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Lifetime estimation on moving sub-cellular objects in frequency domain FLIM imaging

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Fluorescence lifetime is usually defined as the average nanosecond-scale delay between excitation and emission of fluorescence. It has been established that lifetime measurement yields numerous indications on cellular processes such as inter-protein and intra-protein mechanisms through fluorescent tagging and Förster resonance energy transfer (FRET). In this area, frequency domain fluorescence lifetime imaging microscopy (FD FLIM) is particularly well appropriate to probe a sample non-invasively and quantify these interactions in living cells. The aim is then to measure fluorescence lifetime in the sample at each location in space from fluorescence variations observed in a temporal sequence of images obtained by phase modulation of the detection signal. This leads to a sensitivity of lifetime determination to other sources of fluorescence variations such as intracellular motion. In this paper, we propose a robust statistical method for lifetime estimation on both background and small moving structures with a focus on intracellular vesicle trafficking. © 2015 Optical Society of America

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1. INTRODUCTION

The fluorescence lifetime of a given fluorescent species is defined as the average time during which the fluorophore stays in excited state before relaxing to its ground state and possibly emitting a photon. The fluorophore energy can be relaxed to its ground state through various de-excitation pathways. Those pathways are influenced by various biophysical phenomena such as ion concentration, collision or non-radiative energy transfer. Exploiting those properties, Fluorescence Lifetime Imaging Microscopy (FLIM) is now a commonly-used technique for sensing fluorophore environment in a living biological sample (pH, ions, temperature ...). Fluorescence lifetime measurement is also particularly useful to detect Förster resonance energy transfer (FRET) between fluorescent molecules in close proximity [1, 2].

In this area, two main technologies can be considered: time domain methods and frequency domain methods. Time domain methods such as TCSPC (Time-Correlated Single-Photon Counting) measure the mean delay between a photon emission and a pulsed excitation. This approach is now well established but mainly recommended for fixed sample with no motion during acquisition [3, 4]. Unlike TCSPC, a FLIM measurement in

the frequency domain is defined as a short sequence of images presenting a sinusoidal intensity footprint that reflects the sinusoidal excitation of the sample followed by the phase modulation of the emitted signal. A fluorescence lifetime variation results in a phase variation in the detection signal. The acquisition process is 10 to 500 times faster than time domain methods [5–7]. FD FLIM is then intrinsically more suitable for dynamical process analysis and for the study of living cells.

In FD FLIM, Fourier transform or equivalently least mean squares fitting of a sine function is used traditionally to estimate lifetime maps from FD FLIM measurements [2, 8]. Two problems arise from those approaches. First, it is implicitly assumed that the noise induced during acquisition is temporally stationary and Gaussian. This traditional assumption does not hold because of the photonic nature of the emitted signal and the high intensity gradient present in FD FLIM measurement. Secondly, the specimen is assumed to remain immobile during acquisition. Even if frequency domain methods are faster than time domain methods, transient motions of intracellular structures can still happen during FD FLIM acquisitions and thus inducing artifacts in the estimated lifetime map [5, 9]. In this paper, we propose to overcome those two difficulties to address

Fig. 1. Instrumental setup showing the system components and the corresponding FD FLIM signals. We denote $E(t)$ the excitation light, $F(t)$ the fluorescence signal, $G_k(t)$ the detection modulation signal and S_θ the FD FLIM measurement.

important biological issues in dynamic intra-cellular traffic experiments. Our line of work consists in exploiting the usual FD FLIM intensity model combined with the original noise modeling described in [10] to estimate accurately fluorescence lifetime on static background and on moving vesicles that underwent non-negligible motions during the acquisition. In what follows, we propose a novel and original statistical framework based on robust statistics to estimate alternately the trajectory and the fluorescence lifetime of moving vesicles. To our knowledge, this issue is not addressed in the literature and this paper describes an innovative solution to solve this problem.

The remainder of the paper is organized as follows. Section 2 presents the principles of fluorescence lifetime imaging techniques in the frequency domain (FD FLIM). Section 3 describes our estimators and correction terms for fluorescence lifetime on the static background. Section 4 describes our lifetime estimation approach on dynamical structures. In Section 5, the proposed methodology is evaluated on experimental image series with controlled fluorophores and simulated measurements. On experimental data, it is extremely difficult to obtain a ground truth for the fluorescence lifetime of moving vesicles. Accordingly, the accuracy of our method is tested *via* the motion estimation results that are compared against competitive particle trackers. Finally, the complete lifetime map reconstruction framework is applied to the study of a GFP-tagged donor-receptor expressed in RPE1 cells.

2. FD FLIM AND SIGNAL MODELING

A. Basics in frequency domain FLIM measurement

Instead of the pulsed laser excitation that characterizes time-domain methods, the FD FLIM excitation source is a sinusoidally modulated signal $E(t)$ (blue ray Figure 19) defined as:

$$E(t) = C_E + A_E \sin(\omega_E t + \phi_E) \quad (1)$$

where ω_E denotes the radial frequency of the excitation signal, C_E the offset, A_E the amplitude, ϕ_E is the phase delay and t

the temporal variable. The excitation $E(t)$ can be operated with various illumination methods, such as wide-field or confocal microscopy. The signal emitted by the sample $F(t)$ (green ray in Figure 19) is defined as :

$$\begin{aligned} F(t) &= E(t) * R(t) \\ &= C_F + A_F \sin(\omega_E t + \phi_E - \text{atan}(\omega_E \tau)) \end{aligned} \quad (2)$$

where $*$ denotes the convolution operator, $R(t) = R_0 e^{-t/\tau}$ is the response of a fluorescent sample at time t with a given lifetime τ and R_0 is the response at time $t = 0$. Accordingly, the phase delay $\phi_E - \text{atan}(\omega_E \tau)$ allows us to recover the fluorescence lifetime τ .

Experimentally, for measuring lifetime in the order of nanoseconds (ns) the frequency ω_E has to be > 1 MHz, too high to be compatible with the frame rate of a CCD sensor. Therefore the signal $F(t)$ is phase-modulated with a finite set of K sinusoidal signals $M_k(t)$, $k \in [1, K]$ (see black signal in Figure 19) at the same frequency ω_E (homodyne detection) defined as:

$$M_k(t) = C_M + A_M \sin\left(\omega_E t + \phi_{M_0} + \frac{2\pi k}{K}\right) \quad (3)$$

where C_M , A_M and ϕ_{M_0} denote respectively the offset and the amplitude of M_k and the phase of signal M_k for $k = 0$. This modulation is operated by applying a time-varying voltage to the photocathode terminal in the intensified CCD. The intensified CCD is in turn required to acquire the lowest intensity images resulting from modulation. Thanks to the low pass effect of the CCD detector, the higher frequencies of the phase modulated-signal $F(t)M_k(t)$ are attenuated and become negligible. After phase-modulation, the K time-independent signals are described as a function of $k \in [1, K]$:

$$S_\theta(k) = \mathcal{G}_{\text{ccd}}(F(t)M_k(t)) = C_S + A_S \cos\left(\frac{2\pi k}{K} + \phi_\tau\right) \quad (4)$$

where $\mathcal{G}_{\text{ccd}}$ represents the CCD function, $C_S = C_F C_M$, $A_S = A_F A_M$ and $\phi_\tau = (\phi_{M_0} - \phi_E + \text{atan}(\omega_E \tau))$ is the phase.

Fig. 2. Principle of FD FLIM measurement and post-processing A: Example of an image sequence showing the intensity variation ($S_\theta(k)$) at different phase shift values ($K=12$ values distributed on 2π radians). B: Example of $S_\theta(k)$ observed at a given pixel (black dots) with a sine fitting curve (red). C: Fluorescence lifetime map reconstructed from the image sequence in B. Scale bar is $10 \mu\text{m}$,

In what follows, we will denote $\theta = (C_S, A_S, \phi_\tau)^T \in \mathbb{R}^3$ the FD FLIM parameters at a given pixel. FD FLIM thus consists in K images presenting a sinusoidal footprint (see Figure 2.A). For each pixel in the image domain, the phase ϕ_τ combined with proper calibration yields the lifetime estimate τ . A detailed derivation for the fluorescent signal expression and the phase modulation can be found in Appendix A.

B. Conventional estimation of FD FLIM parameters

Estimating fluorescence lifetime from FD FLIM measurements amounts to estimating the parameters that control the sinusoidal footprint (see Figure 2.B).

Let us denote a measured image sequence of K images $I : \Omega \times [1, K] \rightarrow \mathbb{R}_+$ where $\Omega \subset \mathbb{R}^2$ is the image domain. Let $\phi_\tau : \Omega \rightarrow \mathbb{R}_+$ and $\tau : \Omega \rightarrow \mathbb{R}_+$ be respectively the unknown phase and lifetime map. The usual approach to estimate the phase $\phi_\tau(\mathbf{x})$ is based on the Fourier transform and is defined as:

$$\hat{\phi}_\tau(\mathbf{x}) = -\text{atan} \left(\frac{\sum_{k=1}^K I(\mathbf{x}, k) \cos(\frac{2\pi k}{K})}{\sum_{k=1}^K I(\mathbf{x}, k) \sin(\frac{2\pi k}{K})} \right) + \frac{\pi}{2}, \quad (5)$$

where $I(\mathbf{x}, k)$ is the intensity measured at pixel $\mathbf{x}$ in the phase-modulated image k . The large majority of contributions in FD FLIM exploits this estimator [8, 11–13] as much as commercial software such as LI-FLIM by Lambert Instrument. In order to handle the impact of photobleaching, [12] proposed to randomize the order of phase modulation. The noise induced by photobleaching is thus shown to be mixed into the uncorrelated measurement noise.

Let us denote the true signal $S_\theta : \Omega \times [1, K] \rightarrow \mathbb{R}_+$ where $S_\theta(\mathbf{x}, k)$ is the signal value at a given pixel $\mathbf{x} \in \Omega$ and frame $k \in [1, K]$. The equivalent statistical model put forward in those

approaches is as follows:

$$I(\mathbf{x}, k) = S_\theta(\mathbf{x}, k) + \varepsilon(\mathbf{x}, k), \quad (6)$$

where $\varepsilon(\mathbf{x}, k)$ is assumed to be a white Gaussian noise such as $\mathbb{E}[\varepsilon(\mathbf{x}, k)] = 0$ and $\text{Var}[\varepsilon(\mathbf{x}, k)] = v^2$. The associated optimal estimator is found as the minimization of the following least mean squares problem:

$$\hat{\theta}(\mathbf{x}) = \underset{\theta(\mathbf{x})}{\text{argmin}} \sum_{k=1}^K (I(\mathbf{x}, k) - S_\theta(\mathbf{x}, k))^2. \quad (7)$$

Very few studies have been focusing on improving this phase estimator. For instance, in [14, 15], the authors proposed a linearization of the problem in the Fourier domain. Finally, given a two-dimensional map of phase estimation, the fluorescence lifetime map at pixel $\mathbf{x} \in \Omega$ is given by (see Appendix A.2):

$$\hat{\tau}(\mathbf{x}) = \frac{\tan(\hat{\phi}_\tau(\mathbf{x}) - \phi_{M_0} + \phi_E)}{\omega_E}. \quad (8)$$

Note that the phase delay $(\phi_{M_0} - \phi_E)$ of the system must be calibrated by recording a reference sequence of a fluorescent sample with known lifetime τ_{ref} while the frequency ω_E is controlled during acquisition.

3. ROBUST ESTIMATION OF FD FLIM PARAMETERS OVER STATIC BACKGROUND

As described above, least mean squares minimization assumes that the errors are expected to be Gaussian distributed and homoscedastic. While the Gaussian approximation is correct in most cases due to the high counts of photons after intensification, the homoscedastic assumption does not hold in FD FLIM

imaging because of photon counting noise, large intensity gradient and the specificity of the FD FLIM image formation process (see Figure 3). In this Section, we describe a novel estimator based on our previous work on ICCD noise analysis that takes this heteroscedasticity into account and provide a complementary scheme to detect outlier induced by small object motions.

A. Noise variance modeling and estimator

In [10], we described a study of instrument-induced noise based on sensor theory combined with experimental consideration. We showed that the quadratic relationship between the noise variance and intensity level demonstrated in [16] is strongly perturbed by optical aberration induced by the intensifier setup. We thus proposed in our previous work a correction term to the parametric noise variance model $v^2(\mathbf{x}, k)$. The new model for the noise variance is thus defined as the product of two terms at pixel $\mathbf{x}$ and frame k :

$$v^2(\mathbf{x}, k) = (a\mathbb{E}[I(\mathbf{x}, k)]^2 + b\mathbb{E}[I(\mathbf{x}, k)] + c) \left(e^{-\frac{(x-x_0)^2}{h_x} - \frac{(y-y_0)^2}{h_y}} + o \right). \quad (9)$$

where $\mathbb{E}[\cdot]$ denotes the mathematical expectation and a, b, c are the coefficients of a quadratic polynome. The first term describes the overall contribution of the photon counting noise, read out noise and intensifier variability as described in [10, 16]. The second term is a spatial correction term determined empirically to cope with spherical aberration. In [10], we have proposed a Gaussian function parametrized by x_0, y_0, h_x, h_y and an offset o to take the blur observed on real images into account. The blurring effect increases with the distance from the center $\mathbf{x}_0 = (x_0, y_0)^T$ of the aberration (the center of the lens inside the intensifier). The x_0, y_0, h_x and h_y parameters are estimated from the same plain fluorescein sample used for phase calibration (see [10] for details).

Equipped with this non-stationary noise model, the FD FLIM parameters over the background are estimated by minimizing the following weighted least mean squares criterion:

$$\hat{\theta}(\mathbf{x}) = \underset{\theta}{\operatorname{argmin}} \sum_{k=1}^K \frac{(I(\mathbf{x}, k) - S_{\theta}(\mathbf{x}, k))^2}{v^2(\mathbf{x}, k)} \quad (10)$$

where $S_{\theta}(\mathbf{x}, k)$ is the signal defined as in (4). The potential of this weighted least mean squares estimation method is discussed in Section 5.

B. Initialization and convergence

The initialization for the parameter $\theta(\mathbf{x})$ takes into account the set of measurements as well as the expected value taken by the fluorescent sample. Indeed, while the offset the amplitude of the sinusoidal signal are predicted using the mean and the range of the measure intensities, the phase is initialized using the lifetime of the fluorophore in the absence of FRET and the calibrated phase of the system. More formally, the initialization is carried out as follows:

$$\begin{cases} C_S(\mathbf{x}) &= \frac{1}{K} \sum_{k=1}^K I(\mathbf{x}, k) \\ A_S(\mathbf{x}) &= \frac{1}{2} \left(\max_{k \in [1, K]} I(\mathbf{x}, k) - \min_{k \in [1, K]} I(\mathbf{x}, k) \right) \\ \phi_{\tau}(\mathbf{x}) &= \phi_{M_0} - \phi_E + \operatorname{atan}(\omega_E \tau_{ref}) \end{cases}$$

The evaluation of the magnitude of our weighted residuals over a large parameter space constrained by $\phi_{\tau}(\mathbf{x}) \in [0, 2\pi]$

Fig. 3. Residuals of least-squares estimates of the parameters controlling the sinusoid described by the phase modulation images (see (4)). The residuals associated to each pixel in the whole stack presented in Figure 2 are plotted. The variability of the residual amplitude highlights the presence of strong heteroscedasticity.

Fig. 4. The magnitude of the residuals (see (10)) evaluated on a large parameter space for a FD FLIM measurement highlights the existence of a single local minimum (in red). Initialization is shown in green. Left: Minimum value projection along the “constant component” (offset) axis. Right: Minimum value projection along the “amplitude” axis.

highlights a single minimum (see Figure 4). Given our efficient initialization procedure, a conventional Gauss-Newton algorithm converges quickly towards the global minimum.

C. Motion-induced outlier management

Assuming a lifetime map $\hat{\tau} : \Omega \rightarrow \mathbb{R}^+$ computed from (8)-(10), we propose here a complementary approach to estimate the background lifetime in pixels “occluded” by moving objects (see Figs. 6 and 12). To tackle this problem, we analyze the value $\hat{\tau}(\mathbf{x})$ with respect to the lifetime values $\hat{\tau}(\mathbf{y})$ estimated at pixels $\mathbf{y}$ in the neighborhood $V(\mathbf{x})$ of $\mathbf{x}$. We appeal to tools from robust statistics to determine how large the lifetime $\hat{\tau}(\mathbf{x})$ can be before we consider this value to be an “outlier”. We define the Median Absolute Deviation $\kappa_{\tau}(\mathbf{x})$ to determine an adaptive threshold (see [17]):

$$\begin{aligned} \kappa_{\tau}(\mathbf{x}) &= 1.4826 \times \operatorname{median}_{\mathbf{y} \in V(\mathbf{x})} |\hat{\tau}(\mathbf{y}) - \operatorname{median}_{\mathbf{y} \in V(\mathbf{x})} \hat{\tau}(\mathbf{y})|, \quad (11) \\ &\triangleq 1.4826 \times \operatorname{MAD}(\hat{\tau}(\mathbf{y}), \mathbf{y} \in V(\mathbf{x})), \end{aligned}$$

where “MAD” denotes the median absolute deviation and $V(\mathbf{x}) \subset \Omega$ denotes a spatial neighborhood (or window) cen-

tered at location $\mathbf{x}$. The constant in (14) is calculated from the fact that the MAD of a zero-mean normal distribution with unit variance is $0.6745 = 1/1.4826$. Given an arbitrary real value $\lambda > 0$, we can detect outliers by using the usual Chebyshev's inequality which can be applied to any arbitrary distribution:

$$\text{Prob}(|\mathcal{T}(\mathbf{x}) - \mu_\tau(\mathbf{x})| \geq \lambda \kappa_\tau(\mathbf{x})) \leq \frac{1}{\lambda^2} \quad (12)$$

where $\mathcal{T}(\mathbf{x})$ denotes a random variable (representing lifetime at pixel $\mathbf{x}$) with finite (robust) expected value:

$$\mu_\tau(\mathbf{x}) = \text{median}_{\mathbf{y} \in V(\mathbf{x})}(\tau(\mathbf{y})) \quad (13)$$

and finite non-zero variance κ_τ . Setting $\lambda = 3$ as chosen in all our experiments, means that the probability that the values $\hat{\tau}(\mathbf{x})$ lying outside the interval $[\mu_\tau(\mathbf{x}) - 3\kappa_\tau(\mathbf{x}), \mu_\tau(\mathbf{x}) + 3\kappa_\tau(\mathbf{x})]$ does not exceed 88.89% and the probability for a pixel to be classified as an outlier is 11.11%. The neighborhood size $|V(\mathbf{x})|$ is typically set to 30×30 pixels to account for large objects.

Let $O : \Omega \rightarrow \{0, 1\}$ be the binary "outlier" map such that:

$$O(\mathbf{x}) = \begin{cases} 1 & \text{if } |\hat{\tau}(\mathbf{x}) - \mu_\tau(\mathbf{x})| > \lambda \kappa_\tau(\mathbf{x}), \\ 0 & \text{otherwise.} \end{cases} \quad (14)$$

For each pixel $\mathbf{x} \in \Omega$ such that $O(\mathbf{x}) = 1$ ("outlier"), let's denote $\mathcal{X}_{NESW} = \{\mathbf{x}_N, \mathbf{x}_E, \mathbf{x}_S, \mathbf{x}_W\}$ the set of 4 closest "inlier" lifetime locations along respectively the North, East, South and West directions such as $O(\mathbf{x}_N) = O(\mathbf{x}_E) = O(\mathbf{x}_S) = O(\mathbf{x}_W) = 0$. The fluorescence lifetime is then linearly interpolated as:

$$\hat{\tau}_b(\mathbf{x}) = \begin{cases} \frac{1}{C(\mathbf{x})} \sum_{\mathbf{y} \in \mathcal{X}_{NESW}} w(\mathbf{x}, \mathbf{y}) \hat{\tau}(\mathbf{y}) & \text{if } O(\mathbf{x}) = 1 \text{ and } \mathbf{x} \neq \mathbf{y}, \\ \hat{\tau}(\mathbf{x}) & \text{otherwise,} \end{cases}$$

with $C(\mathbf{x}) = \sum_{\mathbf{y} \in \mathcal{X}_{NESW}} w(\mathbf{x}, \mathbf{y})$, $w(\mathbf{x}, \mathbf{y}) = \|\mathbf{x} - \mathbf{y}\|_2^{-2}$ and $\|\cdot\|_2$ denotes the Euclidean norm. This interpolation technique can be very efficiently implemented as a two-pass filter and a demonstration of the method is illustrated in Figure 12 (simulation) and Figure 14 (experimental data). A more sophisticated but more computationally intensive interpolation method could be also considered as in [18]. In this approach, the idea is to exploit the set of pixels located at the periphery of each connected component of the binary map O to compute a weighted average lifetime.

In the next section, we describe a fluorescence lifetime estimation method dedicated to moving objects and particles.

4. ESTIMATION OF FD FLIM PARAMETERS ON MOVING OBJECTS AND ENDOSOMES

One of the advantage of FD FLIM over time domain techniques is the speed of operation that puts time lapse measurements of fluorescence lifetime on dynamic sample within reach. However, in the conventional estimation methods presented in Section 2 B as well as the approach proposed above, the fluorescent objects are supposed to be immobile. Nevertheless, biological processes often involve motion at various temporal and spatial scales, and often faster than the acquisition time of a lifetime map (a few seconds typically). As a result, motions can occur on living cells during the phase modulation. Figure 5 illustrates the impact of motions on the phase estimation. The lifetime mis-estimation results in a so-called "doppler effect" on the lifetime

Fig. 5. Visualization of sub-cellular structure motion during image acquisition and effect on the FLIM images. A: Example of image sequence with moving sub-structures (i.e. series of zoomed thumbnails) and of modulation signal S_θ with sine fitting in red. B: Global estimation of motion by maximum intensity projections (MIP) and ratio with initial intensity R_0 (right). C: Reconstructed lifetime image showing outliers in red. D: Distribution of lifetime values in the reconstructed lifetime map in C showing extreme, non physical (negative or too high) values. Scale bar is 10 μm .

map as illustrated in Figure 6 with an artefactual gradient in the lifetime map in the direction of motion. Indeed, when analyzing the intensity over the phase sequence to determine the lifetime, deviations from a sine function can be observed (see Figure 5 A) due to the object motion.

The issue has been previously diagnosed in [9] and further analyzed in [5] where the impact of moving object on lifetime variance is measured. However, to our knowledge, no method has been proposed to estimate the fluorescence lifetime on moving structures.

In this section, we propose a statistical framework for robust lifetime estimation on moving endosomes in order to provide a corrected reconstruction of the lifetime map τ . To alleviate complexity while taking into account the correlation of parameters, we propose an iterative estimation procedure where the estimation of FD FLIM and lifetime parameters and motion parameters are decoupled. The proposed method needs only the setting two non-critical input parameters: an upper-bound on the object number and the diffracted endosome size (e.g. a square window of 7×7 pixels is a typical choice and the scale may vary on microscope magnification).

Fig. 6. “Doppler effect” on lifetime measurement (color bar in nanoseconds) due to motion (see Figure 5). Scale bar is 10 μm .

A. Modeling spatial and temporal variations

In order to model the lifetime of a moving object, we propose a parametrization of both spatial and temporal intensity variations of a sub-resolved object. In our approach the microscope point spread function (PSF), corresponding to a Bessel function, is approximated as a Gaussian function as it is a common practice in spot detection literature [19]. Secondly the signal is composed of a background and a pinpoint vesicle before microscope diffraction. The background follows the model (4) due to the membrane imaging and/or cytoplasmic fluorescence. Finally, the background is assumed to be smooth over the PSF support.

More formally, $S(\mathbf{x}, k)$ is the signal defined at location $\mathbf{x} \in \Omega$ (pixel grid) as the sum of N object signals (e.g. vesicle, endosomes, ...) and a background signal:

$$S(\mathbf{x}, k) = S_{\theta_b}(\mathbf{x}, k) + \sum_{i=1}^N S_{\theta_i}(\mathbf{x}_i, k) e^{-\frac{\|\mathbf{x} - \mathbf{x}_{i,k}\|_2^2}{2\sigma_{i,k}^2}}, \quad \mathbf{x} \in \Omega \quad (15)$$

where $S_{\theta_b}(\mathbf{x}, k) \in \mathbb{R}_+$ defined in (4) is the spatially varying background signal parametrized by $\theta_b(\mathbf{x}) = (C_{S_b}(\mathbf{x}), A_{S_b}(\mathbf{x}), \phi_{\tau_b}(\mathbf{x}))^T$, $S_{\theta_i}(k)$ is the i -th object intensity parametrized by $\theta_i = (C_{S_i}, A_{S_i}, \phi_{\tau_i})^T$. In this modeling, $\sigma_{i,k}$ is the standard deviation representing the isotropic Gaussian shape of object i at frame k and $\mathbf{x}_{i,k}$ is the location of object i at frame k .

Given an object trajectory $\{\mathbf{x}_{i,1}, \dots, \mathbf{x}_{i,K}\}$ over K frames, we consider the following approximation for $\mathbf{x} \in \mathcal{B}(\mathbf{x}_{i,k}) \subset \Omega$ where $\mathcal{B}(\mathbf{x}_{i,k})$ is a local neighborhood (circle or square window) centered at $\mathbf{x}_{i,k}$ whose size is constant for all objects and for all frames in our study:

$$S(\mathbf{x}, k) \approx S_{\theta_i}(\mathbf{x}_{i,k}) e^{-\frac{\|\mathbf{x} - \mathbf{x}_{i,k}\|_2^2}{2\sigma_{i,k}^2}}, \quad \mathbf{x} \in \mathcal{B}_{i,k}. \quad (16)$$

In other words, we assume that the signal of an object i along its trajectory dominates locally other signal contributions.

B. Estimation procedure

Assuming a given set of images $\mathbf{I}_{1:K} = \{I(1, \cdot) : \Omega \rightarrow \mathbb{R}_+, \dots, I(K, \cdot) : \Omega \rightarrow \mathbb{R}_+\}$. We must first detect the N objects (e.g. vesicles) at locations $\{\mathbf{x}_{1,1}, \dots, \mathbf{x}_{N,1}\}$ on the first frame ($k = 1$). Object detection can be performed by using any robust algorithms (e.g. see [19–22]). We have considered the method #10 in [23] based on structure tensors [24] and an optimal histogram based thresholding [25] since this combination of algorithms was able to reliably

Fig. 7. Moving vesicle model.

determine the location of spots in fluorescence imaging. An initial value for all values of σ_i is set by the user to initialize the estimation procedure and the shape σ_i of all objects.

Given the initial set of objects, the trajectory $\mathbf{x}_{i,1:K} = \{\mathbf{x}_{i,1}, \dots, \mathbf{x}_{i,K}\}$, the shape parameter $\sigma_{i,1:K} = \{\sigma_{i,1}, \dots, \sigma_{i,K}\}$ and the parameters θ_i of object i are individually estimated using an iterative procedure described in the next section.

B.1. Maximum likelihood estimation framework

The basic idea of estimation for a given object indexed by i is to consider the parameter $\mathbf{x}_{i,1:K}$, $\sigma_{i,1:K}$ and θ_i on the same level. Let $p_{\theta_i}(\mathbf{x}_{i,1:K}, \sigma_{i,1:K}, \mathbf{I}_{1:K})$ the joint probability distribution of unknown variables and observations (i.e. measurements). The Generalized Maximum Likelihood (GML) estimation problem [26] amounts to optimizing the following criteria:

$$(\hat{\mathbf{x}}_{i,1:K}, \hat{\sigma}_{i,1:K}, \hat{\theta}_i) = \arg \max_{\mathbf{x}_{i,1:K}, \sigma_{i,1:K}, \theta_i} p_{\theta_i}(\mathbf{x}_{i,1:K}, \sigma_{i,1:K}, \mathbf{I}_{1:K}). \quad (17)$$

Under the proper assumptions of independence, the probability distribution can be generally factorized as follows:

$$p_{\theta_i}(\mathbf{x}_{i,1:K}, \sigma_{i,1:K}, \mathbf{I}_{1:K}) \propto \prod_{k=1}^K \prod_{\mathbf{x} \in \mathcal{B}(\mathbf{x}_{i,k})} p_{\theta_i}(\mathbf{x}, \mathbf{x}_{i,k}, \sigma_{i,k}, I(\mathbf{x}, k)). \quad (18)$$

This optimization problem can be implemented by successive and alternative minimization of criteria with respect to θ_i and $(\mathbf{x}_{i,1:K}, \sigma_{i,1:K})$ as:

$$\begin{cases} \hat{\theta}_i^{(m)} &= \arg \min_{\theta_i} -\log p_{\theta_i}(\hat{\mathbf{x}}_{i,1:K}^{(m)}, \hat{\sigma}_{i,1:K}^{(m)}, \mathbf{I}_{1:K}), \\ (\hat{\mathbf{x}}_{i,1:K}^{(m+1)}, \hat{\sigma}_{i,1:K}^{(m+1)}) &= \arg \min_{\mathbf{x}_{i,1:K}, \sigma_{i,1:K}} -\log p_{\hat{\theta}_i^{(m)}}(\mathbf{x}_{i,1:K}, \sigma_{i,1:K}, \mathbf{I}_{1:K}). \end{cases} \quad (19)$$

where m denotes the iteration number. Those minimizations, described in the following subsection, are carried out iteratively until convergence of the parameter of interest $|(\hat{\theta}_i^{(m)} - \hat{\theta}_i^{(m+1)})/\hat{\theta}_i^{(m)}| \leq \epsilon$ where ϵ is a constant (set to 10^{-6} in our experiments). Given $\hat{\theta}_i$, the decay time $\hat{\tau}_i$ for each detected vesicle is then estimated as (see Appendix A.2):

$$\hat{\tau}_i = \frac{\tan(\hat{\phi}_{\tau_i} - \phi_{M_0} + \phi_E)}{\omega_E}. \quad (20)$$

The final estimated value $\hat{\tau}_i$ of each object i is reported on the static lifetime map as a circular patch of value $\hat{\tau}_{i,0}$ at location $\hat{\mathbf{x}}_{i,0}$ for each vesicle.

B.2. Iterative and robust estimation of parameters

In the case of Gaussian model (10) with (18) and (16), we have

$$p_{\theta_i}(\mathbf{x}_{i,1:K}, \sigma_{i,1:K}, \mathbf{I}_{1:K}) \propto \exp \left(- \sum_{k=1}^K \sum_{\mathbf{x} \in \mathcal{B}(\mathbf{x}_{i,k})} \rho(r_{\theta_i}(\mathbf{x}, \mathbf{x}_{i,k}, \sigma_{i,k})) \right) \quad (21)$$

$$\text{with } r_{\theta_i}(\mathbf{x}, \mathbf{x}_{i,k}, \sigma_{i,k}) \triangleq \frac{I(\mathbf{x}, k) - S_{\theta_i}(\mathbf{x}_{i,k}) e^{-\frac{\|\mathbf{x} - \mathbf{x}_{i,k}\|_2^2}{2\sigma_{i,k}^2}}}{v(\mathbf{x}, k)}. \quad (22)$$

If the distribution of residuals $\{r_{\theta_i}(\mathbf{x}, \mathbf{x}_{i,k}, \sigma_{i,k})\}$ in the neighborhood $\mathcal{B}(\mathbf{x}_{i,k})$ is a zero-mean Gaussian, then $\rho(z) = z^2, z \in \mathbb{R}$ equivalent to a weighted least-mean square estimator. To increase robustness and reject possible outliers due to dead-pixels or model approximation, we consider a ρ -function that increases less rapidly than z^2 and becomes constant beyond a fixed point determined by a “scale” parameter. The ρ -function minimizes the effect of large residuals induced by model approximations and $\rho(z) \approx z^2$ for residuals smaller than the “scale” parameter. Our experiments showed that the Leclerc’s influence function defined as $\rho(z; s) = 1 - e^{-z^2/s^2}$ where $s > 0$ denotes the scale parameter, gives more stable results while compared with other M-estimators (e.g. Tukey Bi-weighted or German-McClure).

Object tracking. At iteration m of the algorithm, $(\mathbf{x}_{i,1:K}, \sigma_{i,1:K})$ are obtained from (19)-(21) as follows:

$$(\hat{\mathbf{x}}_{i,1:K}^{(m)}, \hat{\sigma}_{i,1:K}^{(m)}) = \arg \min_{(\mathbf{x}_{i,1:K}, \sigma_{i,1:K})} \sum_{k=1}^K \sum_{\mathbf{x} \in \mathcal{B}(\mathbf{x}_{i,k})} \rho \left(r_{\theta_i^{(m-1)}}(\mathbf{x}, \mathbf{x}_{i,k}, \sigma_{i,k}); s_{i,k}^{(m-1)} \right). \quad (23)$$

Estimation can be performed frame-by-frame and the parameters at frame k are given by:

$$(\hat{\mathbf{x}}_{i,k}^{(m)}, \hat{\sigma}_{i,k}^{(m)}) = \arg \min_{(\mathbf{x}_{i,k}, \sigma_{i,k})} \rho \left(r_{\theta_i^{(m-1)}}(\mathbf{x}, \mathbf{x}_{i,k}, \sigma_{i,k}); s_{i,k}^{(m-1)} \right). \quad (24)$$

The optimization (24) is thus performed using the iteratively reweighted least square method. The scale parameter is updated at each frame k as the median absolute deviation of the residuals as (see (14) and [17]):

$$s_{i,k}^{(m)} = \lambda \times 1.4826 \times \text{MAD} \left((r_{\theta_i^{(m-1)}}(\mathbf{x}, \mathbf{x}_{i,k}^{(m)}, \sigma_{i,k}^{(m)}))^2, \mathbf{y} \in \mathcal{B}(\mathbf{x}_{i,k}^{(m)}) \right) \quad (25)$$

and λ is set to 3 as explained in Section 3 according to the Chebychev’s inequality. For each frame k , the parameters are initialized with the estimators obtained on the previous frame: $\sigma_{i,k+1}^{(m)} = \hat{\sigma}_{i,k}^{(m)}, \mathbf{x}_{i,k+1}^{(m)} = \hat{\mathbf{x}}_{i,k}^{(m)}$ for $k \in [1, K]$. Let us note that in this framework, the motion is only a “nuisance parameter” and it is not required to track a object along the full FD FLIM measurement to estimate the fluorescence lifetime.

FD FLIM parameter estimation. Given an estimation $(\hat{\mathbf{x}}_{i,1:K}^{(m)}, \hat{\sigma}_{i,1:K}^{(m)})$, the FD FLIM parameters θ_i are estimated at iteration m as:

$$\hat{\theta}_i^{(m)} = \arg \min_{\theta_i} \sum_{k=1}^K \sum_{\mathbf{x} \in \mathcal{B}(\mathbf{x}_{i,k}^{(m)})} \rho \left(r_{\theta_i}(\mathbf{x}, \mathbf{x}_{i,k}^{(m)}, \sigma_{i,k}^{(m)}); s_{i,k}^{(m)} \right). \quad (26)$$

In this case, no model approximation is carried out and no outliers has been detected experimentally. To save time and ease the estimation process, the Leclerc ρ -function is then approximated by a quadratic function, i.e. $\rho(z) = z^2$ (“small” residuals), and (26) is equivalent to a weighted least squares minimization problem. The solution can be then found by using a

conventional Gauss-Newton algorithm as explained in Section 3.

B.3. Initialization of the estimation procedure

In our experiments, we considered small neighborhood $|\mathcal{B}(\mathbf{x}_i)|$ of about 7×7 pixels.

Initialization of object tracking parameters. In order to initialize the iteration process, each object detected on the first frame is tracked over the K frames of the FD FLIM measurement using simple frame-by-frame ordinary least-square model fitting over a local neighborhood $\mathcal{B}(\mathbf{x}_{i,k})$ centered at $\mathbf{x}_{i,k}$. In this initialization step, we have considered the following intensity model:

$$S(\mathbf{x}, k) \approx a(\mathbf{x}_{i,k}) e^{-\frac{\|\mathbf{x} - \mathbf{x}_{i,k}\|_2^2}{2\sigma_{i,k}^2}} + b(\mathbf{x}_{i,k}), \quad \mathbf{x} \in \mathcal{B}_{i,k}, \quad (27)$$

whose parameters $a(\mathbf{x}_{i,k})$ and $b(\mathbf{x}_{i,k})$ are constant values for $\mathbf{x} \in \mathcal{B}(\mathbf{x}_{i,k})$, jointly estimated at each frame k with the parameters $\mathbf{x}_{i,k}$ and $\sigma_{i,k}$. We denote $\hat{\mathbf{x}}_{i,k}^{(0)}$ and $\hat{\sigma}_{i,k}^{(0)}$ those parameters which are used as initializations of (24).

Initialization of FD FLIM parameters. At each iteration m , $\theta_i = (C_{S_i}, A_{S_i}, \phi_{\tau_i})^T$ is initialized using the intensity followed by the tracked object and physical property of the fluorescent donor. More specifically, the mean and amplitude of the normalized intensity of each spot over the object support are used to initialize C_{S_i} and A_{S_i} respectively. The phase ϕ_{τ_i} is initialized by using the fluorophore lifetime τ_{ref} in the absence of FRET. The fluorescence lifetime is determined from the literature or experimentally on fixed samples (e.g. 2.5 ns for the EGFP). For object i , $(\phi_{M_0} - \phi_E)$ is obtained by calibration of the optical system. The computation is carried out as follows:

$$\begin{cases} C_{S_i} &= \frac{1}{|\mathcal{B}(\mathbf{x}_i)|K} \sum_{k=1}^K \sum_{\mathbf{x} \in \mathcal{B}(\mathbf{x}_{i,k}^{(m)})} \tilde{I}(\mathbf{x}, k) \\ A_{S_i} &= \max_{\mathbf{x} \in \mathcal{B}(\mathbf{x}_{i,k}^{(m)}), k \in [1, K]} \tilde{I}(\mathbf{x}, k) - C_{S_i} \\ \phi_{\tau_i} &= \phi_{M_0} - \phi_E + \text{atan}(\omega_E \tau_{ref}) \end{cases}$$

$$\text{where } \tilde{I}(\mathbf{x}, k) \triangleq I(\mathbf{x}, k) e^{-\frac{\|\mathbf{x} - \mathbf{x}_{i,k}^{(m)}\|_2^2}{2(\sigma_{i,k}^{(m)})^2}}.$$

C. Summary of the algorithm

The iterative framework is summarized Figure 8. In a nutshell, the full lifetime map reconstruction algorithm is composed of noise calibration, static lifetime estimation and dynamical lifetime estimation on intracellular structure as given in Algorithm 1. Estimations converge after three to five iterations.

5. EXPERIMENTAL RESULTS

We tested our estimation method on both simulated and experimental sequences. As there can be no ground truth for the fluorescence lifetime on vesicles in experimental sequences, we have been using diverse strategies depending on the scenario at hand. To test our static phase estimator, we exploit a plain reference stack acquisition of fluorescein where a single mode in the phase distribution is expected. To evaluate the performances of dynamic object lifetime estimation, we use simulated data and compare the motion estimation with competitive multiple particle tracking algorithm.

Fig. 8. Iterative and alternative procedure for fluorescence lifetime estimation on moving endosomes.

Algorithm 1. Fluorescence Lifetime Estimation (FLE)

procedure NOISE CALIBRATION

1. Calibration of the spatially varying noise variance:
 $v^2(\mathbf{x}, k), \mathbf{x} \in \Omega$ and $k \in [1, K]$

procedure STATIC BACKGROUND ESTIMATION

1. Estimation of $\theta_b(\mathbf{x}), \forall \mathbf{x} \in \Omega$
2. Lifetime interpolation of pixels corrupted by motion

procedure FD FLIM, LIFETIME AND MOTION ESTIMATION

1. Detection and initial tracking of object $i \in [1, N]$
2. Iterative estimation of $\hat{\mathbf{x}}_{i,k}^{(m)}, \hat{\sigma}_{i,k}^{(m)}$ and parameters $\hat{\theta}_i^{(m)}, \forall i \in [1, N]$
3. Reporting of values $\hat{\tau}_i, \forall i \in [1, N]$ in the lifetime map.

A. FD FLIM setup

FLIM measurements were done in frequency domain by phase modulation on a custom system exploiting a third generation intensifier from Lambert Instruments (model II18MD). The module was attached to a NikonTE2000 TIRF inverted microscope equipped with a Coolsnap HQ CCD camera (Photometrics), a 473 nm modulated laser diode (Omicron) used for excitation of the donor fluorophore and a 100x 1.49 NA TIRF objective. The laser light was coupled using an optical fiber to the TIRF illumination arm or to a Yokogawa CSU10 spinning-disk module. Fluorescence emission was selected by a band-pass filter (500-550 nm). FLIM images were acquired either in widefield illumination using the TIRF illumination set to send the laser perpendicularly to the objective, or in confocal mode using the spinning-disk module. Transfected cells were also imaged in widefield illumination by a mercury lamp with standard Nikon filter cubes and a Coolsnap Ez CCD camera (Photometrics) prior to FLIM measurements in order to estimate the expression level of both GFP- and mCherry-tagged proteins. All the experiments were performed at 37°C.

B. Lifetime map estimation on static background

B.1. Phase estimation on a reference FD FLIM measurement

Figure 10 shows the phase distribution estimated on a FD FLIM acquisition of a sample of fluorescein presenting a single fluorescence lifetime that is constant throughout the sample. As a result, a single mode is expected for the estimated phase distribution. The fluorescence sample has been acquired with a wide-field setup. On this graph, we compared the Fourier decomposition used in the literature (or equivalently, the least mean squares estimator) and our heteroscedastic modeling approach. Those results highlights the significant shift caused by the weighting of least squares residuals (means are 2.46 ns for Fourier vs 2.45 ns, modes are the same) and difference in standard deviation (0.021 ns for Fourier vs 0.015 ns). Also, our correction results in a 10% gain in kurtosis which is expected for a plain fluorescein sample [8].

B.2. Lifetime estimation on a control experimental sample

Figure 9 presents the lifetime distribution estimated on an epithelial cell membrane (RPE1) using the Fourier decomposition and our weighted least squares method on a control sample. Experimental data are a subset of the acquisition carried out in the context of the experiment described Section D. In this experiment, fluorescence from GFP-tagged receptors localized on plasma membrane and early endosomes was acquired in the absence of an acceptor. Experiments were conducted using a confocal microscope with spinning disk setup and the same intensifier as above.

Our method estimates a narrower distribution, highlighting a standard deviation 0.27 ns while the Fourier decomposition estimates a standard deviation of 0.48 ns (means are 2.11 ns for Fourier vs 2.15 ns, modes are 2.08 ns and 2.09 ns). It is hard to assess the quality of a lifetime distribution in experimental conditions. However, the sample under study cannot present any transfer of energy and should thus present a single lifetime distribution. It follows that a narrow lifetime distribution with a single mode is expected. One can note the significant difference in the impact of our noise model on the phase of reference sequence and the lifetime estimate on the living cells. Indeed, two phase estimates are required to obtain the lifetime map, one applied on the reference sequence and an other for the sample at hand. This accentuates the effect on distribution kurtosis.

B.3. Quantification of phase estimation accuracy with increasing photon count

To further quantify the performances of our weighted least-square approach on static lifetime estimation, we have simulated plain fluorescein reference FD FLIM measurement and measured the phase estimation accuracy as a function of the simulated photon count. As shown in Figure 11, our method shows a 40% decrease in root mean square error (0.062 ns vs 0.038 ns) when compared to the Fourier (or equivalently least-square technique) using a sinusoidal signal presenting an offset ranging from 32 to 1000 photon counts. In those simulations, the minimum sample in each sinusoidal signals is 10% of the offset, reaching a minimum photon count of three.

C. Lifetime map reconstruction on dynamical sequences

We first illustrate our lifetime reconstruction and motion estimation on both noise-free and noisy simulations, we then highlight our results on experimental biological data.

Fig. 12. Noise-free simulation of motion impact on lifetime map and reconstruction. **a)** Simulated sequence with 5 vesicles exhibiting free Brownian motion modeled by a normal law for displacement with a standard deviation of 2 pixels (contrast has been enhanced for visualization). **b)** The estimated lifetime map on this FD FLIM measurement with the “doppler effect” footprint. **c)** Patches taking the value of the estimated fluorescence lifetime are added on the initial vesicle detection locus. The patch size is set to three times the size of the estimated spot scale **d)** A simple interpolation algorithm is applied to correct secondary artifacts for cosmetic purpose. **e)** True lifetime simulated on background and vesicles **f)** Motion estimation carried out during the iterative process. Scale bar is $10 \mu\text{m}$.

Fig. 9. Phase histograms using the Fourier method and the weighted least squares method exploiting the estimated noise scale on a reference FD FLIM measurement of a plain fluorescein sample.

Fig. 10. Lifetime histograms on live cell RPE1 with GFP-tagged receptor.

C.1. Lifetime map reconstruction on synthetic data

In order to focus the reconstruction of artifacts due to motion only, we applied our algorithm on simulations (see Figure 12) where the lifetime estimate is only affected by the vesicle motion without additional noise. The tracking is performed without issue, while parameter estimation takes 7 to 8 iterations to converge. The precision is about 10^{-6} after the first iteration on this simple example and additional iterations are not necessary.

We also tested the impact of our iterative approach on more challenging data. Those new simulations are based on our intensity model presented in Section A and the noise model presented in [10]. Background is simulated with an experimental vesicle-free sample and cytoplasmic fluorescence is simulated

Fig. 11. Phase estimation error with respect to the minimum photon count on static FD FLIM measurement ($4 \cdot 10^4$ FD FLIM measurements have been simulated for each photon count).

Table 1. Tracking and lifetime errors computed on simulated noisy data.

	Iteration m	0	1	2	3
Tracking mean error (pixel)		0.240	0.121	0.081	0.080
		(12%)	(6%)	(4%)	(4%)
Lifetime mean error (ns)		0.387	0.248	0.247	0.247
		(15%)	(9%)	(9%)	(9%)

in the background with constant parameter θ_b . Vesicle intensities follow (15) and motions are limited to a two-pixel displacement. The non-stationary Gaussian noise parameters are estimated on an experimental image series and then imposed on synthetic vesicles. Table 1 shows the convergence of lifetime and motion parameter on simulated vesicles. Convergence of tracking and lifetime errors is reached after $m = 3$ iterations on 30 simulated vesicles and 95 % of vesicles tracked. Similarly to our experiment on static fluorescence lifetime estimation, we measured the evolution of the root mean square error of the phase estimate on moving object under the same conditions (see Figure). In our iterative framework, we compare the use of a weighted least square estimator against a classic Fourier estimator for the iterative phase estimation. One can observe that the impact of the noise modeling is less strong on the dynamic sub-resolved objects (only 8%). This is expected given the impact of other variable such as motion estimation and model fitting error.

C.2. Endosome tracking accuracy in experimental conditions

We tested our reconstruction algorithm on living RPE1 cells expressing GFP-tagged Human Interferon Alpha Receptor Subunit-1 (IFN-R1), acquisitions that are further described in Section D. We highlight the result of our reconstruction method on experimental samples in Figure 14. The “doppler effect” can easily be seen on lifetime map computed using the usual Fourier method, and greatly reduced using our method.

However, it is very challenging to assess the accuracy of the estimated endosome fluorescence lifetime on experimental samples. Nevertheless, the lifetime parameter is closely related to the quality of the displacement estimate, and we can much easily measure the quality of motion tracking. To test our algorithm in these conditions, 30 moving spots localizing the GFP-tagged receptors have been imaged with our FD FLIM setup

Fig. 13. Phase estimation error with respect to the minimum photon count on dynamic FD FLIM measurement (100 objects have been simulated for each photon count).

and manually tracked using the MtrackJ software [19]. In order to highlight the effect of the proposed intensity modeling, we first compared the performances of our method against a frame-by-frame Gaussian fitting [27]. This simple method reuses the estimated position on the previous frame as an initialization for spot fitting. In order to compare our approach with a more competitive method, we also tracked those vesicles with the “probabilistic particle tracker” described in [28] and implemented in the ICY software [29]. This multiple particle tracking algorithm detects spots using an undecimated wavelet transform [20] on each frame before building tracks frame-by-frame. For each frame pair, the correspondences between the two sets of detections are optimized thanks to combinatorial optimization of the set of distances between detections. The distances are determined using the parameter of a Brownian model estimated through Kalman filtering. To address the problem of intensity variations that we modeled in our method, images are normalized using the Midway equalization method described in [30].

The detector proposed in [20] cannot adapt to local changes of SNR. As a result, the lowest SNR images that are normalized by our histogram equalization present an higher number of detected objects due to noise artifacts. However, as objects are sparsely localized in our experimental dataset, the three compared algorithms managed to track the 30 objects on the entire stack of 12 frames. The mean localization errors are represented in Figure 15 and the tracking with our method is visually illustrated Figure 16. As one could expect, the localization errors occur as the signal-to-noise ratio drops at lower intensities. Thanks to our thorough intensity modeling, our method presents a more precise spot localization throughout the whole sequence, a difference that intensifies on more noisy images.

D. Application to protein interaction analysis at the plasma membrane on endosomes

An usual approach to determine molecular nm-scale proximity in living cells is the measuring the FRET (Forster Resonance Energy Transfer) efficiency between the fluorophores linked to the two proteins of interest. Thanks to our motion compensation method, we can measure lifetime loss on dynamical structures such as trafficking endosomes as well as at the membrane. To trigger the FRET phenomenon on protein proximity, we considered IFN- α 1 GFP-tagged receptor as the donor protein and its associated kinase Tyk2 tagged with mCherry at different tagging sites as its acceptor (see Appendix 3 for an illustration of

Fig. 14. Lifetime reconstruction on experimental data. Left: Lifetime map using Fourier decomposition. Right: lifetime map reconstruction using our method. Scale bar is 10 μm .

Fig. 15. Mean localization error using Gaussian fitting [27], Probabilistic particle tracking [28] and our method on fluorescently tagged protein in living epithelial cells. Endosomes were tracked manually using the MtrackJ software [19].

their emission and excitation spectra). The sub-resolved spot that are localized outside of the cluttered recycling compartment are the endosomes of interest and the plasma membrane is described by the cell background.

To differentiate the lifetime estimated on the early endosomes from the membrane, we exploit the lifetime estimated on tracked vesicles for the former and we automatically segment the cell background for the latter (Figure 17).

We summarize the results obtained on a set of acquisition using cells expressing IFN- α 1-GFP alone or with Tyk2 fused to mCherry in C- or N-terminus in Figure 18:

- The first conclusion we can draw is that lifetime is lower in the condition where the acceptor is present, an indication that FRET might occur and that IFN- α 1 and its associated kinase Tyk2 may be in close proximity (at the nm scale) in RPE1 cells. This observation is consistent with the IFN- α 1/Tyk2 basal association described in the literature [31].
- No significant difference was found between endosomes and membranes, which validates also the lifetime estimation on moving endosomes. Without our algorithm, values would be probably different due to motion-induced artifact. We can also notice that the chosen site for tagging may have an impact on FRET efficiency. Indeed, both experiments with the two different tagging positions of mCherry on Tyk2 kinase show an higher lifetime loss.

Fig. 16. An example of IFN- α 1-GFP tracking on a 12 frames FD FLIM acquisition (images are normalized for visualization). Scale bar is 10 μm .

Fig. 17. Tracked IFN- α 1-GFP on endosomes (left) and membrane lifetime map (right, from 0 to 5ns, 0 is assigned to background and endosomes) in the absence (up) or presence (down) of acceptor. Scale bar is 10 μ m.

Fig. 18. Summary of lifetime analysis on endosomes (blue boxes) and membrane (red boxes) for RPE1 cells expressing IFN- α 1-GFP alone and in combination with Tyk2 labeled with mCherry at the C or N terminus (respectively mCherry-Tyk2 and Tyk2-mCherry) using the proposed algorithm.

6. CONCLUSION

In the first part of this paper, we have presented a dedicated signal processing method to reconstruct the fluorescein lifetime map corrupted by instrumental noise and intracellular motions in FD FLIM. In the second part, we proposed a fluorescence lifetime estimation method for dynamical analysis of endosomes inside the cell. Endosomes are detected on the first frame of the FD FLIM measurement. For each detected object, the lifetime and motion parameters are jointly estimated in an iterative and alternative fashion. On experimental sequences, the performance of the motion estimation is shown to outperform tracking algorithms that does not take the specific intensity model into account. The proposed methods are fully automatic and need only minimal and non-critical parameter setting.

The application of our method on the biological experiment presented in the last section demonstrated its efficiency in living cells. Thanks to those contributions we are hoping to pave the way for fluorescence lifetime imaging microscopy of dynamical intracellular structures. In particular, our method will allow the studies of protein-protein interactions in living cells during endocytosis process or signaling events on endosomes.

Two improvements could be envisaged for future work. To further improve the noise variance analysis, the next step is the study of the spatial correlation on the ICCD output. The estimation of multiple fluorescence lifetime on a pixel basis could be also interesting in the context of fluorophore with overlapping emission spectra. FD FLIM exploiting multiple frequencies have shown to be able to resolve multiple lifetimes [32, 33]. To improve the tracking capacity, more diversified intracellular structures should be considered. A few leads are already in this direction work, such as pixel soft classification between object and background.

More generally, the long-term goal is to perform time lapse FD FLIM measurement on a longer acquisition time. The motivation is to quantify protein interaction throughout for several cell cycles. If our method can already handle time lapse FD FLIM measurements, the time between each measurement must be short enough for the local motion compensation to perform properly. By spacing out FD FLIM measurements, the reduced photo-toxicity and photo-bleaching would allow much longer time-lapse acquisitions. The motion estimation between each FD FLIM measurement should thus be carried out with multiple particle tracking method. Semi-global optimization strategies to optimize the set of trajectories needs then to be investigated.

7. MATERIAL AND METHODS

Human Retinal Pigment Epithelial (RPE1) cells were maintained in culture in DMEM/F-12 medium (Gibco) supplemented with 10% bovine fetal serum. RPE1 cells were transfected with plasmids with X-tremeGene HP DNA (Roche) according to manufacturer guidelines. DNA constructs were verified by sequencing. Human Interferon Alpha Receptor Subunit-1 (IFN-R1) cDNA was cloned into pCAGGS (kind gift of Pr. Fukata). Human Tyk2 cDNA was extracted from pRC-HuTyk2-VSV plasmid (kind gift of S. Pellegrini), and cloned into pmCherry-N1 or -C1 (Clontech).

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APPENDIX A: DERIVATION OF THE PHASE MODULATED SIGNAL FOR FD FLIM

This appendix presents the derivation of the model used to recover the fluorescence lifetime from an excited sample. Mono-exponential decay and simple sinusoidal excitation is assumed. We start by deriving the equation of the excited signal, then we explain in details the experimental method to estimate the fluorescence lifetime of a given sample. To our knowledge, this derivation has not been described to such a full extent before in the literature.

A.1. The excited signal

Let $E(t)$ and $R(t)$ be an emission signal and the response of a fluorescent molecule respectively at time t and let τ be the unknown lifetime. The excited fluorescent signal is given by $F(t) = E(t) * R(t)$ where $*$ denotes the convolution operator. Assume a mono-exponential decay

and a simple sinusoidal excitation of the following form:

$$E(t) = C_E + A_E \sin(\omega_E t + \phi_E), \quad (28)$$

$$R(t) = R_0 e^{-t/\tau}, \quad (29)$$

where R_0 denote the response at time $t = 0$. It follows that

$$\begin{aligned} F(t) &= \int_0^\infty E(t-t') R(t') dt' \\ &= R_0 \left(C_E \int_0^\infty e^{-t'/\tau} dt + A_E \int_0^\infty \sin(\omega_E(t-t') + \phi_E) e^{-t'/\tau} dt' \right) \\ &= R_0 (C_E \tau + A_E B(t)) \end{aligned}$$

where

$$\begin{aligned} B(t) &= \int_0^\infty \sin(\omega_E(t-t') + \phi_E) e^{-t'/\tau} dt' \\ &= \left[-\tau \sin(\omega_E(t-t') + \phi_E) e^{-t'/\tau} \right]_0^\infty \\ &\quad - \omega_E \tau \int_0^\infty \cos(\omega_E(t-t') + \phi_E) e^{-t'/\tau} dt' \\ &= \tau \sin(\omega_E t + \phi_E) - \left[-\omega_E \tau^2 \cos(\omega_E(t-t') + \phi_E) e^{-t'/\tau} \right]_0^\infty \\ &\quad + (\omega_E \tau)^2 \int_0^\infty \sin(\omega_E(t-t') + \phi_E) e^{-t'/\tau} dt' \\ &= \tau \sin(\omega_E t + \phi_E) - \omega_E \tau^2 \cos(\omega_E t + \phi_E) - (\omega_E \tau)^2 B(t). \end{aligned}$$

It follows that

$$\begin{aligned} B(t) &= \frac{\tau}{1 + (\omega_E \tau)^2} (\sin(\omega_E t + \phi_E) - \omega_E \tau \cos(\omega_E t + \phi_E)) \\ &= \frac{\tau}{1 + (\omega_E \tau)^2} \left(\sin(\omega_E t + \phi_E) - \frac{\sin(\text{atan}(\omega_E \tau)) \cos(\omega_E t + \phi_E)}{\cos(\text{atan}(\omega_E \tau))} \right) \\ &= \frac{\tau \cos(\text{atan}(\omega_E \tau)) \sin(\omega_E t + \phi_E)}{(1 + (\omega_E \tau)^2) \cos(\text{atan}(\omega_E \tau))} \\ &\quad - \frac{\tau \sin(\text{atan}(\omega_E \tau)) \cos(\omega_E t + \phi_E)}{(1 + (\omega_E \tau)^2) \cos(\text{atan}(\omega_E \tau))} \\ &= \frac{\tau \sin(\omega_E t + \phi_E - \text{atan}(\omega_E \tau))}{(1 + (\omega_E \tau)^2) \cos(\text{atan}(\omega_E \tau))}. \end{aligned}$$

Finally, we have

$$\begin{aligned} F(t) &= R_0 \tau \left(C_E + A_E \frac{\sin(\omega_E t + \phi_E - \text{atan}(\omega_E \tau))}{\cos(\text{atan}(\omega_E \tau)) (1 + (\omega_E \tau)^2)} \right) \\ &= C_F + A_F \sin(\omega_E t + \phi_E - \text{atan}(\omega_E \tau)) \end{aligned}$$

where $C_F = C_E R_0 \tau$ (see [11]) and

$$A_F = A_E \frac{1}{\cos(\text{atan}(\omega_E \tau)) (1 + (\omega_E \tau)^2)} = \frac{A_E}{\sqrt{1 + (\omega_E \tau)^2}} \quad (30)$$

This result shows that the fluorescence response of the sample induces a phase delay and an amplitude decrease in the excited signal which is coherent with physical intuition.

A.2. Lifetime estimation

As we control the frequency of the emission signal, (30) shows that the phase and amplitude of the signal yield the lifetime τ (using a calibration reference measurement to retrieve ϕ_E). Actually, the experimental frequency in such experiment is too high for the phase to be directly measurable on a CCD sensor. To overcome this problem, it is usual to modulate the detection sensitivity with an other sinusoidal signal $M(t)$ (see (44)):

$$M(t) = C_M + A_M \sin(\omega t + \phi_M). \quad (31)$$

In the homodyne case, we use the same frequency as the original signal: $\omega_E = \omega_M$. The result of this modulation then yields to:

$$\begin{aligned} M(t)F(t) &= C_M C_F + A_M A_F \sin(\omega_E t + \phi_M) \sin(\omega_E t + \phi_E - \text{atan}(\omega_E \tau)) \\ &\quad + A_F C_M \sin(\omega_E t + \phi_E - \text{atan}(\omega_E \tau)) \\ &\quad + A_M C_F \sin(\omega_E t + \phi_M) \\ &= C_M C_F + A_F A_M \cos(\phi_E - \text{atan}(\omega_E \tau) - \phi_M) \\ &\quad + A_M A_F \cos(2\omega_E t + \phi_E - \text{atan}(\omega_E \tau) + \phi_M) \\ &\quad + A_F C_M \sin(\omega_E t + \phi_E - \text{atan}(\omega_E \tau)) \\ &\quad + A_M C_F \sin(\omega_E t + \phi_M). \end{aligned}$$

An intensified CCD is then required to acquire the lowest intensity images resulting from modulation. The higher frequencies of the phase modulated-signal $F(t)M(t)$ are attenuated by the low pass effect of the CCD detector denoted $\mathcal{G}_{\text{ccd}}$ resulting in the following approximation:

$$S(t) = \mathcal{G}_{\text{ccd}}(F(t)M(t)) \approx C_F C_M + A_F A_M \cos(\phi_E - \text{atan}(\omega_E \tau) - \phi_M) \quad (32)$$

The measured signal is then time-independent. One solution to recover the value $\text{atan}(\omega_E \tau)$ is to adjust ϕ_M as follows:

$$\phi_M = \phi_{M_0} + \frac{2\pi k}{K}, \quad k \in [1, K], \quad (33)$$

where ϕ_{M_0} denotes the phase for $k = 0$. We obtain K samples (K must be chosen > 3) defined as:

$$S(k) = C_F C_M + A_F A_M \cos\left(\phi_{M_0} + \frac{2\pi k}{K} - \phi_E + \text{atan}(\omega_E \tau)\right). \quad (34)$$

$$\triangleq C_S + A_S \cos\left(\frac{2\pi k}{K} + \phi_\tau\right), \quad k \in [1, K], \quad (35)$$

where $C_S = C_F C_M$,

$$\phi_\tau = (\phi_{M_0} - \phi_E + \text{atan}(\omega_E \tau)), \quad (36)$$

and

$$A_S = A_F A_M = \frac{A_M A_E}{\sqrt{1 + (\omega_E \tau)^2}}. \quad (37)$$

The observations $S(k)$ can be then fitted with a cosine function (using various methods including Fourier decomposition and robust regression methods) to estimate the parameter controlling (34), most notably the phase ϕ_τ and amplitude A_S . We can thus estimate the fluorescence lifetime, provided that we can calibrate the optical setup to measure $\phi_{M_0} - \phi_E$ or alternatively $A_E A_M$. To do so, a FD FLIM measurement of reference sample is with known lifetime τ_{ref} is acquired in the same condition. By fitting the $S_{\text{ref}}(k)$ with a cosine function, we obtain an estimate for the reference phase $\hat{\phi}_{\tau_{\text{ref}}}$ yielding :

$$(\widehat{\phi_{M_0}} - \widehat{\phi_E}) = \hat{\phi}_{\tau_{\text{ref}}} - \text{atan}(\omega_E \tau_{\text{ref}}) \quad (38)$$

$$\widehat{A_E A_M} = \widehat{A_F A_{M_{\text{ref}}}} \sqrt{1 + (\omega_E \tau_{\text{ref}})^2} \quad (39)$$

Finally, estimating the phase ϕ_τ of the signal described by 34 yields an estimate of the fluorescence lifetime:

$$\hat{\tau}_\phi = \frac{\tan(\hat{\phi} - \hat{\phi}_{\tau_{\text{ref}}} + \text{atan}(\omega_E \tau_{\text{ref}}))}{\omega_E}, \quad (40)$$

and denoted τ throughout the paper. Moreover, the estimation of the amplitude A_S combined with equation 47 yield an other approach to estimate the fluorescence lifetime using:

$$\hat{\tau}_M = \frac{1}{\omega_E} \sqrt{\left(\frac{\widehat{A_E A_M}}{\widehat{A_F A_{M_{\text{ref}}}} \sqrt{1 + (\omega_E \tau_{\text{ref}})^2}} \right)^2 - 1}. \quad (41)$$

APPENDIX B: NOISE VARIANCE MODELING

In this Appendix we summarize the derivation of a theoretical and experimental model for the variance of the noise induced by the ICCD setup previously published in [10]. Considering the gain and read out noise, Boddeke showed that the ICCD response $I(\mathbf{x}, k)$ at location $\mathbf{x} = (x, y) \in \Omega$ is [16]:

$$I(\mathbf{x}) = g_{\text{INT}} g_{\text{CCD}} \aleph(\mathbf{x}) + \zeta(\mathbf{x}) \quad (42)$$

where $\aleph(\mathbf{x})$ is the incident photon number on the ICCD which follows a Poisson law of parameter $\lambda(\mathbf{x})$, $\zeta(\mathbf{x})$ is the CCD read

out Gaussian noise such as $\zeta(\mathbf{x}) \sim (m_\zeta(\mathbf{x}), \sigma_\zeta^2(\mathbf{x}))$, g_{CCD} is the gain of the CCD sensor and g_{INT} is the gain of the intensifier assumed to be a realization of a random variable. Under the proper assumptions of independence of random variables it comes:

$$\begin{aligned} \text{Var}[I(\mathbf{x})] &= g_{\text{CCD}}^2 \text{Var}[g_{\text{INT}} \aleph(\mathbf{x})] + \sigma_\zeta^2(\mathbf{x})^2 \\ &= g_{\text{CCD}}^2 \text{Var}[g_{\text{INT}}] \text{Var}^2[\aleph(\mathbf{x})] \\ &\quad + g_{\text{CCD}}^2 (\mathbb{E}^2[g_{\text{INT}}] + \text{Var}[g_{\text{INT}}]) \text{Var}[\aleph(\mathbf{x})] + \sigma_\zeta^2(\mathbf{x})^2 \end{aligned} \quad (43)$$

By taking the expectation of (42) and since $\text{Var}[\aleph(\mathbf{x})] = \lambda(\mathbf{x}) = \mathbb{E}[\aleph(\mathbf{x})]$, it follows that:

$$\begin{aligned} \text{Var}[I(\mathbf{x})] &= \frac{\text{Var}[g_{\text{INT}}]}{\mathbb{E}^2[g_{\text{INT}}]} (\mathbb{E}[I(\mathbf{x})] - m_\zeta(\mathbf{x}))^2 \\ &\quad + g_{\text{CCD}} (\mathbb{E}[g_{\text{INT}}] + \frac{\text{Var}[g_{\text{INT}}]}{\mathbb{E}[g_{\text{INT}}]}) (\mathbb{E}[I(\mathbf{x})] - m_\zeta(\mathbf{x})) + \sigma_\zeta^2(\mathbf{x})^2. \end{aligned} \quad (44)$$

$$(45)$$

The local variance thus follows a quadratic relationship with the local expectation that we summarize with:

$$\text{Var}[I(\mathbf{x})] = a\mathbb{E}[I(\mathbf{x}, k)]^2 + b\mathbb{E}[I(\mathbf{x}, k)] + c \quad (46)$$

with a, b and $c \in \mathbb{R}$. We have shown experimentally in [10] that this model holds only locally. Indeed, when measuring the noise variance over the full spatial domain Ω , we notice a strong spatial dependency of the parameters (a, b, c). Further experiments in [10] have shown that the spatial dependency follows a bell-shape curve due to optical aberration induced by the lens in-between the phosphore screen and the CCD sensor in the intensifier. To correct for this aberration we have proposed the following model:

$$\text{Var}[I(\mathbf{x})] = (a\mathbb{E}[I(\mathbf{x})]^2 + b\mathbb{E}[I(\mathbf{x})] + c) \left(e^{-\frac{(x-x_0)^2}{2\sigma_x^2} - \frac{(y-y_0)^2}{2\sigma_y^2}} + o \right) \quad (47)$$

where $x_0, y_0, \sigma_x, \sigma_y$ and o are the center, scale and offset of the correction term. In [10], we propose a two-step approach for the estimation of the eight parameter controlling (47). Despite the simplicity of our correction model, the magnitude of the residuals of our estimation process indicate that our model is suitable to take into account the lens-induced aberration and fully explain the measured noise footprint.

APPENDIX C: EXCITATION AND EMISSION SPECTRA OF DONOR AND ACCEPTOR

Fig. 19. Excitation and emission spectra of EGFP (blue and green lines resp.) and mCherry (orange and red lines) fluorophores used for FRET experiments in our study.