $m$  is given in (2.10). We set

$$\delta f_k^p = f(t_k^n, y_k^{n,p}, z_k^{n,p}, u_k^{n,p}) - f(t_k^n, y_k^{n,p-1}, z_k^{n,p-1}, u_k^{n,p-1}).$$

Now, for some  $\beta > 1$  we develop the quantity

$$\beta^n (\delta y_n^{p+1})^2 - \beta^k (\delta y_k^{p+1})^2 = -\beta^k (\delta y_k^{p+1})^2$$

through a discrete (time dependent) Itô Formula (compare with [19, VII §9]):

$$\begin{aligned}
-\beta^k(\delta y_k^{p+1})^2 &= \sum_{i=k}^{n-1} (\beta^{i+1}(\delta y_{i+1}^{p+1})^2 - \beta^i(\delta y_i^{p+1})^2) \\
&= (\beta - 1) \sum_{i=k}^{n-1} \beta^i(\delta y_i^{p+1})^2 + \sum_{i=k}^{n-1} \beta^{i+1} ((\delta y_{i+1}^{p+1})^2 - (\delta y_i^{p+1})^2) \\
&= (\beta - 1) \sum_{i=k}^{n-1} \beta^i(\delta y_i^{p+1})^2 + 2\beta \sum_{i=k}^{n-1} \beta^i \delta y_i^{p+1} (\delta y_{i+1}^{p+1} - \delta y_i^{p+1}) \\
&\quad + \beta \sum_{i=k}^{n-1} \beta^i (\delta y_{i+1}^{p+1} - \delta y_i^{p+1})^2. \tag{3.18}
\end{aligned}$$

From (3.17) follows, that

$$(\delta y_{i+1}^{p+1} - \delta y_i^{p+1})^2 \geq \frac{1}{2} (\delta m_{i+1}^{p+1})^2 - \frac{1}{n^2} (\delta f_i^p)^2.$$

Changing signs, using the previous inequality and (3.17) again, from (3.18) we obtain

$$\begin{aligned}
\beta^k(\delta y_k^{p+1})^2 &+ \frac{\beta}{2} \sum_{i=k}^{n-1} \beta^i (\delta m_{i+1}^{p+1})^2 \\
&\leq (1 - \beta) \sum_{i=k}^{n-1} \beta^i (\delta y_i^{p+1})^2 + 2\beta \sum_{i=k}^{n-1} \beta^i \delta y_i^{p+1} \left( \frac{1}{n} \delta f_{i+1}^p - \delta m_{i+1}^{p+1} \right) \\
&\quad + \frac{\beta}{n^2} \sum_{i=k}^{n-1} \beta^i (\delta f_{i+1}^p)^2. \tag{3.19}
\end{aligned}$$

We now use the inequality, for  $\lambda > 0$ ,

$$(\delta y_i^{p+1}) \left( \frac{2\beta}{n} \delta f_{i+1}^p \right) \leq \lambda (\delta y_i^{p+1})^2 + \frac{2\beta^2}{\lambda n^2} (\delta f_{i+1}^p)^2$$

in the second term of (3.19) to obtain

$$\begin{aligned}
& \beta^k (\delta y_k^{p+1})^2 + \frac{\beta}{2} \sum_{i=k}^{n-1} \beta^i (\delta m_{i+1}^{p+1})^2 \\
& \leq (1 + \lambda - \beta) \sum_{i=k}^{n-1} \beta^i (\delta y_i^{p+1})^2 + \frac{\beta + 4\lambda^{-1}\beta^2}{n^2} \sum_{i=k}^{n-1} \beta^i (\delta f_{i+1}^p)^2 \\
& \quad - 2\beta \sum_{i=k}^{n-1} \beta^i \delta y_i^{p+1} \delta m_{i+1}^{p+1}.
\end{aligned}$$

We now assume that  $1 + \lambda - \beta < 0$  and denote  $B := \beta + 4\lambda^{-1}\beta^2$  to obtain

$$\begin{aligned}
& \beta^k (\delta y_k^{p+1})^2 + \frac{\beta}{2} \sum_{i=k}^{n-1} \beta^i (\delta m_{i+1}^{p+1})^2 \\
& \leq \frac{B}{n^2} \sum_{i=k}^{n-1} \beta^i (\delta f_{i+1}^p)^2 - 2\beta \sum_{i=k}^{n-1} \beta^i \delta y_i^{p+1} \delta m_{i+1}^{p+1}. \quad (3.20)
\end{aligned}$$

Formula (3.20) is our first main inequality. From it we obtain the following two results. First, as the last summand is a martingale, taking expectations with  $k = 0$ , we obtain

$$\frac{\beta}{2} \mathbb{E} \sum_{i=k}^{n-1} \beta^i (\delta m_{i+1}^{p+1})^2 \leq \frac{B}{n^2} \mathbb{E} \sum_{i=k}^{n-1} \beta^i (\delta f_{i+1}^p)^2. \quad (3.21)$$

Second, taking supremum over  $k = 0, \dots, n$  we have

$$\sup_{0 \leq k \leq n} \beta^k (\delta y_k^{p+1})^2 \leq \frac{B}{n^2} \sum_{i=0}^{n-1} \beta^i (\delta f_{i+1}^p)^2 + 4\beta \sup_{0 \leq k \leq n} \left| \sum_{i=0}^k \beta^i \delta y_i^{p+1} \delta m_{i+1}^{p+1} \right|. \quad (3.22)$$

To obtain a convenient bound in the last term of (3.21), we use Davis (see [19, VII §3]) and afterwards Hölder inequalities. With  $F$  an universal constant,

we obtain

$$\begin{aligned}
4\beta\mathbb{E} \sup_{0 \leq k \leq n} \left| \sum_{i=0}^k \beta^i \delta y_i^{p+1} \delta m_{i+1}^{p+1} \right| &\leq F \mathbb{E} \sqrt{\sum_{i=0}^n \beta^{2i} (\delta y_i^{p+1})^2 (\delta m_{i+1}^{p+1})^2} \\
&\leq 4\beta F \mathbb{E} \sqrt{\sup_{0 \leq k \leq n} \beta^k (\delta y_k^{p+1})^2 \sum_{i=0}^n \beta^i (\delta m_{i+1}^{p+1})^2} \\
&\leq \frac{\beta}{2} \mathbb{E} \sup_{0 \leq k \leq n} \beta^k (\delta y_k^{p+1})^2 + 32\beta F^2 \mathbb{E} \sum_{i=0}^n \beta^i (\delta m_{i+1}^{p+1})^2.
\end{aligned}$$

Taking expectation in (3.22), using the previous result, and finally (3.21), we obtain that

$$(1 - \beta/2)\mathbb{E} \sup_{0 \leq k \leq n} \beta^k (\delta y_k^{p+1})^2 \leq \frac{B}{n^2} (1 + 64\beta F^2) \sum_{i=0}^{n-1} \beta^i (\delta f_{i+1}^p)^2. \quad (3.23)$$

Combining (3.21) and (3.23) we arrive to

$$\mathbb{E} \sup_{0 \leq k \leq n} \beta^k (\delta y_k^{p+1})^2 + \mathbb{E} \sum_{i=k}^{n-1} \beta^i (\delta m_{i+1}^{p+1})^2 \leq \frac{C}{n^2} \mathbb{E} \sum_{i=k}^{n-1} \beta^i (\delta f_{i+1}^p)^2, \quad (3.24)$$

with  $C = 2B(1 + 1/\beta + 64\beta F^2)/(1 - \beta/2)$ .

Using now the Lipschitz property **(A)** we see that there exists a constant  $\tilde{K}$  such that

$$|\delta f_i^p|^2 \leq \tilde{K} ((\delta y_i^p)^2 + (\delta z_i^p)^2 + n\kappa_n(1 - \kappa_n)(\delta u_i^p)^2). \quad (3.25)$$

Equations (3.24) and (3.25) give

$$\begin{aligned}
\mathbb{E} \sup_{0 \leq k \leq n} \beta^k (\delta y_k^{p+1})^2 + \mathbb{E} \sum_{i=k}^{n-1} \beta^i (\delta m_{i+1}^{p+1})^2 \\
\leq \frac{C\tilde{K}}{n} \left( \mathbb{E} \sup_{0 \leq k \leq n} \beta^k (\delta y_k^p)^2 + \frac{1}{n} \mathbb{E} \sum_{k=0}^{n-1} \beta^k (\delta z_k^p)^2 \right. \\
\left. + \kappa_n(1 - \kappa_n) \mathbb{E} \sum_{k=0}^{n-1} \beta^k (\delta u_k^p)^2 \right).
\end{aligned}$$

It remains to choose properly  $\beta$  and  $\lambda$  as a function of  $n$ . For some  $\gamma > 1$ , set  $\beta = \gamma^{1/n}$  and  $\lambda = \beta/n$ . The condition  $1 + \lambda < \beta$  is then equivalent to  $\gamma > (1 - 1/n)^{-n}$ . Thus, if  $\gamma > e$ , for  $n$  large enough this choice is suitable with our assumptions on  $\beta$  and  $\lambda$ , and  $C/n$  remains bounded as  $n \rightarrow \infty$ .

Computing  $\mathbb{E} (\delta m_{i+1}^{p+1})^2$ , we conclude the proof.  $\square$

Just like in the continuous case, we can use the Cauchy criterion and the preceding lemma to get the following result.

**Proposition 1.** *Following the notations of (2.14) and (3.15),*

$$\mathbb{E} \left( \sup_{0 \leq t \leq 1} |Y_t^{n,p} - Y_t^n|^2 + \int_0^1 |Z_t^{n,p} - Z_t^n|^2 dt + \lambda \int_0^1 |U_t^{n,p} - U_t^n|^2 dt \right)$$

converges to 0 uniformly in  $n$  as  $p \rightarrow \infty$ .

We may now state a global bound.

**Lemma 2.** *Under Hypotheses (A) and (H),*

$$\sup_{n \in \mathbb{N}} \sup_{p \in \mathbb{N}} \mathbb{E} \left[ \left| \frac{1}{n} \sum_{k=0}^{n-1} f(t_k^n, y_k^{n,p}, z_k^{n,p}, u_k^{n,p}) \right|^q \right] < +\infty.$$

In addition,

$$\sup_{p \in \mathbb{N}} \mathbb{E} \left[ \left| \int_0^1 f(s, Y_s^{\infty,p}, Z_s^{\infty,p}, U_s^{\infty,p}) ds \right|^q \right] < +\infty.$$

*Proof.* Here, we deal only with the discrete case, which is rather similar to the continuous case whose proof is a variation of the one given in [3, 10].

Again, we drop the superscript  $n$ . Here, we assume in a first time that the time horizon is  $T$  and  $h = T/n$ .

We set

$$\|(z^p, u^p)\| := \frac{T}{n} \sum_{k=1}^n (|z_k^p|^2 + |u_k^p|^2).$$

Since  $n\eta_k^2$  is bounded in  $n$  and  $k$ , we get from the Burkholder Davies-Gundy inequality that for some constants  $C_1$  and  $C_2$ ,

$$\mathbb{E}[\|(z^p, u^p)\|^{q/2}] \leq C_1 \mathbb{E} \left[ \left( \sum_{k=1}^n |m_k^p|^2 \right)^{q/2} \right] \leq C_2 \mathbb{E} \left[ \sup_{k=1, \dots, n} |m_k^p|^q \right].$$

On the other hand,

$$\mathbb{E} \left[ \sup_{\ell=1, \dots, n} |m_\ell^{p+1}|^q \right] \leq \mathbb{E} \left[ \sup_{\ell=1, \dots, n} \left| \mathbb{E} \left[ \xi + \frac{T}{n} \sum_{k=0}^{n-1} f(t_k, y_k^p, z_k^p, u_k^p) \middle| \mathcal{F}_\ell^n \right] \right|^q \right].$$

With Hypothesis **(A)**, for some constant  $C_3$  depending only on  $\Lambda(T)$  and  $K$ ,

$$\begin{aligned} \left| \frac{T}{n} \sum_{k=0}^{n-1} f(t_k, y_k^p, z_k^p, u_k^p) \right| &\leq \frac{T}{\sqrt{n}} \left( \sum_{k=0}^{n-1} f(t_k, y_k^p, z_k^p, u_k^p)^2 \right)^{1/2} \\ &\leq C_3 \frac{T}{\sqrt{n}} \left( n + \sum_{k=0}^{n-1} (|y_k^p|^2 + |z_k^p|^2 + |u_k^p|^2) \right)^{1/2} \\ &\leq C_3 T + C_3 T \sup_{k=0, \dots, n-1} |y_k^p| + C_3 \sqrt{T} \|(z^p, u^p)\|^{1/2}. \end{aligned}$$

With the Jensen inequality for the conditional expectation and the Doob inequality,

$$\mathbb{E} \left[ \sup_{k=1, \dots, n} |m_k^{p+1}|^q \right] \leq C_4 \mathbb{E}[|\xi|^q] + C_4 T^q \|y^p\|_{q, \star} + C_4 T + C_4 T^q \mathbb{E}[\|(z^p, u^p)\|^{q/2}]$$

with

$$\|y^p\|_{q, \star} := \mathbb{E} \left[ \sup_{k=0, \dots, n-1} |y_k^p|^q \right].$$

On the other hand, we have with similar computations, for some constant  $C_4$  depending only on  $C_3$  and  $q$ ,

$$\begin{aligned} \|y^{p+1}\|_{\star} &\leq \mathbb{E} \left[ \sup_{\ell=0, \dots, n-1} \left| \mathbb{E} \left[ \xi + \frac{T}{n} \sum_{k=\ell}^{n-1} f(t_k, y_k^p, z_k^p, u_k^p) \middle| \mathcal{F}_\ell^n \right] \right|^q \right] \\ &\leq C_4 \mathbb{E}[|\xi|^q] + C_4 T + C_4 T^q \|y^p\|_{q, \star} + C_4 T^q \mathbb{E}[\|(z^p, u^p)\|^{q/2}]. \end{aligned}$$

This proves that for  $C_5 = 2C_4$ ,

$$\begin{aligned} \|y^{p+1}\|_{q, \star} + \mathbb{E}[\|(z^{p+1}, u^{p+1})\|^{q/2}] \\ \leq C_5 \mathbb{E}[|\xi|^q] + C_5 T + C_5 T^q (\|y^p\|_{q, \star} + \mathbb{E}[\|(z^p, u^p)\|^{q/2}]). \end{aligned}$$

If  $T$  is small enough so that  $C_5 T < 1$ , then this proves that

$$\|y^p\|_{q, \star} + \mathbb{E}[\|(z^p, u^p)\|^{q/2}] \leq \frac{1}{1 - C_5 T} (1 + C_5 \mathbb{E}[|\xi|^q]).$$

Here, we have obtained a bound when  $T$  is small enough. Now, in order to consider the Picard scheme on the time interval  $[0, 1]$ , we may find for each  $n$  a time  $T_n$  such that  $C_5 T_n \leq k < 1$  and  $T_n = \ell(n)/n$  for some  $\ell$  for some fixed  $k$ , and solve recursively the Picard scheme on  $[T - T_n, T]$ ,  $[T - 2T_n, T - T_n]$ , ... using the terminal condition  $\xi^n$  and then  $y_{\ell(n)}^{n,p}, \dots$

We then obtain the desired bound.  $\square$

### 3.2 Approximation of Brownian motion and Poisson process

In order to establish convergence in probability, we consider that all the processes are defined on the same probability space.

**Lemma 3.** (I) Let us consider a compensated Poisson process  $\tilde{N}_t = N_t - \lambda t$ , where  $N$  is a Poisson process of intensity  $\lambda$ . For any integer  $n$ , one can find a family of realizations of independent random variables  $(\eta_i^n)_{i=1, \dots, n}$  such that (i) the distributions of the  $\eta_s^n$ 's are given by (2.5). (ii) For  $\tilde{N}_t^n = \sum_{i=1}^{\lfloor t/h \rfloor - 1} \eta_i^n$ , then  $\tilde{N}^n$  is a martingale and there exists a family  $(\phi^n)_{n \in \mathbb{N}}$  of one-to-one random time changes from  $[0, 1]$  to  $[0, 1]$  such that  $\sup_{t \in [0, 1]} |\phi^n(t) - t| \xrightarrow[n \rightarrow \infty]{} 0$  almost surely and  $\sup_{t \in [0, 1]} |\tilde{N}_t - \tilde{N}_{\phi^n(t)}^n| \xrightarrow[n \rightarrow \infty]{} 0$  in probability. In other words,  $\tilde{N}^n$  converges in probability to  $\tilde{N}$  in the  $J_1$ -Skorokhod topology.

(II) Let  $W$  be a Brownian motion. Then there exists a family of realizations independent random variables  $\epsilon_k^n$  such that  $\mathbb{P}(\epsilon_k^n = 1) = \mathbb{P}(\epsilon_k^n = -1) = 1/2$  and  $W^n = h \sum_{i=1}^{\lfloor t/h \rfloor - 1} \epsilon_i^n$  converges uniformly in probability to  $W$ .

(III) The couple  $(W^n, \tilde{N}^n)$  converges in the  $J_1$ -Skorokhod topology in probability to  $(W, \tilde{N})$ .

*Proof.* (I) Let  $(\Omega, \mathcal{F}, \mathbb{P})$  be the probability space on which  $N$  is defined. Denote by  $\{\tau_i\}_{i=1, \dots, \ell-1}$  the time jumps of  $N$ . To simplify the notations, we set  $\tau_0 = 0$  and  $\tau_\ell = T$ . Let  $A^n$  the event

$$A^n = \left\{ \begin{array}{l} \text{there is at most one jump of } N \\ \text{on any interval } [kh, h(k+1)) \text{ for } k = 0, \dots, n \end{array} \right\}.$$

We denote by  $\bar{A}^n$  the complementary event of  $A^n$ .

Let  $\phi^n(t, \omega)$  be the random piecewise linear function defined so that for  $i = 0, \dots, \ell$ ,  $\phi^n(\tau_i) = c_n(\tau_i)$  on  $A^n$  and by  $\phi^n(t) = t$  on  $\bar{A}^n$ .

It is easily checked that  $0 \leq t - \phi^n(t) \leq h$  for  $t \in [0, 1]$ .

Then for  $t \in [\tau_i, \tau_{i+1}]$ ,

$$\tilde{N}_{\phi^n(t)}^n - \tilde{N}_{c_n(\tau_i)}^n = (e^{-\lambda/n} - 1)(c_n(\phi_n(t)) - c_n(\tau_i)).$$

On the other hand,  $\tilde{N}_t - \tilde{N}_{\tau_i} = \lambda(t - \tau_i)$  and then there exists some constant  $K$  such that

$$|\tilde{N}_{\phi^n(t)}^n - \tilde{N}_{c_n(\tau_i)}^n - \tilde{N}_t + \tilde{N}_{\tau_i}| \leq Kh \quad (3.26)$$

for  $t \in [\tau_i, \tau_{i+1}]$ .

On the other hand, on  $A^n$ , for  $h$  small enough,

$$|\tilde{N}_{c_n(\tau_i)}^n - \tilde{N}_{c_n(\tau_i)-}^n - \tilde{N}_{\tau_i} - \tilde{N}_{\tau_i-}| = e^{-\lambda/n} - 1 \leq 2\lambda h. \quad (3.27)$$

Combining (3.26) and (3.27), one gets that on  $A_n$ ,

$$\sup_{t \in [0,1]} |\tilde{N}_{\phi^n(t)}^n - \tilde{N}_t| \leq (2 + K)\ell h.$$

In addition,  $\mathbb{E}[\ell] < +\infty$  so that

$$\mathbb{E} \left[ \sup_{t \in [0,1]} |\tilde{N}_{\phi^n(t)}^n - \tilde{N}_t|; A_n \right] \xrightarrow{n \rightarrow \infty} 0.$$

On  $\bar{A}^n$ , then  $\phi^n(t) = t$  and then

$$\tilde{N}_{ih} - \tilde{N}_{ih}^n = -\lambda ih - h(\kappa - 1)i + \sum_{j=0}^{i-1} (N_{(j+1)h} - N_j - 1) \mathbf{1}_{\{N_{(j+1)h} - N_j \geq 2\}}.$$

On the other hand, from the very definition of  $A^n$  and since  $N_{(i+1)h} - N_{ih}$  has the distribution of a Poisson random variable with intensity  $\lambda h$ ,

$$\mathbb{P} [\bar{A}^n] \leq \sum_{i=0}^{n-1} \mathbb{P} [|N_{(i+1)h} - N_{ih}| \geq 2] \leq n(1 - e^{-\lambda h} - \lambda h e^{-\lambda h}) \leq \frac{\lambda^2 T^2}{n}.$$

Thus, for any  $C > 0$ ,

$$\mathbb{P} \left[ \sup_{t \in [0,1]} |\tilde{N}_t - \tilde{N}_{\phi^n(t)}^n| > C \right] \leq \mathbb{P} [\bar{A}^n] + \frac{1}{C} \mathbb{E} \left[ \sup_{t \in [0,1]} |\tilde{N}_t - \tilde{N}_{\phi^n(t)}^n|; A^n \right].$$

Point (II) follows from the Donsker theorem, when one uses for example the Skorokhod embedding theorem to construct the  $\epsilon_k$ 's from the Brownian path [17] and (III) holds since  $W$  is continuous so that the 2-dimensional path  $(W^n, \tilde{N}^n)$  converges in the  $J_1$ -topology to  $(W, \tilde{N})$ .  $\square$

*Remark 3.* With a bit more work, it may be proved that  $\tilde{N}$  converges indeed to  $N$  in  $L^1(\mathbb{P})$ .

### 3.3 Convergence of martingales

Following the ideas given in [5, Theorem 3.1], consider a probability space  $(\Omega, \mathcal{F}, \mathbb{P})$  and the following processes defined on it:

- two independent random walks with independent increments, denoted by  $(W_t^n)_{0 \leq t \leq 1}$  and  $(\tilde{N}_t^n)_{0 \leq t \leq 1}$  respectively, and  $H^n = (W^n, \tilde{N}^n)$ ;
- a Brownian motion  $(W_t)_{0 \leq t \leq 1}$ , and a compensated Poisson process  $(\tilde{N}_t)_{0 \leq t \leq 1}$ , with  $H = (W, \tilde{N})$ .

We denote by  $(\mathcal{F}_t^n)$  (resp.  $(\mathcal{F}_t)$ ) the right continuous filtration generated by  $H^n$  (resp.  $H$ ).

Let us consider finally a sequence  $X^n$  of  $\mathcal{F}_1^n$ -measurable integrable random variables, and  $X$  an  $\mathcal{F}_1$ -measurable integrable random variables, together with the càdlàg martingales

$$M_t^n = \mathbb{E}(X^n | \mathcal{F}_t^n) \quad \text{and} \quad M_t = \mathbb{E}(X | \mathcal{F}_t). \quad (3.28)$$

We denote by  $[M^n, M^n]$  (resp.  $[M, M]$ ) the quadratic variation of  $M^n$  (resp.  $M$ ) and by  $[M^n, W^n]$ ,  $[M^n, \tilde{N}^n]$  (resp.  $[M, W]$ ,  $[M, \tilde{N}]$ ) the cross variation of  $M^n$  and  $W^n$  (resp.  $\tilde{N}^n$ ).

**Proposition 2.** *Let us consider the following assumptions:*

- (H1) *for each  $n$ ,  $W^n$  and  $\tilde{N}^n$  are two independent square integrable  $\mathcal{F}^n$ -martingales with independent increments;*
- (H2)  *$W^n \rightarrow W$  in probability for the topology of uniform convergence of càdlàg processes indexed by  $t \in [0, 1]$ ;*
- (H3)  *$\tilde{N}^n \rightarrow \tilde{N}$  in probability for the  $J_1$ -Skorokhod topology;*
- (H4)  $\mathbb{E}[X^2] + \sup_n \mathbb{E}[(X^n)^2] < \infty$ ;
- (H5)  $\mathbb{E}[|X^n - X|] \rightarrow 0$ .

Then, if conditions (H1)-(H5) are satisfied,

$$(H^n, M^n, [M^n, M^n], [M^n, W^n], [M^n, \tilde{N}^n]) \rightarrow (H, M, [M, M], [M, W], [M, \tilde{N}])$$

in probability for the  $J_1$ -Skorokhod topology.

*Proof.* The proof of this proposition follows directly from [5] Theorem 3.1, and [7] for the convergence of filtrations. We point out three remarks:

- First, since  $W^n \xrightarrow{\|\cdot\|_\infty} W$  and  $\tilde{N}^n \xrightarrow{J_1} \tilde{N}$  as  $n \rightarrow \infty$  in probability, then  $H^n \xrightarrow{J_1} H$ , as  $n \rightarrow \infty$ . So, we prove the convergence in probability for the  $J_1$ -Skorokhod topology instead the uniform convergence.
- Second, since  $H^n \xrightarrow{J_1} H$ , as  $n \rightarrow \infty$ , then we can get the convergence in probability for sequence of martingales  $M^n$  to  $M$  in the  $J_1$ -Skorokhod topology.
- The convergence of the brackets follows from the fact that

$$\sup_n \mathbb{E} \left[ \sup_{t \in [0,1]} (|M_t^n|^2 + |H_t^n|^2) \right] < +\infty,$$

which holds thanks to (H1) and (H4).  $\square$

*Remark 4.* Under the stronger assumption that

$$(H4') \text{ For some } \delta > 0, \sup_{n \in \mathbb{N}} \mathbb{E}[|X^n|^{2+\delta}] < +\infty,$$

then  $M^n$  converges to  $M$  in  $L^2(\mathbb{P})$  in the  $J_1$ -Skorokhod topology.

**Corollary 1.** *Let  $W, W^n, \tilde{N}$  and  $\tilde{N}^n, n \in \mathbb{N}$ , be the standard Brownian motion, the symmetric random walk defined in (2.3), a compensated Poisson Process and the random walk defined in (2.4). Let us consider, on the same space,  $X$  and  $X^n$  satisfying Hypothesis (H4) and*

$$(H5') \mathbb{E}[|X^n - X|^2] \xrightarrow{n \rightarrow \infty} 0.$$

*Then there exist three sequences  $(Z_t^n)_{0 \leq t \leq 1}$ ,  $(V_t^n)_{0 \leq t \leq 1}$  and  $(U_t^n)_{0 \leq t \leq 1}$  of  $\mathcal{F}^n$ -predictable processes, and two independent  $(Z_t)_{0 \leq t \leq 1}$  and  $(U_t)_{0 \leq t \leq 1}$   $\mathcal{F}$ -predictable processes such that*

$$\forall t \in [0, 1], \begin{cases} M_t^n = \mathbb{E}[X^n] + \int_0^t Z_{s-}^n dW_s^n + \int_0^t U_{s-}^n d\tilde{N}_s^n + \int_0^t V_{s-}^n d\tilde{M}_s^n \\ M_t = \mathbb{E}[X] + \int_0^t Z_{s-} dW_s + \int_0^t U_{s-} d\tilde{N}_s \end{cases}$$

with

$$\mathbb{E} \left[ \int_0^1 (Z_t^n - Z_t)^2 dt + \lambda \int_0^1 (U_t^n - U_t)^2 dt \right] \xrightarrow{n \rightarrow \infty} 0.$$

*Proof.* The first part is related to the predictable representation of  $\mathcal{F}^n$ -martingales in terms of stochastic integrals with respect to three independent random walks,  $W^n$ ,  $\tilde{N}^n$  and  $\tilde{M}^n$ . The increments of  $\tilde{M}^n$  may take up to four different values, which means that we need three orthogonal martingales to represent it (see [8]). It is then easily obtained that  $\tilde{M}^n$  is a martingale which is orthogonal to both  $\tilde{N}^n$  and  $W^n$ . This is why we introduce it. The predictable representation of  $M$  with respect to  $W$  and  $\tilde{N}$  is classical.

From the Doob inequality,

$$\mathbb{E}[|M^n, M^n|_1] \leq \mathbb{E}[|M_1^n|^2] \leq 2\mathbb{E}[|X^n|^2].$$

With (H4),

$$\sup_{n \in \mathbb{N}} \mathbb{E} \left[ \frac{1}{n} \int_0^1 (Z_{s-}^n)^2 dc_n(s) + \int_0^1 (\eta_{c_n(s)}^n)^2 (U_{s-}^n)^2 dc_n(s) + \frac{1}{n} \int_0^1 (\eta_{c_n(s)}^n)^2 (V_{s-}^n)^2 dc_n(s) \right] < +\infty. \quad (3.29)$$

Since  $\mathbb{E}[n(\eta_k^n)^2] \sim_{n \rightarrow \infty} \lambda$  and  $U^n$  is predictable with respect to  $\mathcal{F}^n$ , one easily get that

$$\limsup_{n \rightarrow \infty} \mathbb{E} \left[ \int_0^1 (Z_s^n)^2 ds + \int_0^1 \lambda (U_s^n)^2 ds \right] < +\infty. \quad (3.30)$$

Let  $(\mathcal{F}_t^W)_{t \geq 0}$  be the filtration generated by the Brownian motion. Since  $W$  and  $N$  are independent, for  $\bar{X} = X - \mathbb{E}[X]$ ,

$$\mathbb{E}[\bar{X} | \mathcal{F}_1^W] = \int_0^1 Z_s dW_s. \quad (3.31)$$

Let also  $\mathcal{G}$  be the  $\sigma$ -algebra generated by  $\epsilon_1, \dots, \epsilon_n$ . Hence, for  $\bar{X}^n = X^n - \mathbb{E}[X^n]$ ,

$$\mathbb{E}[\bar{X}^n | \mathcal{G}] = \int_0^1 Z_{s-}^n dW_s^n. \quad (3.32)$$

It follows that

$$\mathbb{E} \left[ \mathbb{E}[\bar{X}^n | \mathcal{G}]^2 \right] = \mathbb{E} \left[ \int_0^1 (Z_{s-}^n)^2 dc_n(s) \right].$$

Since the  $\epsilon_k$ 's are constructed from the trajectories of  $W$ , one has  $\mathcal{F} \subset \mathcal{F}_1^W$ . Hence

$$\mathbb{E} \left[ \mathbb{E} [\bar{X}^n | \mathcal{G}]^2 \right] \leq 2\mathbb{E} \left[ \mathbb{E} [\bar{X}^n - \bar{X} | \mathcal{G}]^2 \right] + 2\mathbb{E} \left[ \mathbb{E} [\bar{X} | \mathcal{G}]^2 \right].$$

With the Jensen inequality on conditional expectation and (H5'), one gets that

$$\mathbb{E} \left[ \mathbb{E} [\bar{X}^n - \bar{X} | \mathcal{G}]^2 \right] \xrightarrow{n \rightarrow \infty} 0$$

and

$$\mathbb{E} \left[ \mathbb{E} [\bar{X} | \mathcal{G}]^2 \right] \leq \mathbb{E} \left[ \mathbb{E} \left[ \mathbb{E} [\bar{X} | \mathcal{F}_1^W] | \mathcal{G} \right]^2 \right] \leq \mathbb{E} \left[ \mathbb{E} [\bar{X} | \mathcal{F}_1^W]^2 \right].$$

From (3.31) and (3.32), one gets that

$$\limsup_{n \rightarrow \infty} \mathbb{E} \left[ \int_0^1 (Z_s^n)^2 ds \right] \leq \mathbb{E} \left[ \int_0^1 Z_s^2 ds \right]. \quad (3.33)$$

Similar arguments prove that

$$\limsup_{n \rightarrow \infty} \mathbb{E} \left[ \int_0^1 (U_s^n)^2 ds \right] \leq \mathbb{E} \left[ \int_0^1 U_s^2 ds \right].$$

From Proposition 2,  $[M^n, W^n] \rightarrow [M, W]$  in probability for the  $J_1$ -Skorokhod topology, as well as  $[M^n, M^n]$ . Then

$$\sup_{0 \leq t \leq 1} \left| \int_0^{\psi^n(t)} Z_{s-}^n dc_n(s) - \int_0^t Z_s ds \right| \xrightarrow{n \rightarrow \infty} 0 \quad (3.34)$$

in probability, where  $\psi^n(t) \uparrow t$ . Then we get easily that

$$\sup_{0 \leq t \leq 1} \left| \int_0^t Z_s^n ds - \int_0^t Z_s ds \right| \xrightarrow{n \rightarrow \infty} 0 \quad (3.35)$$

in probability and with (3.30), in  $L^1(\mathbb{P})$ .

On the other hand,  $[M^n, \tilde{N}^n] \rightarrow [M, \tilde{N}]$  in probability for the  $J_1$ -Skorokhod topology. This implies that

$$\sup_{0 \leq t \leq 1} \left| \int_0^{\psi^n(t)} \eta_{c_n(s)}^2 U_{s-}^n dc_n(s) - \lambda \int_0^t U_s ds \right| \xrightarrow{n \rightarrow \infty} 0. \quad (3.36)$$

We can apply the same arguments used for (3.35), Burkholder-Davis-Gundy inequality to control the distance between  $\eta_k^2$  and  $\kappa_n(1 - \kappa_n)$ , and the fact that  $1 - \kappa_n \sim_{n \rightarrow \infty} \lambda/n$  to get

$$\sup_{0 \leq t \leq 1} \lambda \left| \int_0^t U_s^n ds - \int_0^t U_s ds \right| \xrightarrow{n \rightarrow \infty} 0 \quad (3.37)$$

in probability and in  $L^1(\mathbb{P})$ .

The second part relies on the following argument: let  $(g_n)_{n \in \mathbb{N} \cup \{\infty\}}$  be a sequence of functions on  $[0, 1] \times \Omega$  such that

$$\limsup_{n \rightarrow \infty} \mathbb{E} \left[ \int_0^1 (g_n(s, \omega))^2 ds \right] \leq \mathbb{E} \left[ \int_0^1 (g_\infty(s, \omega))^2 ds \right] < +\infty, \quad (3.38)$$

$$\text{and } \mathbb{E} \left[ \sup_{t \in [0, 1]} \left| \int_0^t (g_n(s, \cdot) - g_\infty(s, \cdot)) ds \right| \right] \xrightarrow{n \rightarrow \infty} 0. \quad (3.39)$$

For any given function  $h_\infty$  in  $L^2([0, 1] \times \Omega)$ , there exists a sequence of functions  $(h_n)_{n \in \mathbb{N}}$  in  $L^2([0, 1] \times \Omega)$  such that  $h_n(s, \omega)$  is of form  $\sum_{i=1}^p c_i(\omega) \mathbf{1}_{[t_i, t_{i+1}]}(s)$  and  $h_n$  converges to  $h$  in  $L^2([0, 1] \times \Omega)$ .

With (3.38),

$$\begin{aligned} \left| \mathbb{E} \left[ \int_0^1 g_n(s, \omega) h_\infty(s, \omega) ds \right] - \mathbb{E} \left[ \int_0^1 g_n(s, \omega) h_m(s, \omega) ds \right] \right| \\ \leq \sup_{n \in \mathbb{N}} \|g_n\|_{L^2([0, 1] \times \Omega)} \|h_\infty - h_m\|_{L^2([0, 1] \times \Omega)} \end{aligned}$$

and with (3.38)-(3.39),

$$\mathbb{E} \left[ \int_0^1 g_n(s, \omega) h_m(s, \omega) ds \right] \xrightarrow{n \rightarrow \infty} \mathbb{E} \left[ \int_0^1 g_\infty(s, \omega) h_m(s, \omega) ds \right].$$

It follows that  $g_n$  converges weakly in  $L^2([0, 1] \times \Omega)$  to  $g_\infty$ . In addition, (3.38) implies indeed the strong convergence of  $g_n$  to  $g_\infty$ , which means that  $\mathbb{E} \left[ \int_0^1 |g_n(s, \omega) - g_\infty(s, \omega)|^2 ds \right]$  converges to 0.

It is now possible to apply this argument to both  $Z^n$  and  $U^n$ .  $\square$

### 3.4 Convergence of the solution of the BSDE

The idea is now to prove that if  $(Y^{n,p}, Z^{n,p}, U^{n,p})$  converges to  $(Y^{\infty,p}, Z^{\infty,p}, U^{\infty,p})$  in a given sense, then this is also true for the  $(p + 1)$ -th Picard iteration.

Here, the notion of convergence is

$$\sup_{0 \leq t \leq 1} \left| Y_{\psi^n(t)}^{n,p} - Y_t^{\infty,p} \right|^2 + \int_0^1 |Z_{s-}^{n,p} - Z_{s-}^{\infty,p}|^2 ds + \lambda \int_0^1 |U_{s-}^{n,p} - U_{s-}^{\infty,p}|^2 ds \xrightarrow{n \rightarrow \infty} 0 \quad (3.40)$$

in  $L^1(\mathbb{P})$ , where  $\psi^n$  is a random one-to-one continuous mapping from  $[0, 1]$  to  $[0, 1]$  that converges uniformly to  $t \mapsto t$  almost surely.

Let us set

$$A_t^{n,p} = \int_0^t f(c_n(s), Y_{s-}^{n,p}, Z_{s-}^{n,p}, U_{s-}^{n,p}) dc_n(s)$$

and

$$A_t^{\infty,p} = \int_0^t f(s, Y_s^{\infty,p}, Z_s^{\infty,p}, U_s^{\infty,p}) ds.$$

**Lemma 4.** *If for some integer  $p$ ,  $(Y^{n,p}, Z^{n,p}, U^{n,p})$  converges to  $(Y^{\infty,p}, Z^{\infty,p}, U^{\infty,p})$  in the sense of (3.40), then  $A_{\psi^n(t)}^{n,p}$  converges uniformly in  $t$  to  $A_t^{\infty,p}$  in  $L^2(\mathbb{P})$ .*

*Proof.* Let us note first that  $A^{n,p}$  is piecewise constant on the intervals  $[k/n, (k+1)/n)$ . Let  $\xi^n(t)$  be the inverse of  $\psi^n(t)$ . Then

$$\begin{aligned} \sup_{t \in [0,1]} |A_{\psi^n(t)}^{n,p} - A_t^{\infty,p}| &= \sup_{t \in [0,1]} |A_t^{n,p} - A_{\xi^n(t)}^{\infty,p}| \\ &= \sup_{k=0, \dots, n-1} |A_{k/n}^{n,p} - A_{k/n}^{\infty,p}| + \sup_{k=0, \dots, n-1} \sup_{t \in [k/n, (k+1)/n]} |A_{k/n}^{\infty,p} - A_{\xi^n(t)}^{\infty,p}|. \end{aligned}$$

Because of **(A)** and (3.16), we easily get that the last term above converges to 0 in  $L^2(\mathbb{P})$  uniformly in  $p$ .

On the other hand, with **(A)**,

$$|A_{k/n}^{n,p} - A_{k/n}^{\infty,p}| \leq K \int_0^1 (|Y_s^{n,p} - Y_s^{\infty,p}| + |Z_s^{n,p} - Z_s^{\infty,p}| + |U_s^{n,p} - U_s^{\infty,p}|) ds.$$

Yet let us note that

$$\int_0^1 |Y_s^{n,p} - Y_s^{\infty,p}| ds \leq \int_0^1 |Y_s^{n,p} - Y_{\xi^n(s)}^{\infty,p}| ds + \int_0^1 |Y_{\xi^n(s)}^{\infty,p} - Y_s^{\infty,p}| ds.$$

The first term in the right-hand side of the previous inequality converges to 0 since  $Y^{n,p} - Y^{\infty,p} \circ \xi^n$  converges uniformly to 0 in  $L^2(\mathbb{P})$ .

Regarding the second term,  $s \mapsto Y_s^{\infty,p}$  is continuous except at the times at which the Poisson process jumps. Hence,  $Y_{\xi^n(s)}^{\infty,p}$  converges to  $Y_s^{\infty,p}$  for almost every  $s \in [0, 1]$  and then  $\int_0^1 |Y_{\xi^n(s)}^{\infty,p} - Y_s^{\infty,p}| ds$  converges to 0 almost surely.

With Lemma 2,  $\sup_{t \in [0,1]} |A_{\psi^n(t)}^{n,p} - A_t^{\infty,p}|$  converges to 0 in  $L^2(\mathbb{P})$ .  $\square$

**Proposition 3.** *Under Hypotheses (A), (B) and (H), let us consider the scaled random walks  $W^n$  and  $\tilde{N}^n$ . Then for any  $p \in \mathbb{N}$ ,  $(Y^{n,p}, Z^{n,p}, U^{n,p})$  converges to  $(Y^{\infty,p}, Z^{\infty,p}, U^{\infty,p})$  in the sense of (3.40).*

*Proof.* This will be done by induction on  $p$ . We rewrite (3.15) as

$$Y_t^{n,p+1} = \xi^n + A_1^{n,p} - A_t^{n,p} - \int_t^1 Z_{s-}^{n,p+1} dW_s^n - \int_t^1 U_{s-}^{n,p+1} d\tilde{N}_s^n - \int_t^1 V_s^n d\tilde{M}_s^n. \quad (3.41)$$

The induction hypothesis is that  $(Y^{n,p}, Z^{n,p}, U^{n,p})$  converges to  $(Y^{\infty,p}, Z^{\infty,p}, U^{\infty,p})$  in the sense of (3.40) so that our aim is to prove that the triplet  $(Y^{n,p+1}, Z^{n,p+1}, U^{n,p+1})$  converges to  $(Y^{\infty,p+1}, Z^{\infty,p+1}, U^{\infty,p+1})$  in the same sense.

As  $(Y^{n,0}, Z^{n,0}, U^{n,0}) = (0, 0, 0)$  and  $s \mapsto f(s, 0, 0, 0)$  is continuous, the first step of the induction is immediate from Corollary 1 using (H).

Taking conditional expectations w.r.t.  $\mathcal{F}_k^n$  in (3.15) and using the fact that  $Y_{t_k}^{n,p+1}$  is  $\mathcal{F}_k^n$ -measurable, we find that for  $t_k \leq c_n(t) < t_{k+1}$ ,

$$Y_t^{n,p+1} = \mathbb{E} \left[ \xi^n + \int_t^1 f(c_n(s), Y_{s-}^{n,p}, Z_{s-}^{n,p}, U_{s-}^{n,p}) dc_n(s) \mid \mathcal{F}_k^n \right].$$

So that

$$\begin{aligned} M_t^{n,p+1} &:= Y_t^{n,p+1} + \int_0^t f(c_n(s), Y_{s-}^{n,p}, Z_{s-}^{n,p}, U_{s-}^{n,p}) dc_n(s) = Y_t^{n,p+1} + A_t^{n,p} \\ &= \mathbb{E} \left[ \xi^n + \int_0^1 f(s-, Y_{s-}^{n,p}, Z_{s-}^{n,p}, U_{s-}^{n,p}) dc_n(s) \mid \mathcal{F}_k^n \right] \\ &= \mathbb{E} [M_1^{n,p+1} \mid \mathcal{F}_k^n] \quad \text{is a } \mathcal{F}^n \text{ martingale.} \end{aligned}$$

Moreover, we have the representation

$$\begin{aligned}
 M_t^{n,p+1} &= \mathbb{E} \left[ \int_0^1 Z_{s-}^{n,p+1} dW_s^n + \int_0^1 U_{s-}^{n,p+1} d\tilde{N}_s^n + \int_0^1 V_{s-}^{n,p+1} d\tilde{M}_s^n \middle| \mathcal{F}_k^n \right] \\
 &= \int_0^{t_k} Z_{s-}^{n,p+1} dW_s^n + \int_0^{t_k} U_{s-}^{n,p+1} d\tilde{N}_s^n + \int_0^{t_k} V_{s-}^{n,p+1} d\tilde{M}_s^n \\
 &= \int_0^t Z_{s-}^{n,p+1} dW_s^n + \int_0^t U_{s-}^{n,p+1} d\tilde{N}_s^n + \int_0^t V_{s-}^{n,p+1} d\tilde{M}_s^n.
 \end{aligned}$$

The last decomposition corresponds to the martingale representation theorem given in Corollary 1. In order to apply this corollary to the sequence of martingales  $\left\{ (M_t^{n,p+1})_{0 \leq t \leq 1} ; n \in \mathbb{N} \right\}$ , we have to prove the  $L^2(\mathbb{P})$  convergence of  $M_1^{n,p+1}$  (the terminal value).

Using the fact that  $Y^{n,p}$ ,  $Z^{n,p}$  and  $U^{n,p}$  are piecewise constant, we have that

$$|M_1^{n,p+1} - \xi - A_1^{\infty,p}| \leq |\xi^n - \xi| + |A_1^{n,p} - A_1^{\infty,p}|.$$

With Lemma 2, Lemma 4 and **(H)**, this last quantity tends to zero in  $L^2(\mathbb{P})$ .

Applying Corollary 1, we see that  $M^{n,p+1}$  converges to

$$M_t^{\infty,p+1} := \mathbb{E} \left( \xi + \int_0^1 f(s, Y_s^p, Z_s^p, U_s^p) ds \middle| \mathcal{F}_t \right) = Y_t^{\infty,p+1} + A_t^{\infty,p}, \quad (3.42)$$

in the sense that

$$\begin{aligned}
 \sup_{0 \leq t \leq 1} \left| M_{\psi^n(t)}^{n,p+1} - M_t^{\infty,p+1} \right|^2 + \int_0^1 |Z_s^{n,p+1} - Z_s^{\infty,p+1}|^2 ds \\
 + \lambda \int_0^1 |U_s^{n,p+1} - U_s^{\infty,p+1}|^2 ds \rightarrow 0
 \end{aligned}$$

in  $L^2(\mathbb{P})$ , where  $\psi^n$  is a random one-to-one continuous mapping from  $[0, 1]$  to  $[0, 1]$  that converges uniformly to  $t \mapsto t$  almost surely.

With (3.42) and Lemma 4, then we get the convergence of  $(Y^{n,p+1}, Z^{n,p+1}, U^{n,p+1})$  to  $(Y^{\infty,p+1}, Z^{\infty,p+1}, U^{\infty,p+1})$  in the sense of (3.40).  $\square$

## 4 Applications to decoupled system of SDE and BSDEJ and to the numerical computations of the solutions of PIDE

Let  $X$  be the solution of the  $d$ -dimensional SDE with jumps

$$X_t = x + \int_0^t \sigma(s-, X_{s-}) dW_s + \int_0^t b(s-, X_{s-}) ds + \int_0^t c(s-, X_{s-}) d\tilde{N}_s, \quad (4.43)$$

where we have assumed that

$$|b(t, x) - b(t, y)| + |\sigma(t, x) - \sigma(t, y)| + |c(t, x) - c(t, y)| \leq K'|x - y|, \quad (4.44)$$

$$\sup_{t \in [0,1]} (|b(t, 0)| + |\sigma(t, 0)| + |c(t, 0)|) \leq K'' \quad (4.45)$$

for all  $t \in [0, 1]$  and for all  $x, y \in \mathbb{R}$ . Of course, the BSDEJ

$$Y_t = \xi + \int_t^1 f(s, X_{s-}, Y_s, Z_s, U_s) ds - \int_t^1 Z_s \sigma(s-, X_{s-}) dW_s - \int_t^1 U_s d\tilde{N}_s \quad (4.46)$$

is linked to the non-linear PIDE (with  $a = \sigma \cdot \sigma^T$ ) by

$$\begin{aligned} \frac{\partial u(t, x)}{\partial t} + \sum_{i,j=1}^d \frac{1}{2} a_{i,j}(t, x) \frac{\partial^2 u(t, x)}{\partial x_i \partial x_j} + \sum_{i=1}^d b_i(t, x) \frac{\partial u(t, x)}{\partial x_i} \\ + \int_{\mathbb{R}} \left( u(t, x + c(t, x, z)) - u(t, x) - c_i(t, x, z) \frac{\partial u(t, x)}{\partial x_i} \right) \Pi(dz) \\ = f(t, x, u(t, x), \nabla u(t, x) \sigma(t, x), u(t, x + c(t, x, \cdot)) - u(t, x)) \end{aligned} \quad (4.47)$$

with the terminal condition  $u(1, x) = g(x)$ . It is standard in the theory of BSDE that  $u(t, X_t) = Y_t$  and thus  $u(0, x) = Y_0$ . One can compute similarly  $u(s, x)$  for any  $s \in [0, t)$  by using the solution to (4.43) starting from  $(s, x)$  instead of  $(0, x)$ .

We use now for  $\Pi(dz)$  the measure  $\Pi(dz) = \sum_{x_i} \pi_i \delta_{x_i}$ , where  $x_i$  belongs to the interval  $I_i$ .

Following Remark 2, we assume for the sake of simplicity that indeed  $\Pi(dz) = \lambda \delta_0$  and that the dimension of the Brownian motion  $W$  is 1. Hence, we rewrite  $c(t, x, z)$  as  $c(t, x)$  since only  $c(t, x, 0)$  is used.

The SDE (4.43) may be discretized the following way for an integer  $n$ : we set  $\chi_0^n = x$  and for  $i = 0, \dots, n$ ,

$$\begin{aligned} \chi_{i+1}^n &= \chi_i^n + hb((i+1)h, \chi_i^n) \\ &\quad + \sqrt{h}\sigma(((i+1)h, \chi_i^n))\epsilon_{i+1}^n + c((i+1)h, \chi_i^n)\eta_{i+1}^n \end{aligned} \quad (4.48)$$

where  $\eta^n$  is a Bernoulli approximation of the compensated Poisson process with intensity  $\lambda$  and  $\epsilon^n$  is a Bernoulli approximation of the Brownian motion  $W$ . Of course, the  $\chi_i^n$ 's are easily simulated.

Using our algorithm, it is then possible to find  $(y_i^n, z_i^n, u_i^n)_{i=1, \dots, n}$  adapted to  $(\mathcal{F}_i^n)_{i=0, \dots, n}$  that solves the discrete BSDE

$$y_i^n = y_{i+1}^n + hf((i+1)h, \chi_i^n, y_i^n, z_i^n, u_i^n) - z_i^n \epsilon_{i+1}^n - u_i^n \eta_{i+1}^n - v_i^n \epsilon_{i+1}^n \eta_{i+1}^n \quad (4.49)$$

for  $i = 0, \dots, n-1$  with the terminal condition  $y_n^n = g(\chi_n^n)$ .

We are looking for a function  $v^n(i, z)$  such that  $y_k^n = v^n(i, \chi_i^n)$  and  $v^n$  solves a discrete PDE.

For a function  $v$  on  $\{0, \dots, n\} \times \mathbb{R}$ , we define using (2.12) the discrete operators

$$\begin{aligned} D_0^n v(k, x) &= \mathbb{E}[v(k, x) | \mathcal{F}_k^n] \\ &= \frac{1-\kappa}{2} \left( v(k, x + hb(kh, x)) + \sqrt{h}\sigma(kh, x) + \kappa c(kh, x) \right) \\ &\quad + v(k, x + hb(kh, x) - \sqrt{h}\sigma(kh, x) + \kappa c(kh, x)) \\ &+ \frac{\kappa}{2} \left( v(k, x + hb(kh, x)) + \sqrt{h}\sigma(kh, x) + (\kappa-1)c(kh, x) \right) \\ &\quad + v(k, x + hb(kh, x) - \sqrt{h}\sigma(kh, x) + (\kappa-1)c(kh, x)) \end{aligned}$$

and

$$D_1^n v(k, x) = \mathbb{E}[v(k, x)\epsilon_{k+1}^n | \mathcal{F}_k^n] \text{ and } D_2^n v(k, x) = \mathbb{E}[v(k, x)\eta_{k+1}^n | \mathcal{F}_k^n],$$

for which formulae similar to the one for  $D_0^n v(k, x)$  can be given.

From these results, one can deduce the representation of the solution of a discrete PDE with the help of the  $\chi_i^n$ . This representation is similar to the representation of the solution of the BSDEJ in term of  $Y_t = u(t, X_t)$ , where  $u$  is the solution to the PIDE (4.47).

**Proposition 4.** *Let  $v^n$  be the solution to the discrete PDE*

$$\begin{aligned} v_i^n(i, x) &= D_0^n v^n(i+1, x) \\ &+ hf((i+1)h, x, v_i^n(i, x), h^{-1/2} D_1^n v^n(i+1, \chi_i^n), D_2^n v^n(i+1, x)) \\ &\text{for } i = 0, \dots, n-1, x \in \mathbb{R}, \end{aligned} \quad (4.50)$$

with the terminal condition  $v_i^n(1, x) = g(x)$ . If  $hK < 1$ , then this solution exists and is unique. In addition, the solution  $(y^n, z^n, u^n)$  to the discrete BSDE (2.10) with the terminal condition  $\xi = g(\chi_n^n)$  which we construct using our algorithm satisfies  $y_i^n = v^n(i, \chi_i^n)$ ,  $z_i^n = h^{-1/2} D_1^n v^n(i+1, \chi_i^n)$  and  $u_i^n = D_2^n v^n(i+1, \chi_i^n)$ .

*Proof.* As  $hK < 1$  the existence and uniqueness of  $v^n(i, \cdot)$  follows from the existence of the solution  $\rho(x)$  to

$$\begin{aligned} \rho(x) &= D_0^n v^n(i+1, x) + hf((i+1)h, x, \rho(x), h^{-1/2} D_1^n v^n(i+1, x), \\ &\quad (1-\kappa)^{-1} D_2^n v^n(i+1, x)) \end{aligned}$$

for any  $x \in \mathbb{R}$ , once  $v^n(i+1, \cdot)$  is known. Thus, one can proceed recursively with  $i = n-1$  down to 0.

Let  $(y^n, z^n, u^n)$  be given by our algorithm. We assume  $v^n(i+1, \chi_{i+1}^n) = y_{i+1}^n$ , which is true for  $i+1 = n$ . Using (2.7), (2.8), and the definitions of  $D_1^n$  and  $D_2^n$ ,

$$\begin{aligned} z_i^n &= h^{-1/2} \mathbb{E}[v^n(i+1, \chi_{i+1}^n) \epsilon_{i+1}^n | \mathcal{F}_i^n] = h^{-1/2} D_1^n v^n(i+1, \chi_i^n), \\ u_i^n &= \mathbb{E}[v^n(i+1, \chi_{i+1}^n) \eta_{i+1}^n | \mathcal{F}_i^n] = D_2^n v^n(i+1, \chi_i^n). \end{aligned}$$

Taking conditional expectation with respect to  $\mathcal{F}_i^n$  in (4.50), since  $D_0^n v^n(i+1, \chi_i^n) = \mathbb{E}[v^n(i+1, \chi_{i+1}^n) | \mathcal{F}_i^n]$ , we get

$$\begin{aligned} v^n(i, \chi_i^n) &= \mathbb{E}[y_{i+1}^n | \mathcal{F}_i^n] \\ &+ hf((i+1)h, \chi_i^n, v^n(i, \chi_i^n), h^{-1/2} D_1^n v^n(i+1, \chi_i^n), D_2^n v^n(i+1, \chi_i^n)), \end{aligned} \quad (4.51)$$

while taking the conditional expectation with respect to  $\mathcal{F}_i^n$  in (2.10), we get

$$\begin{aligned} y_i^n &= \mathbb{E}[y_{i+1}^n | \mathcal{F}_i^n] \\ &+ hf((i+1)h, \chi_i^n, y_i^n, h^{-1/2} D_1^n v^n(i+1, \chi_i^n), D_2^n v^n(i+1, \chi_i^n)). \end{aligned} \quad (4.52)$$

As  $hf(\cdot, \cdot, \cdot, \cdot, \cdot)$  is  $Kh$ -Lipschitz in its third argument with  $Kh < 1$ , we obtain that  $y_i^n$  and  $v^n(i, \chi_i^n)$  are equal.  $\square$

## 4.1 Convergence of the discrete SDE

We may also prove the convergence of the solution of the discrete SDE

$$X_t^n = x + \int_0^t \sigma(s-, X_{s-}^n) dW_s^n + \int_0^t b(s-, X_{s-}^n) dc^n(s) + \int_0^t c(s-, X_{s-}^n) d\tilde{N}_s^n \quad (4.53)$$

to the solution  $X$  of (4.43). The solution  $X^n$  is equal to the solution  $\chi^n$  of (4.48). The proof is similar to the one for BSDE, but is much more simpler.

**Proposition 5.** *The solution  $X^n$  to (4.53) converges in probability in the  $J_1$ -Skorokhod topology to the solution  $X$  to (4.43).*

*Proof.* We use the Picard scheme on the discrete SDE (4.48) on the time interval  $[0, 1]$ . For this, we set

$$\begin{aligned} \chi_{i+1}^{n,p+1} &= \chi_i^{n,p} + hb((i+1)h, \chi_i^{n,p}) \\ &\quad + \sqrt{h}\sigma(((i+1)h, \chi_i^{n,p}))\epsilon_{i+1}^n + c((i+1)h, \chi_i^{n,p})\eta_i^n \end{aligned} \quad (4.54)$$

with  $h = T/n$ ,  $\chi_i^{n,0} = x$  for  $i = 0, \dots, n$  and  $\chi_0^{n,p} = x$  for  $i = 0, \dots, n$ .

Let us note that for any  $n, p \geq 0$ ,  $\chi_i^n$ ,  $\epsilon_i^n$  and  $\eta_i^n$  are independent for any  $n = 1, \dots, n$ . In addition, the means of  $\epsilon_i^n$  and  $\eta_i^n$  are equal to 0, and  $\mathbb{E}[(\eta_i^n)^2] = \kappa_n(1 - \kappa_n)$ . Thus, for  $i = 1, \dots, n-1$ ,  $\Delta^p \chi_i^n = \chi_i^{n,p} - \chi_i^{n,p-1}$ , we obtain that

$$\begin{aligned} \mathbb{E}((\Delta^{p+1} \chi_{i+1}^n)^2) &\leq 3h^2 \mathbb{E} \left( \left( \sum_{j \leq i} (b((j+1)h, \chi_j^{n,p}) - b((j+1)h, \chi_j^{n,p-1})) \right)^2 \right) \\ &\quad + 3h \mathbb{E} \left( \sum_{j \leq i} (\sigma((j+1)h, \chi_j^{n,p}) - \sigma((j+1)h, \chi_j^{n,p-1}))^2 \right) \\ &\quad + 3\kappa_n(1 - \kappa_n) \mathbb{E} \left( \sum_{j \leq i} (c((j+1)h, \chi_j^{n,p}) - c((j+1)h, \chi_j^{n,p-1}))^2 \right). \end{aligned}$$

Furthermore, since the functions  $\sigma$ ,  $c$  and  $b$  are uniformly Lipschitz with a

constant  $K'$  with respect to the  $x$ -variable, we have that

$$\begin{aligned} \mathbb{E}((\Delta^{p+1}\chi_{i+1}^n)^2) &\leq 3(K')^2 h^2 \mathbb{E} \left( \left( \sum_{j \leq i} |\Delta^p \chi_j^n| \right)^2 \right) \\ &\quad + 6K^2 (h + \kappa_n (1 - \kappa_n)) \mathbb{E} \left( \sum_{j \leq i} |\Delta^p \chi_j^n|^2 \right). \end{aligned}$$

With the convexity inequality and since  $h = T/n$ , we get

$$\mathbb{E}((\Delta^{p+1}\chi_{i+1}^n)^2) \leq 3K^2 \left( \frac{T^2}{n} + 2 \left( \frac{T}{n} + \kappa_n (1 - \kappa_n) \right) \right) \mathbb{E} \left( \sum_{j \leq i} |\Delta^p \chi_j^n|^2 \right).$$

We set  $\|\chi^n\| := \sum_{i=0}^{n-1} \mathbb{E}((\chi_i^n)^2)$  for a family  $\chi^n = (\chi_0^n, \dots, \chi_n^n)$ . It defines a norm on such families for which  $\chi_0^n = 0$ . Since  $\Delta^p \chi_0^n = 0$ ,

$$\|\Delta^{p+1}\chi^n\| \leq 3(K')^2 (T^2 + 2T + 2n\kappa_n(1 - \kappa_n)) \|\Delta^p \chi^n\|.$$

As  $n(1 - \kappa_n)\kappa_n$  converges to  $\lambda T$ , we obtain that there exist  $k < 1$  and  $n_0$  such that  $\|\Delta^{p+1}\chi^n\| \leq k \|\Delta^p \chi^n\|$  for any  $n \geq n_0$  when  $T$  is small enough.

This prove that  $\chi^{n,p}$  converges to the solution of (4.48) as  $p \rightarrow \infty$  at a speed that is uniform in  $n$ . This also proves that

$$\forall p \geq 1, \sup_{n \geq n_0} \mathbb{E} \left( \sup_{i=1, \dots, n} (\chi_i^{n,p} - x)^2 \right) \leq \sup_{n \geq n_0} \mathbb{E} \left( \sum_{i=0}^n (\chi_i^{n,p} - x)^2 \right) \leq C \quad (4.55)$$

$$\sup_{n \geq n_0} \mathbb{E} \left( \sup_{i=1, \dots, n} (\chi_i^{n,p} - \chi_i^n)^2 \right) \leq \sup_{n \geq n_0} \mathbb{E} \left( \sum_{i=0}^n (\chi_i^{n,p} - \chi_i^n)^2 \right) \xrightarrow{p \rightarrow \infty} 0 \quad (4.56)$$

for some constant  $C$  that does not depend on  $n$  and  $p$ . The bound (4.55) also holds when  $\chi^{n,p}$  is replaced with the solution  $\chi^n$  to (4.48).

Let us set  $X_t^{n,p} = \chi_{h[t/h]}^{n,p}$ ,  $X_t^n = \chi_{h[t/h]}^n$  and  $X^{\infty,p}$  be the  $p$ -th step in the Picard scheme of the solution  $X^{\infty,\infty}$  to (4.43), with  $X_t^{\infty,0} = x$ . These processes  $X^{\infty,p}$  are known to converge to  $X^{\infty,\infty}$  in  $J_1$  in probability.

The idea of the proof is the same as the one of Corollary 1. We prove the following by induction: if  $(X^{n,p}, W^n, \tilde{N}^n)$  converges in probability to  $(X^{p,\infty}, W, \tilde{N})$  in the  $J_1$ -Skorokhod topology, then  $(X^{n,p+1}, W^n, \tilde{N}^n)$  converges

to  $(X^{p+1,\infty}, W, \tilde{N})$  in the  $J_1$ -Skorokhod topology. If this induction is true (it is of course true that  $X^{n,0} = x = X^{\infty,0}$ ), then it follows easily from (4.56) that  $X^n$  converges in probability in the  $J_1$ -Skorokhod topology to  $X^{\infty,\infty}$ .

For this, let us remark first that  $B(x) = x \mapsto \left( t \mapsto \int_0^t b(s-, x_{s-}) ds \right)$  is continuous from the space of càdlàg functions with the  $J_1$ -topology to the space of continuous function with the uniform norm. From the continuity of  $\sigma$  and  $c$  follows the convergence in probability of the 6-uple

$$(X^{n,p}, B(X^{n,p}), \sigma(\cdot, X^{n,p}), c(\cdot, X^{n,p}), W^n, \tilde{N}^n)$$

to  $(X^{\infty,p}, B(X^{\infty,p}), \sigma(\cdot, X^{\infty,p}), c(\cdot, X^{\infty,p}), W, \tilde{N})$  in the  $J_1$ -Skorokhod topology.

In addition,  $[W^n, W^n]_T = T$  and  $\mathbb{E}[[N^n, N^n]_T] = Tn\kappa_n(1 - \kappa_n) \xrightarrow{n \rightarrow \infty} \lambda T$ . Thus, this is sufficient to assert from the previous convergences and Proposition 2 that

$$\left( X^{n,p}, B(X^{n,p}), W^n, \tilde{N}^n, \int_0^\cdot \sigma(s-, X_{s-}^{n,p}) dW_s^n, \int_0^\cdot c(s-, X_{s-}^{n,p}) d\tilde{N}_s^n \right)$$

converge in probability to

$$\left( X^{\infty,p}, B(X^{\infty,p}), W, \tilde{N}, \int_0^\cdot \sigma(s-, X_{s-}^{\infty,p}) dW_s, \int_0^\cdot c(s-, X_{s-}^{\infty,p}) d\tilde{N}_s \right)$$

in the  $J_1$ -Skorokhod topology. This proves the induction step, and then  $X^n$  converges in probability in  $X^{\infty,\infty}$  in the  $J_1$ -Skorokhod topology.

The convergence is proved on a time interval  $[0, 1]$  where  $T$  is small enough, but it is easy to get it, using  $X_T^{n,p}$  instead of  $x$  as the initial condition, to prove the result on  $[0, 2T]$ , and so on.  $\square$

## 5 A numerical example

In this section, we deal with a numerical example. We consider  $N$  a Poisson process with  $\lambda = 1$  and  $c < 1$ , and the following BSDEJ:

$$dY_t = -cU_t dt + Z_t dW_t + U_t(dN_t - dt), \quad (5.57)$$

with  $\xi = N_T$ .

The explicit solution of (5.57) is given by

$$(Y_t, Z_t, U_t) = (N_t + (1 + c)(T - t), 0, 1).$$

Furthermore if  $\xi = 0$  then the solution is equal to  $(Y_t, Z_t, U_t) = (0, 0, 0)$ . This example is borrowed from [2].

We have implemented this method on a standard personal computing platform (PC), and have observed that it performs very well using simulated data, as can be seen from the simulated data in the Table 1 and Figure 1. Despite the apparent algebraic complexity of the equations (2.7), (2.8) and (2.11) one needs to solve at each step the conditional expectation to obtain  $y_{t_i}^n$ , the problem poses no difficulty. Using MATLAB's simulations and algebra capabilities (Version 7.0 running on the University of Valparaíso CIMFAV cluster) yielded best computing times.

In our implementation, and for computational conveniences we consider the case when  $T = 1$ . The iteration of the algorithm begins from  $y_{t_n}^n = \xi^n = N_1^n$  at time  $t_n = T = 1$  and proceeds backward to solve  $(y_{t_j}^n, z_{t_j}^n, u_{t_j}^n)$ , where  $t_j = j/n$ , at time  $j = 0$ . Values are given with 4 significant digits.

In the following table and picture we summarize the results.

$n$	$c = 0.3$	$c = 0.9$	$c = 0.1$	$c = 0.5$
10	1.27	1.81	1.09	1.45
100	1.2970	1.8910	1.0990	1.495
1000	1.2997	1.8991	1.0999	1.4995
2000	1.2999	1.8996	1.1	1.4998
3000	1.2999	1.8997	1.1	1.4998
4000	1.2999	1.8998	1.1	1.4999
4500	1.2999	1.8998	1.1	1.4999
4900	1.2999	1.8998	1.1	1.4999
5000	1.2999	1.8998	1.1	1.4999
<b>Real Value <math>Y_0</math></b>	<b>1.3</b>	<b>1.9</b>	<b>1.1</b>	<b>1.5</b>

Table 1: Numerical Scheme for  $dY_t = -cU_t dt + Z_t dB_t - U_t(dN_t - dt)$  with  $n$  from 10 until  $n = 5000$  steps,  $\lambda = 1$ ,  $T = 1$  and different values of  $c$ .

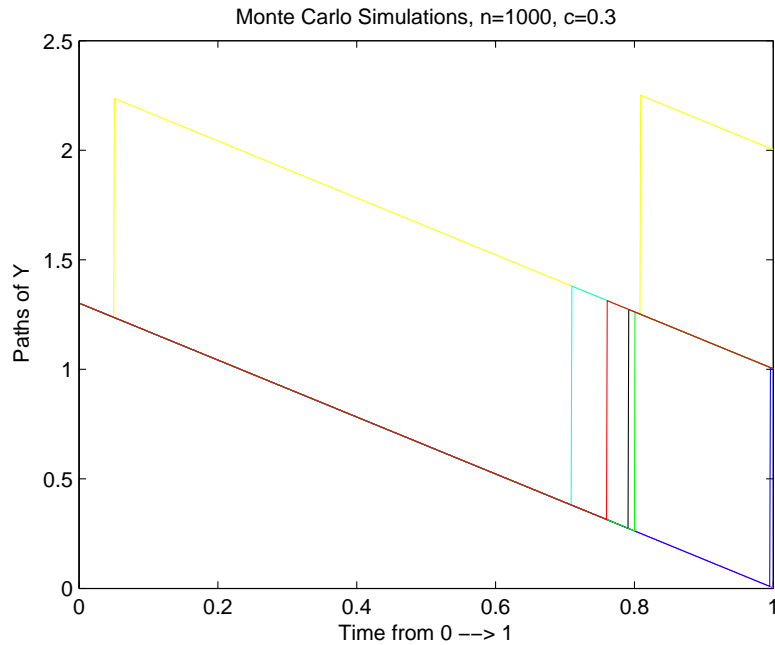


Figure 1: Monte Carlo Simulation;  $c = 0.3$ ,  $\lambda = 1$ ,  $n = 1000$  and  $T = 1$ .

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