Clustering in Stochastic Asynchronous Algorithms for Distributed Simulations

Anatoly Manita, François Simonot

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On the cascade rollback synchronization

Anatoli Manita*

François Simonot[†]

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Abstract

We consider a cascade model of N different processors performing a distributed parallel simulation. The main goal of the study is to show that the long-time dynamics of the system has a cluster behavior. To attack this problem we combine two methods: stochastic comparison and Foster–Lyapunov functions.

1 Introduction

The present paper contains a probabilistic analysis of some mathematical model of asynchronous algorithm for parallel simulation. For the detailed discussion of synchronization issues in parallel and distributed algorithms we refer to [2, 14]. Here we give only a brief description of the problem. In large-scale parallel computation it is necessary to coordinate the activities of different processor which are working together on some common task. Usually such coordination is implemented by using a so-called message-passing system. This means that a processor shares data with other processors by sending timestamped messages. Between sending or receiving the messages the processors work independently. It can be happened that till the moment of receiving of a message some processor can proceed farther in performing its program than the value of timestamp indicated in this newly received message; in this case the state of the processor should be *rolled back* to the indicated value. It is clear that due to possible rollbacks the mean speed of a given processor

*The work of this author is supported by Russian Foundation of Basic Research (grant 02-01-00945). Postal address: Faculty of Mathematics and Mechanics, Moscow State University, 119992, Moscow, Russia. E-mail: manita@mech.math.msu.su

[†]IECN, Université Henri Poincaré Nancy I, Esstin, 2, Rue J. Lamour, 54500 Vandoeuvre, France. E-mail: francois.simonot@esstin.uhp-nancy.fr

in the computing network will be lower than its proper speed. One of the most important performance characteristics of the system is the progress of the computing network on large time intervals.

Probabilistic models for such system are studied already for twenty years. From the probabilistic point of view these models consists of many relatively independent components which synchronize from time to time their states according to some special algorithm. The detailed review of all existing publications is out of range of this paper. We would like to mention only that the bibliography on this subject consists mostly of two group of papers. The first group of publication [3, 10, 11, 9, 6] are devoted to the case of two processors. The paper [3] is of special interest since it contains an exhaustive basic analysis of the two-dimensional model and had a big influence on the further research. The case of many processors was studied in [5, 6, 7, 15, 12, 13]. An important difference of such models from two-dimensional case is that in a realistic model with more than two processors one message can provoke a multiple rollback of a chain of processors. Since the multi-dimensional model is much more complicated for a rigorous study, in the above papers the authors deal with the set of identical processors and their mathematical results are contained in preparatory sections before large numerical simulations.

It should be noted also that probabilistic models with synchronization mechanism are interesting also for modelling database systems (see for example, [14]). Moreover, now synchronization-like interactions are considered as well in the framework of interaction particle systems [16, 17, 18].

The model considered in the present paper is of special interest for the following reasons. We deals with a nonhomogeneous model consisting of several different processors. We consider case of message-passing topology other from the topology of complete graph which was considered in all previous papers. Our main interest is the cascade model which pressupose a subordination between processors. We put forward a conjecture on the cluster behavior of the system: processors can be divided into separated groups which are asymptotically independent and have their own proper performance characteristics. Our main goal is to justify this conjecture. One should point out that in the case of complete graph topology the cluster decomposition into group is degenerated and, thus, not interesting.

We describe our model in terms of multi-dimensional continuous time Markov process. To get asymptotical performance characteristics of the model we combine two probabilistic methods (stochastic comparison and Foster–Lyapunov functions).

The paper is organized as follows. In Section 2 we give introduce a general continuous

time Markov model and define a cascade model as a special subclass of the general model. In Section 3 we pass to the embedded Markov chain. Main problem is now to study a long-time behavior of Markov chain with highly nonhomogeneous transition probabilities. To do this we consider relative coordinates and find groups of processors whose evolution is ergodic (converges to a steady state) in these relative coordinates. To our opinion the method of Foster-Lyapunov functions seems to be the only one to prove the stability in relative coordinates for the Markov chain under consideration. First of all in Section 5 we start from the case of two processors ($N = 2$) and the analysis here is rather simple and similar to [3]. In the study of the three-dimensional case (Section 7) the main point is the proof of ergodicity. We propose an explicit construction of some nonsmooth Foster-Lyapunov function. Our construction is rather nontrivial as it can be seen by comparing with already existing explicit examples of Lyapunov functions (see [1]). All this analysis brings us to some conclusions presented in Section 8. This section contains decomposition into groups (clusters) in the case of cascade model with any number of processors N and our main Conjecture 14. We show that future progress in proving this conjecture may be related with progress in explicit construction of multi-dimensional Foster-Lyapunov functions. Analysis of random walks in $\mathbf{Z}_+^n$ (which was done in [1]) shows that, in general, this technical problem may be very difficult. In the next papers we hope to overcome these difficulties by using specific features of our concrete Markov processes.

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2 Description of continuous time model

2.1 General model

We present here some mathematical model for parallel computations. There are N computing units (processors) working together on some common task. The state of a processor k is described by an integer variable $x_k \in \mathbf{Z}$ which is called a local (or inner) time of the processor k and has a meaning of amount of job done by the processor k up to the given time moment.

Assume that the state $(x_1, x_2, \dots, x_N)$ of the system evolves in continuous time $t \in \mathbf{R}_+$. Any change of a state is possible only at some special random time instants. Namely,

with any processor k we associate a Poissonian flow $\Pi^k = \{0 = \sigma_0^k < \sigma_1^k < \dots < \sigma_n^k < \dots\}$ with intensity λ_k and with a pair (k, l) of processors we associate a Poissonian flow $\Pi^{kl} = \{0 = \sigma_0^{kl} < \sigma_1^{kl} < \dots < \sigma_n^{kl} < \dots\}$ with intensity β_{kl} . This means, for example, that $\{\sigma_n^k - \sigma_{n-1}^k\}_{n=1}^\infty$ is a sequence of independent exponentially distributed random variables with mean λ_k^{-1} : $\forall n = 1, 2, \dots \quad \mathbf{P}\{\sigma_n^k - \sigma_{n-1}^k > 0\} = \exp(-\lambda_k s)$, and similarly for the flows Π^{kl} . We assume also that all these flows Π^k and Π^{kl} are mutually independent.

Let us now define a stochastic process $(X(t) = (x_1(t), \dots, x_N(t)), t \in \mathbf{R}_+)$ on the state space $\mathbf{Z}^N$ according to the following rules.

1) At time instants σ_n^k the processor k increases its local time x_k by 1: $x_k(\sigma_n^k + 0) = x_k(\sigma_n^k) + 1$.

2) There is an exchange of information between different processors. At time instant σ_i^{kl} the processor k sends a message $m_{kl}^{(x_k)}$ to the processor l . We assume that messages reach their destination immediately. A message $m_{kl}^{(x_k)}$ coming to node l from node k contains an information about local time $x_k(\sigma_i^{kl}) = x_k$ of the sender k . If at the time instant σ_i^{kl} (when the message $m_{kl}^{(x_k)}$ arrives to the node l) we have $x_l(\sigma_i^{kl}) > x_k(\sigma_i^{kl})$ then the local time x_l rolls back to the value x_k : $x_l(\sigma_i^{kl} + 0) = x_k(\sigma_i^{kl})$. Moreover, if the processor l rolls back, then all messages sent by the processor l during the time interval $\mathcal{I} = (\theta_l(x_k, \sigma_i^{kl}), \sigma_i^{kl})$, where $\theta_l(x, u) := \max\{s \leq u : x_l(s) = x, x_l(s+0) = x+1\}$, should be eliminated. This can provoke a cascading rollback of local times for some subset of processors. For example, assume that there is a processor q which received a message $m_{lq}^{(x'_l)}$ at some time instant $s' \in \mathcal{I}$ and $x_q(\sigma_i^{kl}) > x_l(s') = x'_l$. Then the local clock of q should be rolled back to the value $x_l(s')$: $x_q(\sigma_i^{kl} + 0) = x_l(s')$ and, moreover, all messages sent by q during the interval $\mathcal{I} = (\theta_q(x_l(s'), \sigma_i^{kl}), \sigma_i^{kl})$ should be eliminated. And so on. Hence, at time instant σ_i^{kl} a message from k to l can provoke a multiple rollback of processor $l, q, \dots$ in the system.

2.2 Cascade model

From now we shall consider the following special subclass of the above general models.

A chain of processors $1, 2, \dots, N$ is called a *cascade* if any processor j can send a message only to its right neighbour $j+1$. So noone sends messages to the processor 1 and the processor N does not send any messages. In other words, $\beta_{ij} \neq 0 \Leftrightarrow (j = i+1)$. A message sent from j to $j+1$ can provoke a cascading roll-back of processors $j+2, \dots$. Recall that all above time intervals are exponentially distributed and supposed to be independent. Obviously, the stochastic process $X_c^{(N)}(t) = (x_1(t), \dots, x_N(t))$ is Markovian. A very important property is that any “truncated” marginal process $X_c^{(N_1)}(t) = (x_1(t), \dots, x_{N_1}(t))$,

$N_1 \leq N$, is also Markovian.

Assume that for any j the following limit

$$v_j^* = \lim_{t \rightarrow +\infty} \frac{x_j(t)}{t} \quad (\text{in probability}) \quad (1)$$

exists. Then the numbers v_j^* , $j = 1, \dots, N$, characterize *performance* of the model. The main goal of the present paper is to prove existence of these limits and to calculate them.

Note that if we uniformly transform the absolute time scale $t = cs$, where $c > 0$ is a constant and s is a new absolute time scale, the performance characteristics (1) will not change.

3 Definition of discrete time model

Consider a sequence

$$0 = \tau_0 < \tau_1 < \tau_2 < \dots < \dots$$

of time moments when changes of local time at nodes may happen (we mean local time updates and moments of sending of messages). It is clear that $\{\tau_{r+1} - \tau_r\}_{r=0}^{\infty}$ is a sequence of independent identically distributed r.v. having exponential distribution with parameter

$$Z = \sum_{i=1}^N \lambda_i + \sum_{i=1}^{N-1} \beta_{i,i+1}.$$

Observing the continuous time Markov process $(x_1(t), \dots, x_N(t))$ at time moments τ_n we get a so-called embedded discrete time Markov chain $\{X(n), n = 0, 1, \dots\}$ with state space $\mathbf{Z}_+^N$. In the sequel we will be interested in the long-time behavior of the chain $\{X(n), n = 0, 1, \dots\}$.

Transition probabilities. In the MC $\{X(n), n = 0, 1, \dots\}$ there are transitions produced by a free dynamics and transitions generated by rollbacks. By the free dynamics we mean updating of local times

$$P\{X(n+1) = x + \mathbf{e}_j \mid X(n) = x\} = \lambda_j Z^{-1}, \quad j = 1, \dots, N.$$

It is easy to see that if a state $x = (x_1, \dots, x_N)$ is such that for some j $x_j < x_{j+1}$ then a message sended from j to $j + 1$ produces a transition of the following form

$$(x_1, \dots, x_j, x_{j+1}, \dots, x_l, x_{l+1}, \dots, x_N) \rightarrow (x_1, \dots, x_j, w_{j+1}, \dots, w_l, x_{l+1}, \dots, x_N) \quad (2)$$

with probability

$$Z^{-1} \beta_{j,j+1} \prod_{q=j+1}^{l-1} p(w_q, x_q; w_{q+1}, x_{q+1}) \times (1 - b_l)^{\min(x_l, x_{l+1}-1) - w_l + 1}, \quad (3)$$

where

- sequence $(w_{j+1}, \dots, w_l)$ is admissible in the following sense:

$$j < l \leq N, \quad w_{j+1} = x_j \quad w_q \leq w_{q+1} \leq \min(x_q, x_{q+1} - 1), \quad (j < q < l)$$

- $p(w_q, x_q; w_{q+1}, x_{q+1}) = b_q (1 - b_q)^{w_{q+1} - w_q}$
- $b_q = \frac{\lambda_q}{\lambda_q + \beta_{q,q+1}}, q < N.$

Here b_q is the probability of an event that processor q in state x_q sends at least one message to $q + 1$ before updating its state $x_q \rightarrow x_q + 1$. For $q = N$ we put $b_N = 0$. So in the case $l = N$ the probability (3) takes the form

$$Z^{-1} \beta_{j,j+1} \prod_{q=j+1}^{N-1} p(w_q, x_q; w_{q+1}, x_{q+1}).$$

Relative coordinates. Note that the first processor $x_1(t)$ evolves independently of other processors. It is useful to introduce new process $Y_c(t) = (y_2(t), \dots, y_N(t)) \in \mathbf{Z}^{N-1}$ in relative coordinates as viewing by an observer sitting at the point $x_1(t)$:

$$y_j(t) := x_j(t) - x_1(t), \quad j = 2, \dots, N.$$

In a similar way we define $Y(n) = Y_c(\tau_n)$, $n = 0, 1, \dots$. The free dynamics produces the following transitions of $Y(n)$:

$$P \{Y(n+1) = y + \mathbf{e}_j \mid Y(n) = y\} = \lambda_j Z^{-1}, \quad j = 2, \dots, N, \quad (4)$$

$$P \left\{ Y(n+1) = y - \sum_{j=2}^N \mathbf{e}_j \mid Y(n) = y \right\} = \lambda_1 Z^{-1}. \quad (5)$$

Since rollback does not affect on the first processor the corresponding transitions have the same form and the same probabilities as (2) and (3).

4 Stochastic monotonicity

All statements of this section are valid for the both Markov processes $X_c^{(N)}(t)$, $t \in \mathbf{R}_+$, and $X(n)$, $n \in \mathbf{Z}_+$. For the sake of brevity we give here results only for the continuous time model $X_c^{(N)}(t)$.

Theorem 1. *Let us consider two cascade models (say $X_{c,1}^{(n)}(t)$ and $X_{c,2}^{(n)}(t)$) with processors $1, 2, \dots, n$ and parameters $\lambda_1, \dots, \lambda_n$ and $\beta_{12}^{(1)}, \beta_{23}^{(1)}, \dots, \beta_{n-1,n}^{(1)}$ for the first model $X_{c,1}^{(n)}(t)$ and parameters $\lambda_1, \dots, \lambda_n$ and $\beta_{12}^{(2)}, \beta_{23}^{(2)}, \dots, \beta_{n-1,n}^{(2)}$ for the second model $X_{c,2}^{(n)}(t)$. Assume that*

$$\beta_{i,i+1}^{(1)} \leq \beta_{i,i+1}^{(2)} \quad \forall i.$$

Then the first model $X_{c,1}^{(n)}(t)$ stochastically dominates over the second model $X_{c,2}^{(n)}(t)$ in the following sense: if $X_{c,1}^{(n)}(0) = X_{c,2}^{(n)}(0)$ then for any time moment t we have¹

$$X_{c,1}^{(n)}(t) \geq_{st} X_{c,2}^{(n)}(t).$$

Proof may be given by an explicit coupling construction for the processes $X_{c,1}^{(n)}(t)$ and $X_{c,2}^{(n)}(t)$ on the same probability space. The following fact should be used: a Poisson flow with intensity $\beta_{12}^{(1)}$ can be obtained from a Poisson flow with intensity $\beta_{12}^{(2)}$ in which any point (independently from other) is killed with probability $1 - \beta_{12}^{(1)}/\beta_{12}^{(2)}$.

Corollary 2 (Solid barriers). *Fix some $1 \leq r_1 < r_2 < \dots < r_b < n$ and consider two cascade models: $X_{c,1}^{(n)}(t)$ with parameters $(\lambda_1, \dots, \lambda_n; \beta_{12}^{(1)}, \beta_{23}^{(1)}, \dots, \beta_{n-1,n}^{(1)})$ and $X_{c,2}^{(n)}(t)$ with parameters $(\lambda_1, \dots, \lambda_n; \beta_{12}^{(2)}, \beta_{23}^{(2)}, \dots, \beta_{n-1,n}^{(2)})$, where*

$$\beta_{i,i+1}^{(2)} = \beta_{i,i+1}^{(1)} \quad \forall i \notin \{r_1, \dots, r_b\}, \quad \beta_{i,i+1}^{(2)} = 0 \quad \forall i \in \{r_1, \dots, r_b\}.$$

We can say that the model $X_{c,2}^{(n)}(t)$ differs from the model $X_{c,1}^{(n)}(t)$ by the presence of b solid barriers between processors r_1 and $r_1 + 1, \dots, r_b$ and $r_b + 1$. Then by Theorem 1 we have

¹It means that there exists a *coupling* $(\tilde{X}_1^{(n)}(t, \omega), \tilde{X}_2^{(n)}(t, \omega))$ of stochastic processes $X_1^{(n)}(t)$ and $X_2^{(n)}(t)$ such that $P\left\{\omega : \tilde{X}_1^{(n)}(t, \omega) \geq \tilde{X}_2^{(n)}(t, \omega) \quad \forall t\right\} = 1$. If $w, z \in \mathbf{R}^n$ we say $w \geq z$ if $w_i \geq z_i$ for all $i = 1, \dots, n$ (partial order).

that

$$X_{c,1}^{(n)}(t) \leq_{st} X_{c,2}^{(n)}(t).$$

5 Case $N = 2$

We start with the Markov chain $X_c^{(2)}(t)$. Since processor 1 works independently, the main interest has Markov chain $Y_c^{(2)}(t) = x_2(t) - x_1(t)$.

Taking in mind the remark at the end of Subsection 2.2, for brevity of notation let us rescale absolute time in such a way that $Z = 1$. Then Markov chain $Y(n)$ has the following transition probabilities

$$p_{i,i+1} = \lambda_2, \quad p_{i,i-1} = \lambda_1, \quad p_{i,0} = \beta_{12} \ (i \geq 0), \quad p_{i,i} = \beta_{12} \ (i < 0)$$

and $p_{i,j} = 0$ for any another pair i, j .

a

Theorem 3. *If $\lambda_1 < \lambda_2$ then Markov chain $Y(n)$ is ergodic. If $\lambda_1 > \lambda_2$ then Markov chain $Y(n)$ is transient.*

Corollary 4. *If $\lambda_1 < \lambda_2$ then $v_1^* = v_2^* = \lambda_1$. If $\lambda_1 > \lambda_2$ then $v_1^* = \lambda_1$ and $v_2^* = \lambda_2$.*

Proof. Markov chain $Y(n)$ is one-dimensional and its analysis is quite easy. To establish ergodicity under assumption $\lambda_1 < \lambda_2$ we use the Foster-Lyapunov criterion (Theorem 17) with test function $f(y) = |y|$, $y \in \mathbf{Z}$. This implies that $x_2(t) - x_1(t)$ has a limit in distribution as $t \rightarrow \infty$. Recall that $x_1(t)$ is a Poissonian process hence the limit $v_1^* = t^{-1} \lim_t x_1(t) = \lambda_1$ exists (in probability). It follows from this that $v_2^* = t^{-1} \lim_t x_2(t) = \lambda_1$.

Under assumption $\lambda_1 > \lambda_2$ we get transiency by choosing the function $f(y) = \min(e^{\delta y}, 1)$, $y \in \mathbf{Z}$, where we fix sufficiently small $\delta > 0$, and applying Theorem 18. So any trajectory

of $Y(n)$ spends a finite time in any prefixed domain $\{y \geq C\}$ hence $x_2(t) - x_1(t) \rightarrow -\infty$ as $t \rightarrow \infty$ (a.s.). It means that after some time messages $1 \rightarrow 2$ can not produce a rollback anymore, so $x_1(t)$ and $x_2(t)$ become asymptotically independent and hence $v_2^* = t^{-1} \lim_t x_2(t) = \lambda_2$.

□

6 Case $N = 3$

Theorem 5. *Four situations are possible.*

1. If $\lambda_1 < \min(\lambda_2, \lambda_3)$ then $v_1^* = v_2^* = v_3^* = \lambda_1$.
2. If $\lambda_2 > \lambda_1 > \lambda_3$ then $v_1^* = v_2^* = \lambda_1$, $v_3^* = \lambda_3$.
3. If $\lambda_2 < \min(\lambda_1, \lambda_3)$ then $v_1^* = \lambda_1$, $v_2^* = v_3^* = \lambda_2$.
4. If $\lambda_1 > \lambda_2 > \lambda_3$ then $v_1^* = \lambda_1$, $v_2^* = \lambda_2$, $v_3^* = \lambda_3$.

Items 2, 3 and 4 can be reduced in some sense to the results on the case $N = 2$ (see Corollary 4). We prove them in the current section. Proof of item 1 is different and will be given in Sect. 7.

Proof of Theorem 5 (items 2–4). We start from the item 2: $\lambda_2 > \lambda_1 > \lambda_3$. Since the first two processors are governed by the Markov chain $X_c^{(2)}(t)$ and do not depend on the state of processor 3 we apply Corollary 4 and conclude that $X_c^{(2)}(t)$ is ergodic and $v_1^* = v_2^* = \lambda_1$.

Let us compare the following two cascade models

$$X_c^{(3)}(t) : \quad 1 \xrightarrow{\beta_{1,2}} 2 \xrightarrow{\beta_{2,3}} 3$$

$$X_{c,2}^{(3)}(t) : \quad 1 \xrightarrow{\beta_{1,2}} 2 \xrightarrow{0} 3$$

(parameters λ_1 , λ_2 and λ_3 are the same for the both models $X_c^{(3)}(t)$ and $X_{c,2}^{(3)}(t)$).

In the model $X_{c,2}^{(3)}$ the groups of processors $\{1, 2\}$ and $\{3\}$ evolve independently. Evidently, an asymptotic speed of processor 3 in the model $X_{c,2}^{(3)}$ exists and is equal to λ_3 . By Corollary 2 $X_c^{(3)}(t) \leq_{st} X_{c,2}^{(3)}(t)$. Hence in the model $X_c^{(3)}$ an asymptotic speed of the

processor 3 is *not greater* than λ_3 . Since $\lambda_3 < \lambda_1$ we conclude that there exists some time moment T_0 such that for $t \geq T_0$ in the model $X_c^{(3)}$ messages from 2 to 3 that roll back the processor 3 will be very “rare”. So these rare rollbacks will be not essential for an asymptotical speed of the processor 3. In other words, as $t \rightarrow \infty$ the groups of processors $\{1, 2\}$ and $\{3\}$ of the model $X_c^{(3)}$ become asymptotically independent, so the processor 3 will move with the average speed λ_3 .

Items 3 and 4 can be considered in a similar way. Note the item 3 consists of two subcases: $\lambda_1 > \lambda_3 > \lambda_2$ and $\lambda_3 > \lambda_1 > \lambda_2$. We omit details.

7 Explicit construction of Lyapunov function

In this section we prove the item 1 of Theorem 5. Recall that our key assumption here is

$$\lambda_1 < \lambda_2, \quad \lambda_1 < \lambda_3. \quad (6)$$

Main idea is to prove that the Markov chain $Y(n)$ is ergodic. To do this we apply the Foster-Lyapunov criterion (see Theorem 17 in Appendix). As in the case of Corollary 4 ergodicity of $Y(n)$ implies that $v_j^* = \lambda_1$, $j = 1, 2, 3$.

7.1 Transition probabilities

Consider the embedded Markov chain $Y(n)$. A stochastic dynamics produced by this Markov chain consists of two components: transitions generated by the free dynamics and transitions generated by roll-backs. For each transition probability $p_{\alpha\beta}$, $\alpha \neq \beta$, we have the following representation:

$$p_{\alpha\beta} = s_{\alpha\beta} + r_{\alpha\beta}, \quad (7)$$

where $s_{\alpha\beta} \geq 0$ corresponds to a transition $\alpha \rightarrow \beta$ which occurs due to the free dynamics and $r_{\alpha\beta}$ corresponds to a roll-back transition $\alpha \rightarrow \beta$.

Taking into account the remark at the end of Subsection 2.2, without loss of generality we assume that the time is rescaled in such way that $Z = 1$. This slightly simplifies notation for transition probabilities. For example, free dynamics transitions (4)–(5) are equal to λ_2 , λ_3 and λ_1 correspondingly. On the next figure we show all non-zero transitions $\alpha \rightarrow \beta$, ($\alpha \neq \beta$). It is true, of course, that $p_{\alpha\alpha} = 1 - \sum_{\beta \neq \alpha} p_{\alpha\beta}$, but it is useless to put this

information on the picture. Below we give explicit form of rollback transition probabilities:

$$\begin{aligned}
1 \rightarrow 2 : \quad & r_{yz} = \beta_{12} \quad \text{for } 0 < y_2 \\
2 \rightarrow 3 : \quad & r_{yz} = \beta_{23} \quad \text{for } y_2 < y_3 \\
1 \rightarrow 2 \rightarrow 3 : \quad & r_{yz} = \begin{cases} \beta_{12} (1 - b_2)^{z_3} b_2, & z_3 < y_3 \\ \beta_{12} (1 - b_2)^{y_3}, & z_3 = y_3 \end{cases} \quad \text{for } 0 < y_3 \leq y_2 \\
1 \rightarrow 2 \rightarrow 3 : \quad & r_{yz} = \begin{cases} \beta_{12} (1 - b_2)^{z_3} b_2, & z_3 \leq y_2 \\ \beta_{12} (1 - b_2)^{y_2+1}, & z_3 = y_3 \end{cases} \quad \text{for } 0 < y_2 < y_3
\end{aligned}$$

where $y = (y_2, y_3)$, $z = (z_2, z_3)$.

7.2 Contour of level 1

On the plain Oy_2y_3 let us consider an ellipse

$$e(y_2, y_3) = ay_2^2 + b(y_2 - y_3)^2 = 1, \quad a > 0, b > 0,$$

and draw a tangent line to it with normal vector $(-\Delta, 1)$. Evidently, there exist two tangent lines with the same normal vector $(-\Delta, 1)$. If $\Delta > 0$ is sufficiently large then one of this tangent line touches the ellipse at some point T_3 of the domain $y_2 < 0, y_3 < 0$. Take a segment on this line from the point T_3 to a point $K_3 = (0, u_3)$ of intersection with coordinate axis Oy_3 . Now let us draw tangent lines to the ellipse corresponding to a normal vector $(1, -\Delta)$. If $\Delta > 0$ is sufficiently large, then one of these lines touches the ellipse at some point T_2 of the domain $y_3 < 0$. Let us take this tangent line and fix a segment on it from the point T_2 to a point $K_2 = (u_2, 0)$ of intersection with coordinate axis Oy_2 . It is evident that $[K_2K_3] = \mathbf{R}_+^2 \cap \{(y_2, y_3) : y_2/u_2 + y_3/u_3 = 1\}$.

Let us consider now a closed contour L , consisting of subsequently joined segment K_3K_2 , segment K_2T_2 , arc T_2T_3 of the ellipse and segment T_3K_3 . This countour has the following property: any ray of the form $\{cv, c > 0\}$, where $v \in \mathbf{R}^2, v \neq 0$, has exactly one common point with the contour L .

Denote by $n(y)$ a vector of unit length which is an outer normal vector to the contour L at the point $y \in L$. The vector function $n(y)$ is well defined at all points of L except K_2 and K_3 and, moreover, this function is continuous on L except the points K_2 and K_3 . A behaviour of the function $n(y)$ on the arc T_2T_3 is of main interest:

$$n(y) = \frac{\nabla e(y)}{\|\nabla e(y)\|}, \quad \nabla e(y) = 2(ay_2 + b(y_2 - y_3), -b(y_2 - y_3)), \quad y \in T_2T_3 \subset L.$$

It is easy to see that $n(y) = n(T_2)$ for $y \in (K_2T_2]$, $n(y) = n(T_3)$ for $y \in [T_3K_3)$ and

$$n(y) = (u_2^{-1}, u_3^{-1}) \quad y \in (K_3K_2).$$

For the sequel it is important to point out the following points of the arc T_2T_3 : $y^{(3)} = (-a^{-1/2}, -a^{-1/2})$ and $y^{(2)}$, $\{y^{(2)}\} = T_2T_3 \cap \left\{y_3^{(2)} = \frac{a+b}{b}y_2^{(2)}\right\}$. It is easy to check that

$$n(y^{(2)}) \parallel Oy_3, \quad n(y^{(3)}) \parallel Oy_2.$$

Evidently, the both points belong to the domain $\{y_2 < 0, y_3 < 0\}$.

Lemma 6.

The function $n(y)$ has the following properties:

- $\langle n(y), y \rangle \neq 0 \quad \forall y \in L \setminus \{K_2, K_3\}$
- If $\Delta > 0$ is sufficiently large then there exist continuous functions $c_2(y)$ and $c_3(y)$ such that

$$c_2(T_2) = c_3(T_3) = 1, \quad c_2(T_3) = c_3(T_2) = 0, \quad c_2(y) > 0, c_3(y) > 0 \quad y \in (T_2, T_3)$$

$$n(y) = c_2(y)n(T_2) + c_3(y)n(T_3), \quad y \in (T_2, T_3).$$

- $\langle n(y), (0, -1) \rangle < 0$ if $y_2 < 0, y_3 > y_2$, and $\langle n(y), (-1, 0) \rangle < 0$ if $y_2 > 0, y_3 < 0$.

7.3 Definition of function φ

For any point $(y_2, y_3) \in \mathbf{R}^2 \setminus \{0\}$ define $\varphi(y_2, y_3) > 0$ such that

$$\frac{(y_2, y_3)}{\varphi(y_2, y_3)} \in L.$$

For $(y_2, y_3) = 0$ we put $\varphi(0, 0) = 0$. The function $\varphi(y_2, y_3)$ is well-defined and has the following properties:

- $\varphi : \mathbf{R}^2 \rightarrow \mathbf{R}_+$ (positivity)
- $\varphi(r y_2, r y_3) = r \varphi(y_2, y_3)$, $r > 0$, (homogeneity)
- $L = \{y : \varphi(y) = 1\}$.

To any point $y = (y_2, y_3)$ we put in correspondance a point $y^* := \frac{y}{\varphi(y)} \in L$. Therefore, $\varphi(y^*) = 1$.

Lemma 7.

- A gradient $\nabla\varphi(y)$ exists at all points except that y for which $y^* = K_2$ or K_3 and, moreover, the gradient is constant on rays of the form $\{cv, c > 0\}$, $v \in \mathbf{R}^2$:

$$\nabla\varphi(y) = \frac{n(y^*)}{\langle y^*, n(y^*) \rangle}.$$

- Let $y = (y_2, y_3)$ be such that $y^* \in T_2 T_3$. Then

$$|\varphi(w) - \langle \nabla\varphi(y^*), w \rangle| \leq \frac{\text{const}}{\varphi(y)} \|w - y\|^2. \quad (8)$$

In other words, in a neighbourhood of the point y the function φ can be approximated by the linear function $\langle \nabla\varphi(y^*), \cdot \rangle$.

In particular, $\varphi(y) = \langle \nabla \varphi(y^*), y \rangle$.

Proof of Lemma 7 is a straightforward computation.

7.4 Modification of the principle of local linearity

For any state α define a set $T_\alpha = \{\beta : p_{\alpha\beta} > 0\}$. Recall decomposition (7) and define $F_\alpha = \{\beta : s_{\alpha\beta} > 0\}$ and $R_\alpha = \{\beta : r_{\alpha\beta} > 0\}$. It is evident that $T_\alpha = F_\alpha \cup R_\alpha$. The most simple case is $F_\alpha \cup R_\alpha = \emptyset$. The case $F_\alpha \cup R_\alpha \neq \emptyset$ can be reduced to the previous one by a dilatation of the state space. Thus we assume that $F_\alpha \cup R_\alpha = \emptyset$ and consider events $\{Y(n+1) \in F_\alpha\}$ and $\{Y(n+1) \in R_\alpha\}$. On the set $\{\omega \in \Omega : Y(n, \omega) = \alpha\}$ we have $I_{\{Y(n+1) \in F_\alpha\}}(\omega) + I_{\{Y(n+1) \in R_\alpha\}}(\omega) \equiv 1$. Hence,

$$\begin{aligned} \mathbb{E} (f(Y(n+1)) | Y(n) = \alpha) - f(\alpha) &= \mathbb{E} ((f(Y(n+1)) - f(\alpha)) I_{\{Y(n+1) \in F_\alpha\}} | Y(n) = \alpha) + \\ &+ \mathbb{E} ((f(Y(n+1)) - f(\alpha)) I_{\{Y(n+1) \in R_\alpha\}} | Y(n) = \alpha) \end{aligned}$$

It follows from definition of Markov chain $Y(n)$ (see Subsect. 7.1) that diameters $d_\alpha := \text{diam } F_\alpha$ are uniformly bounded in α : $d = \max_\alpha d_\alpha < +\infty$. Define a vector

$$M_F(\alpha) = \mathbb{E} ((Y(n+1) - \alpha) I_{\{Y(n+1) \in F_\alpha\}} | Y(n) = \alpha) = \sum_{\beta \in F_\alpha} (\beta - \alpha) p_{\alpha\beta}.$$

This is an analogue of a notion of mean jump (see (12) in Appendix). In the next subsection we shall need the following modification of the principle of local linearity from [1] (see also Subsect. A.2 in Appendix).

Lemma 8. *Assume that the following condition holds*

$$\inf_l \sup_{\tilde{\alpha} \in \mathbf{R}^n, \|\tilde{\alpha} - \alpha\| \leq d_\alpha} |f(\tilde{\alpha}) - l(\tilde{\alpha})| < \varepsilon,$$

where $\inf$ is taken over all linear functions l . We state that from inequality

$$f(\alpha + M_F(\alpha)) - f(\alpha) < -5\varepsilon,$$

it follows that the following inequality

$$\mathbb{E} ((f(Y(n+1)) - f(\alpha)) I_{\{Y(n+1) \in F_\alpha\}} | Y(n) = \alpha) < -\varepsilon$$

holds.

Proof of this statement repeats the proof of principle of local linearity presented in [1] and, thus, is omitted.

7.5 Proof of Foster condition

The validity of the Foster condition will follow from several auxiliary lemmas dealing with the following different domains of the state space:

$$\begin{aligned} E_- &:= \{y = (y_2, y_3) : \min(y_2, y_3) < 0\}, \\ E_1 &:= \{y = (y_2, y_3) : y_2 > 0, y_3 > 0\}, \\ E_{1,2} &:= \{y_2 > 0, y_3 = 0\}, \\ E_{1,3} &:= \{y_2 = 0, y_3 > 0\}. \end{aligned}$$

Lemma 9. *Consider domain $E_- = \{y = (y_2, y_3) : \min(y_2, y_3) < 0\}$. There exists $C_9 > 0$, such that if $\varphi(y) > C_9$, then*

a)

$$\mathbb{E} \left((\varphi(Y(n+1)) - \varphi(y)) I_{\{Y(n+1) \in R_y\}} \mid Y(n) = y \right) \leq 0,$$

b) *there exists $\varepsilon > 0$ such that*

$$\varphi(y + M_F(y)) - \varphi(y) < -5\varepsilon.$$

Proof. It is evident that the vector

$$M_F(y) = (\lambda_2 - \lambda_1, \lambda_3 - \lambda_1)$$

is constant (does not depend on y). Since the vector $n(T_2)$ is co-directed with the vector $(1, -\Delta)$ and the vector $n(T_3)$ is co-directed with the vector $(-\Delta, 1)$ and the conditions $\lambda_2 > \lambda_1, \lambda_3 > \lambda_1$ hold, we can find a large $\Delta_1 > 0$ such that

$$\langle M_F(y), n(T_2) \rangle < 0, \quad \langle M_F(y), n(T_3) \rangle < 0, \quad \forall \Delta > \Delta_1.$$

Fix this Δ_1 . Hence, by Lemma 6 there exists $\varepsilon > 0$ such that

$$\langle M_F(y), n(y) \rangle < -6\varepsilon \quad \text{if} \quad \min(y_2, y_3) < 0.$$

Put $w = y + M_F(y)$ and consider

$$\begin{aligned}
\varphi(w) - \varphi(y) &= \varphi(w) - \langle \nabla \varphi(y^*), y \rangle \\
&= \varphi(w) - \langle \nabla \varphi(y^*), w \rangle + \langle \nabla \varphi(y^*), w - y \rangle \\
&= \varphi(w) - \langle \nabla \varphi(y^*), w \rangle + \langle \nabla \varphi(y^*), M_F(y) \rangle.
\end{aligned}$$

By (8) for any given $\varepsilon > 0$ we can choose $C_0 > 0$ such that

$$|\varphi(y + M_F(y)) - \langle \nabla \varphi(y^*), M_F(y) \rangle| \leq \frac{\text{const}}{\varphi(y)} d^2 \leq \varepsilon \quad \text{if } \varphi(y) \geq C_0.$$

Now the item b) of the lemma easily follows.

Let us prove the item a) of the lemma. Note that in the domain $y_3 > y_2, y_2 < 0$ a rollback decreases coordinate y_3 : $(y_2, y_3) \rightarrow (y'_2, y'_3) = (y_2, y_2)$. From geometrical properties of level sets of function φ and item 3 of Lemma 6 it follows that any transition generated by a rollback decreases a value of the function φ : $\varphi((y_2, y_3)) < \varphi((y'_2, y'_3))$. In the domain $y_3 < 0, y_2 > 0$ a rollback has the following form: $(y_2, y_3) \rightarrow (y'_2, y'_3) = (0, y_3)$. For similar reasons we again have $\varphi((y_2, y_3)) < \varphi((y'_2, y'_3))$. In the domain $y_3 \leq y_2 < 0$ there is no rollback. Now the item a) easily follows. $\square$

Lemma 10. Consider domain: $E_1 = \{y = (y_2, y_3) : y_2 > 0, y_3 > 0\}$.

1. The conditional expectation

$$\mathbb{E} \left((\varphi(Y(n+1)) - \varphi(y)) I_{\{Y(n+1) \in F_y\}} \mid Y(n) = y \right) = \langle (u_2^{-1}, u_3^{-1}), M_F(y) \rangle$$

does not depend on y .

2. There exist constants $C_{10}, \gamma_{10} > 0$ such that

$$\mathbb{E} \left((\varphi(Y(n+1)) - \varphi(y)) I_{\{Y(n+1) \in R_y\}} \mid Y(n) = y \right) \leq -\gamma_{10} \varphi(y) \quad \text{if } \varphi(y) > C_{10} \quad (9)$$

Proof. The second statement follows from the fact that in this domain $\varphi(y) = \langle (u_2^{-1}, u_3^{-1}), y \rangle$, and the vector $M_F(y)$ does not depend on y .

Let us prove (9). Fix some level set $\{y_2/u_2 + y_3/u_3 = C\} \cap \{y_2 > 0, y_3 > 0\}$ and consider an action of rollbacks. We have three different situations.

a) Let y be such that $y_2 \geq y_3 > 0$. It follows that $y_2 \geq \frac{C}{\frac{1}{u_2} + \frac{1}{u_3}}$. With probability β_{12} we have a rollback of the following form $(y_2, y_3) \rightarrow (0, y'_3)$ where $0 \leq y'_3 \leq y_3$. Hence the

function $\varphi(y)$ will decrease at least on the value

$$\frac{C}{\frac{1}{u_2} + \frac{1}{u_3}} : u_2 = \frac{C}{1 + u_2/u_3},$$

and the conditional mean $\mathbf{E} ((\varphi(Y(n+1)) - \varphi(y)) I_{\{Y(n+1) \in R_y\}} | Y(n) = y)$ will not exceed the value $-\beta_{12} (1 + u_2/u_3)^{-1} C$.

b) Let y be such that $0 < \frac{1}{2}y_3 \leq y_2 < y_3$. It follows that $y_2 \geq \frac{C}{2(\frac{1}{u_2} + \frac{1}{u_3})}$. With probability β_{12} we have a rollback $(y_2, y_3) \rightarrow (0, y'_3)$ where $0 \leq y'_3 \leq y_3$ and with probability β_{23} we have a rollback $(y_2, y_3) \rightarrow (y_2, y_2)$. The both of them will decrease the function φ . But the first rollback will decrease it at least on the value $\frac{C}{2(1 + u_2/u_3)}$. So we conclude that the above conditional mean will not exceed the value $-\frac{1}{2}\beta_{12} (1 + u_2/u_3)^{-1} C$.

c) Now let y be such that $0 < y_2 \leq \frac{1}{2}y_3$. It follows that $y_3 - y_2 \geq K(C)$ where

$$K(C) := \frac{1}{2} \cdot \frac{C}{\left(\frac{1}{2u_2} + \frac{1}{u_3}\right)} = \frac{Cu_3}{u_3/u_2 + 2}.$$

With probability β_{12} we have a rollback $(y_2, y_3) \rightarrow (0, y'_3)$, $0 \leq y'_3 \leq y_3$, which decreases the function φ , and with probability β_{23} we have a rollback $(y_2, y_3) \rightarrow (y_2, y_2)$, which decreases the value of the function φ at least on $K(C)/u_3$. Hence the conditional expectation will decrease at least on the value $-\beta_{23}K(C)/u_3 = -\beta_{23} (u_3/u_2 + 2)^{-1} C$.

□

Lemma 11. *Consider the cases when y belongs to the axes: $y \in E_{1,3}$, $y \in E_{2,3}$. Here*

$$\mathbf{E} ((\varphi(Y(n+1)) - \varphi(y)) I_{\{Y(n+1) \in R_y\}} | Y(n) = y) \leq -\gamma_{11}\varphi(y),$$

and

$$\mathbf{E} ((\varphi(Y(n+1)) - \varphi(y)) I_{\{Y(n+1) \in F_y\}} | Y(n) = y)$$

does not depend on y , where $y \in \{y : \varphi(y) > C_{11}\}$.

Proof. We consider in details the case $E_{1,3} = \{y_2 = 0, y_3 > 0\}$. We start with the free dynamics. The following transition

$$(0, y_3) \rightarrow (y'_2, y'_3) = (-1, y_3 - 1) \in E_3 = \{y_2 < 0, y_3 > y_2\}.$$

occurs with probability λ_1 . It is easy to see that for $y_3 > C_{11}$ values of the function $\varphi(\cdot)$

in the both points $(0, y_3)$ and $(-1, y_3 - 1)$ coincide with values of linear function $\langle n(T_3), \cdot \rangle$.

With probability λ_2 we have a transition

$$(0, y_3) \rightarrow (y'_2, y'_3) = (1, y_3) \in E_1 = \{y_2 > 0, y_3 > 0\},$$

and with probability λ_3 we have a transition

$$(0, y_3) \rightarrow (y'_2, y'_3) = (0, y_3 + 1) \in E_3.$$

Evidently, that in all these points $(0, y_3)$, $(1, y_3)$ and $(0, y_3 + 1)$ values of $\varphi(\cdot)$ coincide with values of a linear function $\langle n_1, \cdot \rangle$. Hence,

$$\begin{aligned} \mathbb{E} \left((\varphi(Y(n+1)) - \varphi(y)) I_{\{Y(n+1) \in F_y\}} \mid Y(n) = y \right) &= \lambda_1 \langle n(T_3), (-1, -1) \rangle \\ &\quad + \lambda_2 \langle n_1, (1, 0) \rangle + \lambda_3 \langle n_1, (0, 1) \rangle. \end{aligned}$$

Since the r.h.s. does not depend on y we get the second statement of the lemma.

Due to a rollback the Markov chain $Y(n)$ goes from the state $(0, y_3) \in E_{1,3}$ to a state $(0, 0)$ with probability β_{23} . Since values of $\varphi(\cdot)$ at these two points can be calculated by using the linear function $\langle n_1, \cdot \rangle$, the rollback decreases the value of φ on C , where $C = \varphi((0, y_3))$. The first statement of the lemma is proved.

The case of the domain $E_{1,3} = \{y_2 > 0, y_3 = 0\}$ is similar.

□

Lemma 12. *For any $C_0 > 0$*

$$\sup_{y: \varphi(y) \leq C_0} \mathbb{E} (\varphi(Y_{t+1}) \mid Y_t = y) < +\infty.$$

Proof. This statement follows from the fact that jumps from any fixed neighbourhood of $(0, 0)$ are bounded and the fact that the function φ is continuous. □

8 Conclusions, conjectures and perspectives

8.1 Decomposition into groups

We shall always assume that all $\lambda_1, \dots, \lambda_N$ are different. Define a function

$$\ell(m) := \min_{i \leq m} \lambda_i .$$

Evidently, this function has the following property:

$$\ell(1) = \lambda_1 \geq \dots \geq \ell(m) \geq \ell(m+1) \geq \dots \geq \min(\lambda_1, \dots, \lambda_N) = \ell(N) .$$

Level sets of function ℓ generate a partition of the set $\{1, \dots, N\}$. Namely, there exists a sequence $j_1 = 1 < j_2 < \dots < j_K < j_{K+1} = N+1$ such that the set of all processors can be divided into several nonintersecting groups

$$\{1, \dots, N\} = \bigcup_{k=1, \overline{K}} G_k , \quad (10)$$

$$G_k := \{j_k, j_k + 1, \dots, j_{k+1} - 1\} , \quad \ell(j_k - 1) > \ell(j_k) = \dots = \ell(j_{k+1} - 1) > \lambda_{j_{k+1}} .$$

Remark 13. An equivalent description of the group is possible. We say, for example, that $1, 2, \dots, k$ is a group if

$$\lambda_1 \leq \min(\lambda_2, \dots, \lambda_k) , \quad \lambda_1 > \lambda_{k+1} . \quad (11)$$

8.2 Long-time behaviour of the groups

Taking into account Theorems 3 and 5 and the above notion of groups of processors we put forward the following Conjecture.

Conjecture 14. Assume that all $\lambda_1, \dots, \lambda_N$ are different. For any j the following limit $v_j^* = \lim_{t \rightarrow +\infty} \frac{x_j(t)}{t}$ exists and $v_j^* = \ell(j)$.

So by this conjecture $v_j^* = \ell(j_k)$ for $j \in G_k$. If for some k the group G_k consists of *more than one* processor we may say that the processors of the group G_k are *synchronized*.

Remark 15 (On monotone cases). If $\lambda_1 < \dots < \lambda_N$ then $v_j^* = \lambda_1$ for any j . If $\lambda_1 > \dots > \lambda_N$ then for all j we have $v_j^* = \lambda_j$.

Let us discuss briefly perspectives of rigorous proof of the above Conjecture for large values of N . In fact, we have already prove this conjectures for wide class of cascade models.

Theorem 16. *Assume that all $\lambda_1, \dots, \lambda_N$ are different and a partition (10) of the set of processors $\{1, \dots, N\}$ is such that $|G_k| \leq 3$ for all k . Then the limits $v_j^* = \lim_{t \rightarrow +\infty} \frac{x_j(t)}{t}$ exist and $v_j^* = \ell(j)$.*

Proof of this statement is just a combination of the result of Theorem 5 (item 1) and arguments of the proof of items 2-4 of Theorem 5. We will not go into details.

So the key to the proof of Conjecture consists in generalization of item 1 of Theorem 5. As we see from Section 7 one of possible way of such generalization is an explicit construction of Foster-Lyapunov function in high dimensions. This seems to be a difficult technical problem which is out of the purposes of this paper.

A Appendix

Let $(\xi_n, n = 0, 1, \dots)$ be a countable irreducible aperiodic Markov chain with the state space $\mathcal{A}$.

A.1 Criteria

We use the following Foster criterion.

Theorem 17 ([1]). *The Markov chain ξ_n is ergodic if and only if there exists a positive function $f(\alpha)$, $\alpha \in \mathcal{A}$, a number $\varepsilon > 0$ and a finite set $A \in \mathcal{A}$ such that*

1)

$$\mathbb{E} (f(\xi_{n+1}) | \xi_n = y) - f(y) < -\varepsilon$$

for all $y \notin A$,

2) $\mathbb{E} (f(\xi_{n+1}) | \xi_n = y) < +\infty$ for all $y \in A$.

The following theorem give a criterion of transiency.

Theorem 18 ([1]). *The Markov chain is transient, if and only if there exists a positive function $f(\alpha)$ and a set A such that the following inequalities are fulfilled*

$$\mathbb{E} (f(\xi_{m+1}) | \xi_m = \alpha_i) - f(\alpha_i) \leq 0, \quad \forall \alpha_i \notin A,$$

$$f(\alpha_k) < \inf_{\alpha_j \in A} f(\alpha_j), \quad \text{for at least one } \alpha_k \notin A.$$

A.2 Principle of local linearity

From now we assume that the state space $\mathcal{A}$ is a subset of some $\mathbf{R}^k$. Define a vector of mean jump from the point α

$$M(\alpha) = \mathbf{E} ((\xi_{n+1} - \alpha) \mid \xi_n = \alpha) = \sum_{\beta} (\beta - \alpha) p_{\alpha\beta}. \quad (12)$$

Assume that $d_\alpha := \max_{\beta} \{|\beta - \alpha| : p_{\alpha\beta} > 0\} < \infty$ for all α .

The following principle of local linearity was proved in [1].

Lemma 19. *Assume that at some point α the following condition*

$$\inf_l \sup_{\tilde{\alpha} \in \mathbf{R}^n, \|\tilde{\alpha} - \alpha\| \leq d_\alpha} |f(\tilde{\alpha}) - l(\tilde{\alpha})| < \varepsilon,$$

holds, where $\inf$ is taken over all linear functions l . Then

$$f(\alpha + M(\alpha)) - f(\alpha) < -5\varepsilon \quad \implies \quad \mathbf{E} (f(\xi_{n+1}) - f(\alpha) \mid \xi_n = \alpha) < -\varepsilon.$$

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