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Statistical Shape Modeling using Morphological Representations

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Abstract

The aim of this paper is to propose tools for statistical analysis of shape families using morphological operators. Given a series of shape families (or shape categories), the approach consists in empirically computing shape statistics (i.e., mean shape and variance of shape) and then to use simple algorithms for random shape generation, for empirical shape confidence boundaries computation and for shape classification using Bayes rules. The main required ingredients for the present methods are well known in image processing, such as watershed on distance functions or log-polar transformation. Performance of classification is presented in a well-known shape database.

1. Introduction

Object recognition based on shape information is a classic problem in image processing. That motivates the development of tools for statistical analysis of shapes in \mathbb{R}^2 . In [5] is described a geometric technique to parametrize curves based on their arc lengths and their angle function to represent and analyze shapes. Techniques for clustering learning in planar shapes are described in [9]. Shape classification problem is referred as “classify a given shape into one of the predefined classes”, it can be solved using a shape-probability per class. Usually the first step towards the design of a shape classifier is a feature extraction followed by a shape matching stage. Curvature, chain codes, Fourier descriptors, tangent space representation [6] and beam angle statistics (BAS) [2] are examples of contour-based descriptors. Dynamic programming (DP)-based shape matching is applied successfully in [8] for shape matching. Hidden Markov models (HMMs) are also being explored as one of the possible shape modeling and classification frameworks [11].

This paper presents an approach to analyze a given family of shapes based on the computation of mean

shape and associated variance of shape, obtained by means of a mathematical morphology algorithm. Starting for the empirically obtained mean and variance, it is introduced a novel methodology of shape classification, using a training stage producing an associated probability map per class. Additionally, random shape generation and empirical confidence shape boundaries computation are also investigated. All the examples of the paper are obtained with shape families of the MPEG-7 database. Additional examples of the different algorithms using the complete database can be found in the website ¹.

2. Statistical Shape Modelling

2.1 Mean shape and variance of shape using mathematical morphology

Let $\mathcal{X}_j = \{X_1^j, X_2^j, \dots, X_{N_j}^j\}$ be a family (or collection) of N_j shapes for the class j , for $j = 1, \dots, J$ where $X_i^j \in P(E)$ represents the set (or binary image) of the shape i in the family of shapes labeled by j , and the support space E is a nonempty set. Typically for the digital 2D images $E \subseteq Z^2$. The set X_i^j is a compact set (and typically a closed simply connected set). The family \mathcal{X}_j can be considered as a random variable of shape, where X_i^j represents a realization of this random variable. We make the assumption that each class of shapes, i.e. the set of shapes that can be identified with a common concept or object, e.g. fish shapes, are they are centered, oriented and scaled.

Let us start by the classical definition of the mean of N samples: μ is the value such that $f(c) = \sum_i^N (x_i - c)^2$ is minimal. In the extension to a shape family \mathcal{X}_j , we start by constructing the sum of distance functions to the frontier sets $\Upsilon(\mathcal{X}_j)$ using the squared Euclidean distance, or an appropriate squared discrete distance [3] (Δ_c^2 denote outer distance and Δ^2 inner distance). Thus, the mean shape $\mu_{\mathcal{X}_j} \in P(E)$ is the compact set associ-

¹<http://cmm.ensmp.fr/~velasco/Shape>

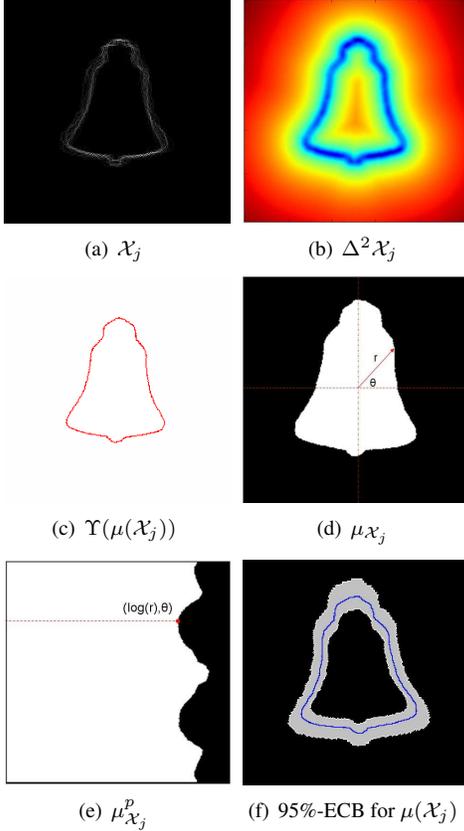


Figure 1. Analysis of “bell” shapes using the proposed approach.

ated to of locally minimal contour of obtained from the functional:

$$\Delta^2 \mathcal{X}_j = \sum_i^N [\Delta_c^2 X_i^j + \Delta^2 X_i^j] \quad (1)$$

The contour, or frontier, of the mean shape $\mu_{\mathcal{X}_j}$ will be denoted by $\Upsilon(\mu_{\mathcal{X}_j})$. As it was introduced in [1], this optimal contour can be easily obtained by computing the watershed line of the inverse (or complement) of $\Delta^2 \mathcal{X}_j$. An example of this analysis is given in Fig. 1. Additionally, using also the generalization of the classical statistical notion, the variance of a shape family is defined for each point in the mean contour as the function $\sigma_{\mathcal{X}_j} : \Upsilon(\mu_{\mathcal{X}_j}) \rightarrow \mathbb{R}^+$, given by [7]:

$$\sigma_{\mathcal{X}_j} = \frac{\Delta^2 \mathcal{X}_j}{N_j} \quad (2)$$

It is important to remark that the variance of shape is a set of nonnegative values associated to each point in $\Upsilon(\mu_{\mathcal{X}_j})$.

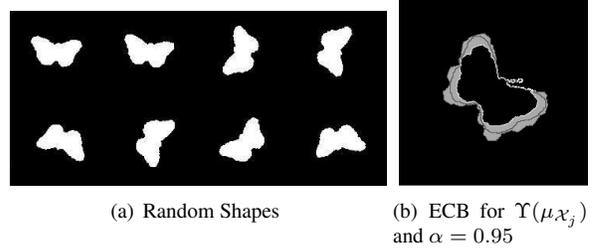


Figure 2. Random shapes and empirical confidence boundaries for $\Upsilon(\mu_{\mathcal{X}_j})$ in the “Butterfly” shapes family

2.2 Random Shapes Generation

A straight statistical application of the representation $(\mu_{\mathcal{X}_j}, \sigma_{\mathcal{X}_j})$ is the generation of random shapes. Our proposal is to generate a random perturbation of $\mu_{\mathcal{X}_j}$ in its log-polar representation [1], and these stochastic perturbations must follows the empirical variance $\sigma_{\mathcal{X}_j}$. More precisely, the approach is detailed in algorithm 1. Fig.1(e) $\mu_{\mathcal{X}_j}^p$ gives the log-polar coordinate representation (superscript p will denote log-polar space) of the mean shape depicted in Fig.1(d). An example of random shape generation according to our algorithm, for a family of butterflies, is presented in Fig. 2(a). More examples can be found in the website¹.

Input : Family of shape \mathcal{X}_j

output: Random shape $R(\mathcal{X}_j)$

Calculate $\mu_{\mathcal{X}_j}$ and $\sigma_{\mathcal{X}_j}$;

Transform to log-polar $\mu_{\mathcal{X}_j}^p$ and $\sigma_{\mathcal{X}_j}^p$;

$u_1 = U(-1, 1), u_2 = U(0, 2\pi)$;

for $i \leftarrow 1$ **to** 2π **do**

$R^p(\theta_i, \text{mod}(\cdot + u_2, 2\pi)) = \mu_{\mathcal{X}_j}^p(\theta_i, \cdot) + u_1 \sigma_{\mathcal{X}_j}^p(\theta_i, \cdot)$;

end

R = polar inverse transformation of R^p ;

Algorithm 1: Random Generation of Shapes

2.3 Empirical Confidence Boundaries for shape mean

Under the assumption of independence gaussian distribution in each direction of the log-polar transformation, i.e., $\mathcal{N}(0, \sigma_{\mathcal{X}_j}^p(\theta_i))$, we can define the notion of Empirical Confidence Boundaries (ECB) of shape. The idea is to extract a confidence boundary around $\mu_{\mathcal{X}_j}$ using the percentiles associated to a gaussian distribution. Technical description of our approach is given in algorithm 2, where Δ denotes the symmetric set difference and Φ^{-1} represents the inverse of the standardized

gaussian cumulative distribution function. Examples of ECB of shape are showed in Figs. 2(b) and 1(f).

Input : Family of shape \mathcal{X}_j , confidence parameter α
output: Boundary CB_α

Calculate $\mu_{\mathcal{X}_j}$ and $\sigma_{\mathcal{X}_j}$;
 Transform to log-polar $\mu_{\mathcal{X}_j}^p$ and $\sigma_{\mathcal{X}_j}^p$;

$$k_1 = \Phi_{(1-\alpha_2)}^{-1};$$

$$k_2 = \Phi_{(\alpha_2)}^{-1};$$

for $i \leftarrow 1$ **to** 2π **do**

$$\left| \begin{array}{l} S_1^p(\theta_i) = \mu_{\mathcal{X}_j}^p(\theta_i) + k_1 \sigma_{\mathcal{X}_j}^p(\theta_i); \\ S_2^p(\theta_i) = \mu_{\mathcal{X}_j}^p(\theta_i) + k_2 \sigma_{\mathcal{X}_j}^p(\theta_i); \end{array} \right.$$

end

$$CB_\alpha = S_1 \Delta S_2 ;$$

Algorithm 2: Empirical confidence boundary of a shapes family

2.4 Classification in families of shapes

Starting again from the pair $(\mu_{\mathcal{X}_j}, \sigma_{\mathcal{X}_j})$, the goal now is to define a probability associated to the family of shapes \mathcal{X}_j . Firstly, we introduce a measure for a point $\vec{v}_i \in \mathbb{R}^2$ w.r.t. the shape family \mathcal{X}_j as follows:

$$\Psi_j(\vec{v}_i) = \max_k \mathcal{N}(\Upsilon(\mu_{\mathcal{X}_j})(k), \Sigma_{\mathcal{X}_j}(k))(\vec{v}_i) \quad (3)$$

where the bivariate gaussian distribution is centered at each point k which belongs to the mean contour $\Upsilon(\mu_{\mathcal{X}_j})(k)$ and has an uncorrelated covariance matrix $\Sigma_k = \mathbf{I}_{(2 \times 2)} \sigma_{\mathcal{X}_j}(k)$, where $\sigma_{\mathcal{X}_j}(k)$ is the variance value at point $\Upsilon(\mu_{\mathcal{X}_j})(k)$. The gaussian distribution in expression (3) can be clearly substituted for another elliptically contoured distributions or even an asymmetric distribution to penalize any direction. Fig. 3(a) shows an illustration of the assumption of bivariate gaussian distribution at each value of $\Upsilon(\mu_{\mathcal{X}_j})$. Secondly, given a set of shapes \mathcal{X}_j , their associated probability map (APM) is defined as

$$APM_j = \Psi_j(\vec{v}_i), \forall \vec{v}_i \in \mathbb{R}^2 \quad (4)$$

Under bivariate gaussian assumption, the expression (4) can be calculated for a given point \vec{v} as

$$\begin{aligned} APM_j &= \max_k \int_0^{2\pi} \int_0^{c_k} r \exp \frac{-r^2}{2\sigma_k^2} dr d\theta \\ &= 1 - \min_k \left\{ \exp \frac{-c_k^2}{2\sigma_k^2} \right\} \end{aligned}$$

where c_k^2 is the distance between \vec{v} and each point $k \in \Upsilon(\mu_{\mathcal{X}_j})$. Fig. 3(b) shows the APM_j for the class of bells. Other examples can be consulted in the website ¹.

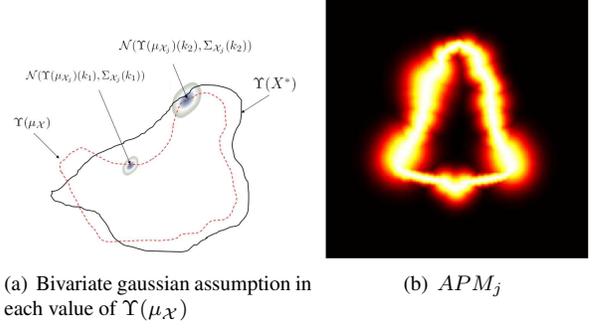


Figure 3. Shape family, mean and variance contours using proposed approach.

The next step is to use APM_j in order to obtain an analytical expression to calculate the empirical conditional probability per class j for a new shape $X^* \in P(E)$, denoted $P[J = j/X^*]$, knowing the empirical mean $\mu_{\mathcal{X}_j}$ and variance $\sigma_{\mathcal{X}_j}$. Intuitively, a shape X^* going through zones of the space with highest probability APM_j has to accumulate more likely in it. On the other hand, if $\int \Upsilon(X^*) \ll \int \Upsilon(\mu_{\mathcal{X}_j})$, its probability has to reduce, because we assume that all the shapes are oriented and rescaled. Therefore, taking into account these ideas, the *a posteriori* probability is given by

$$P[J = j/X^*] = \frac{\int_{APM_j} X^*}{\max(\int X^*, \int \Upsilon(\mu_{\mathcal{X}_j}))} \quad (5)$$

where \int denote the integral over an APM with constant value 1 for every point. Note that the denominator is comparing lengths between the frontier of the mean shape in the class j and the frontier of the new shape to predict.

Consequently, we can use (5) in a Bayes decision rule to design the following statistical classifier $g(\cdot)$:

$$g(X^*) = \operatorname{argmax}_j (\pi_j P[J = j/X^*]) \quad (6)$$

where π_j is the *a priori probability* and it represents the likelihood for the appearance of the class j in the global data set.

3. Experiments

The well-known MPEG-7 test database consists of 70 types of objects, each family having 20 different shapes. The database is challenging due to the presence of examples that are visually dissimilar from other members of their class, and examples that are highly

